

$$\left\{ \begin{array}{l} \cos \theta \quad \text{if } l \geq r_1 \text{ i.e. } \left\{ \begin{array}{l} l \geq r_1, \quad 0 \leq \theta \leq \pi/2 \quad \text{or} \\ l \leq r_1, \quad \cos^{-1} \frac{l}{r_1} \leq \theta \leq \frac{\pi}{2} \end{array} \right. \\ 0 \quad \text{otherwise} \end{array} \right.$$

Since the orientation of slabs is assumed to be statistically isotropic, the probability density of angle θ is the measure of relative solid angle giving θ , i.e.

$$\left\{ \begin{array}{l} \sin \theta \\ 0 \end{array} \right. \quad \begin{array}{l} 0 \leq \theta \leq \pi/2 \\ \text{otherwise} \end{array}$$

Therefore

$$\begin{aligned} m_2 &= \int_{r_1}^{\infty} \int_0^{\pi/2} \sin \theta \cos \theta g(l) d\theta dl \\ &\quad + \int_0^{r_1} \int_{\cos^{-1} \frac{l}{r_1}}^{\pi/2} \sin \theta \cos \theta g(l) d\theta dl \\ &= \int_{r_1}^{\infty} \frac{1}{2} g(l) dl + \int_0^{r_1} \frac{1}{2} \frac{l^2}{r_1^2} g(l) dl \\ &= \int_0^{\infty} \frac{1}{2} g(l) dl + \int_0^{r_1} \frac{1}{2} g(l) \left[-1 + \frac{l^2}{r_1^2} \right] dl \\ &= \frac{1}{2} S - \frac{1}{2} S_{r_1} \left(1 - \frac{\langle l^2 \rangle_{r_1}}{r_1^2} \right) \end{aligned}$$

and, with Equation (3.10)

$$\begin{aligned} \frac{m_2}{m} &= 1 - \frac{S_{r_1}}{S} \left(1 - \frac{\langle l^2 \rangle_{r_1}}{r_1^2} \right) \\ &= 1 - C_7(r_1) \end{aligned} \tag{3.22}$$

Analogous to the spherical model, the corresponding W_2 for both hot-wires inside the slab is

$$W_2 = \begin{cases} W - r_1 = \frac{l}{\cos \theta} - r_1 & \text{if } l \geq r_1 \text{ or} \\ 0 & \text{if } l \leq r_1, \cos^{-1} \frac{l}{r_1} \leq \theta \leq \frac{\pi}{2} \end{cases}$$

Therefore

otherwise

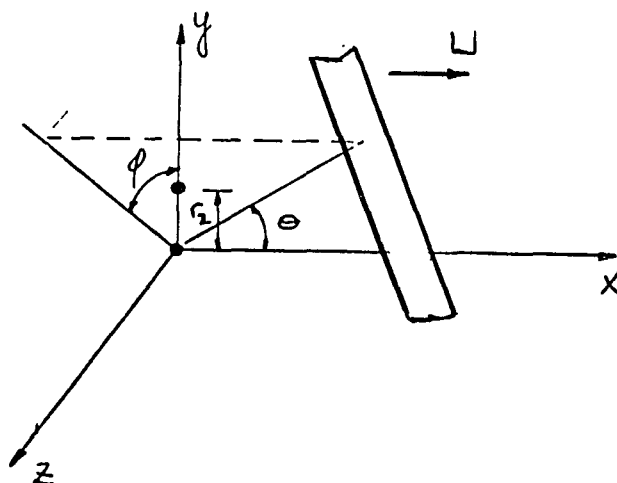
$$\begin{aligned} \gamma_2 &= \int_{r_1}^{\infty} \int_0^{\pi/2} \left(\frac{l}{\cos \theta} - r_1 \right) \cos \theta \sin \theta g(l) d\theta dl \\ &\quad + \int_0^{r_1} \int_{\cos^{-1} \frac{l}{r_1}}^{\pi/2} \left(\frac{l}{\cos \theta} - r_1 \right) \cos \theta \sin \theta g(l) d\theta dl \\ &= \int_{r_1}^{\infty} g(l) \left(l - \frac{r_1}{2} \right) dl + \int_0^{r_1} g(l) \left(\frac{l^2}{r_1} - \frac{1}{2} \frac{l^2}{r_1} \right) dl \\ &= \int_0^{\infty} g(l) \left(l - \frac{r_1}{2} \right) dl + \int_0^{r_1} g(l) \left(-l + \frac{r_1}{2} + \frac{1}{2} \frac{l^2}{r_1} \right) dl \\ &= S \langle l \rangle - \frac{r_1}{2} S + S r_1 \left(\frac{r_1}{2} - \langle l \rangle_{r_1} + \frac{1}{2} \frac{\langle l^2 \rangle_{r_1}}{r_1} \right) \end{aligned}$$

and, with Equations (3.11) and (3.12)

$$\begin{aligned} \frac{\gamma_2}{\gamma} &= 1 - \frac{r_1}{\langle W \rangle} + \frac{S r_1 \left(\frac{r_1}{2} - \langle l \rangle_{r_1} + \frac{1}{2} \frac{\langle l^2 \rangle_{r_1}}{r_1} \right)}{S \langle l \rangle} \\ &= 1 - \frac{r_1}{\langle W \rangle} + C_B(r_1) \end{aligned} \quad (3.23)$$

(B) Two hot-wires separated in the y-direction

As shown in Sketch 11, a slab is convected along the x-axis with its normal at an angle θ to the x-axis and its projection on the y-z plane at an angle ϕ to the y-axis. The slab is cut in a strip by the y-z plane as shown in

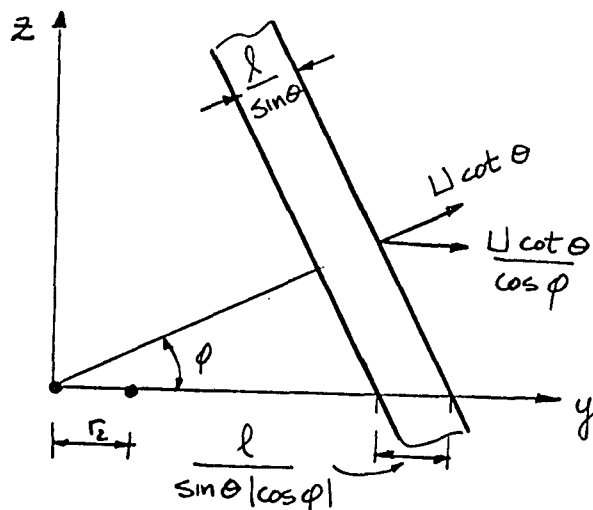


Sketch 11

Sketch 12. If the slab hits the wire at origin, the probability that it hits both wires simultaneously is

$$\begin{cases} 1 & \text{if } \frac{l}{\sin\theta|\cos\phi|} \geq r_2 \\ 0 & \text{if } \frac{l}{\sin\theta|\cos\phi|} < r_2 \end{cases}$$

i.e.,



Sketch 12

$$\begin{cases} 1 & \text{if (i) } l \geq r_2, \quad 0 \leq \theta \leq \frac{\pi}{2}, \quad 0 \leq \phi \leq \pi \\ & \text{(ii) } l < r_2, \quad 0 \leq \theta \leq \sin^{-1} \frac{l}{r_2}, \quad 0 \leq \phi \leq \pi \\ & \text{(iii) } l < r_2, \quad \theta > \sin^{-1} l/r_2, \quad \cos^{-1} \frac{l}{r_2 \sin\theta} < \phi < \pi - \cos^{-1} \frac{l}{r_2 \sin\theta} \\ 0 & \text{otherwise} \end{cases}$$

Similar to Equation (3.9), the expected number of slabs hitting one wire, with thickness between l and $l+dl$, orientation angles between θ and $\theta+d\theta$, φ and $\varphi+d\varphi$, is

$$\frac{1}{\pi} \sin\theta \cos\theta g(l) d\varphi d\theta dl$$

Therefore

$$\begin{aligned} m_2 &= \int_{r_2}^{\infty} \int_0^{\pi/2} \int_0^{\pi} \frac{1}{\pi} \sin\theta \cos\theta g(l) d\varphi d\theta dl \\ &+ \int_0^{r_2} \int_0^{\sin^{-1} l/r_2} \int_0^{\pi} \frac{1}{\pi} \sin\theta \cos\theta g(l) d\varphi d\theta dl \\ &+ 2 \int_0^{r_2} \int_{\sin^{-1} l/r_2}^{\pi/2} \int_{\cos^{-1} \frac{l}{r_2 \sin\theta}}^{\pi/2} \frac{1}{\pi} \sin\theta \cos\theta g(l) d\varphi d\theta dl \\ &= \int_{r_2}^{\infty} \frac{1}{2} g(l) dl + \int_0^{r_2} \int_0^{\sin^{-1} l/r_2} \sin\theta \cos\theta g(l) d\theta dl \\ &\quad + \int_0^{r_2} \int_{\sin^{-1} l/r_2}^{\pi/2} \frac{1}{\pi} g(l) \left[\pi - 2 \cos^{-1} \frac{l}{r_2 \sin\theta} \right] \sin\theta \cos\theta d\theta dl \\ &= \int_0^{\infty} \frac{1}{2} g(l) dl - \int_0^{r_2} \frac{1}{2} g(l) dl + \int_0^{r_2} \frac{1}{2} \left(\frac{l}{r_2} \right)^2 g(l) dl \\ &\quad + \int_0^{r_2} \int_{\sin^{-1} \frac{l}{r_2}}^{\pi/2} \frac{1}{\pi} g(l) 2 \left[\frac{l}{r_2 \sin\theta} + O\left(\frac{l}{r_2 \sin\theta} \right)^3 \right] \sin\theta \cos\theta d\theta dl \\ &= \frac{1}{2} S - \frac{1}{2} \int_0^{r_2} g(l) \left[1 - \left(\frac{l}{r_2} \right)^2 \right] dl \\ &\quad + \int_0^{r_2} \frac{2}{\pi} g(l) \frac{l}{r_2} \left[1 - \frac{l}{r_2} + O\left(\frac{l}{r_2} \right)^2 \right] dl \\ &= \frac{1}{2} S - \frac{1}{2} S_{r_2} \left[1 - \frac{4}{\pi} \frac{\langle l \rangle_{r_2}}{r_2} + \left(\frac{4}{\pi} - 1 \right) \frac{\langle l^2 \rangle_{r_2}}{r_2^2} + O\left(\frac{\langle l^3 \rangle_{r_2}}{r_2^3} \right) \right] \end{aligned}$$

and, with Equation (3.10)

$$\begin{aligned} \frac{m_2}{m} &= 1 - \frac{S_{r_2}}{S} \left[1 - \frac{4}{\pi} \frac{\langle l \rangle_{r_2}}{r_2} + \left(\frac{4}{\pi} - 1 \right) \frac{\langle l^2 \rangle_{r_2}}{r_2^2} + O\left(\frac{\langle l^3 \rangle_{r_2}}{r_2^3} \right) \right] \\ &= 1 - C_q(r_2) \end{aligned} \tag{3.24}$$

IV. EXPERIMENTAL RESULTS

4.1 Visual Observation of Fine-Structure Intermittency

The direct visual observation of over-all turbulence structure has been somewhat successful in a smoke tunnel or water channel (with dye or tiny bubbles). Nevertheless, present flow-visualization techniques are inadequate for a study of the fine-scale structure of turbulent fluctuations, since these techniques fail to extract fine-scale components and reveal them. Also they have not yet proved viable at the relatively high speeds necessary to generate turbulence whose Reynolds number is large enough to have "internal intermittency" of the fine-structure. An indirect visualization may be achieved by electronically examining narrow-band frequency components of the over-all turbulent signal.

Figure 21 shows typical oscillograms of band-pass and high-pass signals from a hot-wire placed in a grid-generated turbulent flow field ($R_\lambda = 110$). The low frequency signal is more or less uniformly distributed in time, while the high frequency signals appear to be intermittent. With Taylor's hypothesis, frequency is proportional to wave number, so the high frequency signals correspond to the velocity fluctuations associated with the fine-scale components of motion, and the temporal variation of a signal from a fixed hot-wire as it is convected past the hot-wire by the mean flow corresponds to the spatial variation of the turbulence pattern. Therefore, the intermittency of high frequency signal in time domain implies the localization of the fine-structure in space domain. The

time interval when the high frequency signal is "zero" corresponds to the time when the hot-wire is in a spatial region in which the fine-scale components are negligible.

Figure 22 shows the oscillograms of the total turbulent signal and of its first and second time derivatives. The derivative signals emphasize the fine-scale components, and their crest factors appear to increase as the order of derivative increases. However, they do not appear intermittent as the high frequency signal does. This is because the Reynolds number of the turbulence is not high enough to remove the spectrum of derivative signals far away from energy containing eddies, which is more or less uniformly distributed in space. Since the viscous dissipation of turbulent kinetic energy occurs primarily in the fine-structure, the fine-structure intermittency implies that the energy dissipation (or the first derivative) will appear intermittent if the Reynolds number of the turbulent flow is high enough such that the spectrum of energy dissipation decouples itself from energy spectrum. See, for example, the high Reynolds number flows of Sheih⁵, of Stegen and Gibson⁶, and of Grant, Stewart and Moilliet.³⁵

4.2 Flatness Factor

The flatness factor F of a random variable e is defined as

$$F \equiv \frac{\langle e^4 \rangle}{\langle e^2 \rangle^2} = \frac{\int_{-\infty}^{+\infty} e^4 P(e) de}{\left\{ \int_{-\infty}^{+\infty} e^2 P(e) de \right\}^2} \quad (4.1)$$

where $P(e)$ is the probability density of e . Since the fourth moment depends heavily on the large values of e , the flatness factor is a measure of the relative extent of the skirts of the probability density curve. A random variable with normal density has a flatness factor of 3.0 and $(F-3.0)$ for an arbitrary random variable is termed the "kurtosis." The roughly normal variable whose probability density function is more peaked in the neighborhood of the mean than is a normal density of the same standard deviation will have positive kurtosis.

As seen in Equation (4.1), a probability density function will uniquely determine the flatness factor, but the converse is not true. Though an intermittent variable is likely to have a large flatness factor, a large flatness factor does not necessarily imply intermittency. Therefore, a flatness factor can be used to indicate the degree of intermittency of a random variable only if it is known by other observations that the variable is intermittent. Batchelor and Townsend¹ suggested a relation between flatness factor and intermittency factor

$$\gamma = \frac{3.0}{F} \quad (4.2)$$

by assuming the intermittent variable varies with a normal probability distribution for a fraction γ of the total time, and is zero for the remainder of the time.

Having observed the fine-structure intermittency on the oscillograms as described in the previous section, the flatness factors of the velocity derivatives and band-pass signals were measured to offer

an indication to the degree of intermittency and the degree of deviation from normal distribution. The flatness factors of the velocity fluctuations u were also measured and found to be quite close to 3.0 as were found before by other investigators.

The noise spectrum of a thermally compensated hot-wire signal increases with frequency, while the energy spectrum of the turbulence decreases sharply with increasing frequency in the high frequency range. Differentiation accentuates the high frequencies present in the total signal, so the differentiation tends to reduce the signal-to-noise ratio. Therefore some kind of low-pass filter is necessary to cut off the high frequencies at which the signal-to-noise ratio is smallest.

The effect of high cut-off frequency on the flatness factors of the first and the second derivatives of the signal are shown in Figures 23 and 24, respectively. The flatness factors tend to increase with increasing cut-off frequency, and the rate of increase is higher for the second derivative than for the first, and for the higher Reynolds number flow; both differentiation and R_λ increase give relatively more high frequency "energy." This tendency agrees with the fact that the flatness factors of band-pass signals increase with frequency, to be discussed later. If the cut-off frequency is too high, however, the output includes more additional noise than signal and the flatness factor begins to level off, since the noise is approximately normal. The flatness factor therefore will eventually decrease as the cut-off frequency is raised to the value where the filter output includes more and more noise, with its flatness

factor of about 3.0. This leveling-off tendency is observed particularly in the first derivative signals of lower Reynolds number flows, which have less high frequency "energy."

The flatness factors of the first and the second derivatives were measured and are shown in Figures 25 and 26 as functions of R_λ , which ranged from 12 in a grid-generated turbulence to 830 on the axis of a round jet. To be consistent for differing Reynolds number flows, the cut-off frequencies of the filters were set at the frequencies of the Kolmogorov microscales.* This cut-off frequency is much higher than the frequency at which the peak of the spectrum of the derivative signal is located. Some measurements by other investigators are included in Figures 25 and 26 for comparison. Since Batchelor and Townsend¹ did not give the cut-off frequency of their filter, an effort was made to discover their procedure and to adjust their second derivative data to the values consistent with cut-off at f^* .* Their first derivative data were not adjusted because the variation of the flatness factor is comparatively small for low Reynolds number flows, as shown by Figure 23. By repetition of their experiments, we established the strong likelihood that they used a single cut-off frequency for all cases. The present flatness factor data were plotted as a function of cut-off frequency, and the flatness factors at various cut-off fre-

i.e., the frequency $f^ \equiv U/2\pi\eta$ corresponding to the convection of η -scale fine-structure past the hot-wire $\eta \equiv (\nu^3/\varepsilon)^{1/4}$.

quencies (for fixed R_λ 's) were obtained by extrapolation or interpolation. The flatness factors as a function of R_λ , with various constant cut-off frequencies, were plotted and compared with the Batchelor and Townsend data. The one with cut-off frequency at 3.5 kHz was found to agree best with their data and is shown in Figure 27. The upper line is a plot of the "correction" factor, i.e., the ratio of the flatness factors with cut-off frequency at f^* to those with cut-off at 3.5 kHz. The "corrected" Batchelor and Townsend data were obtained by multiplying their reported values by correction factors at the appropriate Reynolds numbers.

Batchelor and Townsend^{1,9} inferred from the original Kolmogorov similarity hypothesis that the limiting values of the flatness factors of velocity derivatives are independent of the large-scale properties of the turbulence and should reach universal constant values at large enough Reynolds numbers. Figure 25 shows that the flatness factor of the first derivative behaves like $R_\lambda^{0.2}$ at Reynolds number below 200, followed by a transition zone up to $R_\lambda \approx 500$, and then behaves like $R_\lambda^{0.62}$ instead of leveling off. In Figure 26, the flatness factor of the second derivative behaves like $R_\lambda^{0.25}$ up to $R_\lambda \approx 100$, and like $R_\lambda^{0.75}$ for $R_\lambda \geq 300$. Apparently, the flatness factors of the derivatives show no sign of approaching constants at Reynolds numbers R_λ of the order of thousands, which is believed to be high enough for the universal similarity hypothesis to apply.

In fact, it is doubtful that the flatness factors of the derivatives are determined wholly by the large wave-number components of

the turbulence. Suppose we have a turbulence with large enough Reynolds number that the derivative signal appears intermittent (it was observed that the second derivative signal at $R_\lambda = 830$ appeared intermittent), then the flatness factor of this intermittent signal depends in part on the signal durations over which the signal is zero. But the statistics of these durations are associated with the scale of the energy-containing eddies, not the smallest eddies.

Corrsin¹⁹ and Tennekes²⁰ explored some consequences of simple models which included intermittency of the fine-structure, and estimated the consequent dependence of the flatness factor of the first derivative on Reynolds number to be $\sim R_\lambda^{1.5}$ and $\sim R_\lambda$, respectively. The result of this experiment shows a weaker dependence.

The Yaglom and Gurvich¹⁸ model with log-normal distribution of positive random variables predicts that

$$\frac{\langle \phi^k \rangle}{\langle \phi \rangle^k} \sim R_\lambda^{\frac{3}{4} \mu k(k-1)} \quad (4.3)$$

where ϕ is a non-negative quantity governed by fine-scale components, μ is a universal constant estimated to be 0.4 by Yaglom and Gurvich from measured spectra of $(\partial u / \partial t)^2$ and $(\partial \omega / \partial t)^2$. If we take $\phi = \left(\frac{\partial u}{\partial t} \right)^2$ and $K = 2$, we get the flatness factors of the derivatives proportional to $R_\lambda^{0.6}$, which agrees with the high Reynolds number data of the first derivative signal.

The flatness factors of band-pass signals were measured in the grid-generated turbulence at $R_\lambda = 110$ and 86.5. Figure 28 shows the effect of relative bandwidth on the flatness factors for a fixed mid-band frequency. The data were replotted in Figure 29 on a semi-log scale to accentuate the peakedness of the curves. All the curves peak around $\Delta f / f_m = 0.3$, and the higher the mid-band frequency, the more pronounced the peak is. For a given mid-band frequency, the flatness factor decreases as bandwidth increases from 0.3, since the greater bandwidth passes relatively more low frequency component which is essentially normal. When the bandwidth decreases from 0.3, its time constant increases up to a point where it is comparable with or larger than the time scale of the on-off cycle of the intermittent signal. In these cases, the filter performs some kind of weighted average on the signal and the flatness factor of the filter output decreases.* With a very narrow-band filter of bandwidth 6 Hz between -3db points, Hewlett Packard model 302A wave analyzer, it was found that the flatness factors of filtered signals were very close to 3.0 for all mid-band frequencies.

*It is generally found that any kind of "smoothing" of a non-normal stationary random variable tends to make it tend toward normality. In the limit of "infinite smoothing" this is analogous to the "Central Limit Theorem."³⁶ A more formal formulation of the "Central Limit Theorem" in this respect has been presented by J. L. Lumley in a seminar given in this Department of Mechanics on April 3, 1970.

To investigate the flatness factors of various frequency components, it is desirable to have a filter bandwidth as narrow as possible, but not so narrow that the time constant of the filter will play a smoothing role. It is seen in Figures 28 and 29 that the filter with bandwidth giving "peak" flatness factor plays a significant smoothing role. Therefore, we chose a slightly wider bandwidth, $\Delta f / f_m = 0.52$, which is the narrowest bandwidth of the Krohn-Hite filter. The flatness factors of fixed bandwidth band-passed signals are shown in Figure 30 as functions of mid-band frequency, f_m / f^* . The flatness factors start with 3.0 at low f_m and monotonically increase with f_m . The fall-off at very high f_m is due to the high noise-to-signal ratio at high frequency. The flatness factors of band-passed signals of normal (Gaussian) random noise are also shown in Figures 28 and 30; they are approximately 3.0, independent of both bandwidth and mid-band frequency.

4.3 Probability Density and Distribution Function

Let $P(e)$ be the probability density function of a random variable e ; then by definition,

$$\int_{-\infty}^{\infty} P(e) de = 1 \quad (4.4)$$

A conventional way to nondimensionalize the probability density is by transforming Equation (4.4) into

$$\int_{-\infty}^{\infty} \langle e^2 \rangle^{1/2} P(e) d\left(\frac{e}{\langle e^2 \rangle^{1/2}}\right) = \int_{-\infty}^{\infty} P_n(e_n) de_n = 1 \quad (4.5)$$

Thus, the suitable non-dimensional variables are

$$P_n(e_n) \equiv \langle e^2 \rangle^{1/2} P(e)$$

and $e_n \equiv e / \langle e^2 \rangle^{1/2}$ (4.6)

In Figures 31 and 32, the probability densities of u , $\frac{\partial u}{\partial t}$, and $\frac{\partial^2 u}{\partial t^2}$ in a grid-generated turbulence ($R_\lambda = 72$) and on the axis of a round jet ($R_\lambda = 830$) were plotted in these non-dimensional variables. The normal (Gaussian) density,

$$P_n(e_n) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{e_n^2}{2}\right), \quad (4.7)$$

was also plotted for comparison. We see that the higher the order of the derivative, the more its probability density deviates from the normal density, and that the deviation is more pronounced at higher Reynolds numbers. In all cases in which the probability densities deviate from normality, they all deviate in the same manner, tending to have higher probability than the normal curve in the neighborhood of zero e and at very large values of $e / \langle e^2 \rangle^{1/2}$, and to have lower probability at the intermediate values. This is a feature that an intermittent signal should have. If it were exactly zero during the "off" periods, there would be a Dirac function at $e = 0$.

Since the probability density was obtained by numerically differentiating the measured probability distribution, the accuracy is poor at large values of e , where the probability density is small.

It is difficult to calculate accurately the higher moments, which weigh heavily the large values of e . No report on the higher moments will be made based on the probability density data.

Since the directly measured data are the probability distributions, it is easier and more accurate to compare these with the normal distribution function. The probability distribution functions

$$\text{Prob}(e_n \leq a) \equiv \int_{-\infty}^a P_n(e_n) de_n \quad (4.7)$$

were plotted in a normal probability scale, for which the normal distribution function would appear as a straight line.

The probability distributions of u , $\partial u / \partial t$ and $\partial^2 u / \partial t^2$ at $R_\lambda = 72$ and 830 are shown in Figures 33 and 34. We see that the u 's are nearly normal, the derivatives less so. The higher order derivative deviates more from normality.

To demonstrate the variation with Reynolds number, probability distributions of $\partial^2 u / \partial t^2$ at various Reynolds numbers are shown in Figure 35. $\partial^2 u / \partial t^2$ deviates more and more from normality as Reynolds number increases.

Next the probability distribution of various frequency components (via band-pass filter) were measured in grid-generated turbulence at

$R_\lambda = 110$. They are compared with normal distribution in Figure 36, which shows that the low frequency part of the signal has a normal distribution, while the higher frequency parts do not. The deviation from normality increases as frequency increases. The probability

distribution of the signal from a high-pass Butterworth filter* is shown in Figure 37, which does not agree with normal distribution at all.

To test Kolmogorov's¹⁴ conjecture and Gurvich and Yaglom's¹⁸ prediction that a non-negative quantity governed by the fine-structure of the motion has a log-normal distribution at large Reynolds number, the probability distribution of e^2 was computed from that of e , where e is velocity fluctuations, velocity derivatives, or band-pass signals. If $P_n(e_n)$ is the probability density of e_n , the probability density and distribution of e_n^2 are respectively,

$$P'_n(e_n^2) = [P_n(-e_n) + P_n(e_n)] \frac{1}{2|e_n|} \quad (4.8)$$

and

$$\begin{aligned} \text{Prob}(e_n^2 \leq a^2) &\equiv \int_0^{a^2} P'_n(e_n^2) de_n^2 \\ &= \int_0^a [P_n(-e_n) + P_n(e_n)] de_n \quad (4.9) \\ &= \text{Prob}(e_n \leq a) - \text{Prob}(e_n \leq -a) \end{aligned}$$

where $a \geq 0$. The probability distribution

*Actually, the cascade of Butterworth filter with all the amplifiers constitutes a band-pass filter. But the high cut-off frequency is so high that the energy spectrum is negligibly small there. Therefore, the band-pass filter is essentially a high-pass filter.

$\text{Prob}(e_n^2 \leq a^2)$ was plotted on a normal probability scale against $\ln a^2$, so that log-normal distribution functions would appear as straight lines.

The probability distributions of u^2 , $(\partial u / \partial t)^2$, and $(\partial^2 u / \partial t^2)^2$ at $R_\lambda = 72$ and 830 are shown in Figures 38 and 39. The results suggest that the higher the order of derivative, the better it agrees with log-normality. $(\partial^2 u / \partial t^2)^2$ agrees well over the distribution range 0.35 to 0.9.

To look for a variation with Reynolds number, the probability distributions of $(\partial^2 u / \partial t^2)^2$ at various Reynolds numbers are shown in Figure 40. Evidently the distribution of $(\partial^2 u / \partial t^2)^2$ is approximated better by a log-normal distribution as Reynolds number increases.

The probability distributions of the squares of band-pass signals are presented in Figure 41. The results suggest that with higher mid-band frequency, the agreement with a log-normal distribution improves. The band-pass signal of mid-band frequency 6.3 kHz agrees with log-normal distribution quite well at high amplitude range. When the amplitude $e / \langle e^2 \rangle^{1/2}$ is smaller than $\frac{1}{\sqrt{33}} (\approx \sqrt{\exp[-3.5]})$ i.e., for distribution function less than 0.3, the distribution departs from log-normality. It should be pointed out that in the low amplitude range, electronic noise may contribute appreciably, so the turbulence signal may be log-normal over a broader range of values than indicated by these measurements.

The probability distribution of the square of the high-pass signal from a Butterworth filter is shown in Figure 42. This too

agrees with a log-normal distribution except the low amplitude range.

4.4 Intermittency Characteristics

The intermittency characteristics of band-pass signals were measured in grid-generated turbulence and on the axes of round jets. Their dependences on Reynolds number were investigated. The bandwidth of the Krohn-Hite filter was set at $\Delta f / f_m = 0.52$ and the mid-band frequency at $f_m = f^*$. Two quantities were measured: γ , the intermittency factor, i.e., the fraction of space occupied by the fine-structure, and \bar{n} , the average number of the fine-structure regions detected by the hot-wire per unit time. The average width $\langle W \rangle$ of the interception chord lines crossing fine-structure regions was calculated by the formula

$$\langle W \rangle = \frac{\gamma U}{\bar{n}} \quad (4.10)$$

where U is the mean velocity. The intermittency factor γ and the ratio of $\langle W \rangle$ to the Kolmogorov microscale η are shown in Figure 43 as functions of R_λ . It is seen that both γ and $\langle W \rangle / \eta$ decrease monotonically with R_λ and seem to reach asymptotic values at R_λ of the order of several hundreds. This result implies that the relative volume of the fine-structure regions decreases with increasing Reynolds number.

In the grid-generated turbulence of $R_\lambda = 110$, the intermittency characteristics of band-pass signals were measured to investigate possible dependence on frequency, hence on wave-number, in the

light of the Taylor's hypothesis. The bandwidth of the filter was set at $\Delta f/f_m = 0.52$. The intermittency factor γ and the non-dimensional width $\langle W \rangle k_m$, where $k_m = 2\pi f_m/L$ are shown in Figure 44 as a function of f_m/f^* . $\langle W \rangle k_m$ is a measure of the ratio of an average linear dimension of these "narrow-band" fine-structure regions to the characteristic eddy sizes with which they are active. Both γ and $\langle W \rangle k_m$ decrease monotonically with increasing f_m/f^* . This means that the volume occupied by the eddies with size of the order of $1/k_m$ decreases as k_m increases, which is consistent with a suggestion by Batchelor and Townsend.*

Batchelor and Townsend¹ suggested Equation (4.2) for the relationship between the flatness factor and intermittency factor of an intermittent variable. γ and F of the band-pass signal ($\Delta f/f_m = 0.52$, $f_m = f^*$) from a grid-generated turbulence ($R_\lambda = 110$) were measured to be 0.58 and 13 respectively. This differs significantly from Equation (4.2) and suggests that the intermittent fine-structure signal does not vary with a normal distribution during the time interval when it is not zero.

4.5 Geometric Categorization of the Random Fine-Structure Regions

The flow field chosen for this study is a grid-generated turbulence homogeneously strained by a slight contraction to bring it closer to isotropy.²³ The grid is of 4-inch mesh size and the mean

*see Chapter I

velocity in the test section is 12.7 m/s. Measurements were made at 42 mesh lengths downstream of the grid, where the turbulent field has been found more or less isotropic and $R_\lambda = 110$. The measured one-dimensional spectrum function of u is shown in Figure 45.

To be able to infer the geometry of the fine-structure regions, a choice of fine-structure signal has to be made. A signal which is easy to obtain and has obvious physical meaning is the time derivative of the velocity fluctuations, which is related to strain-rate and energy dissipation. Unfortunately, a turbulent flow field with Reynolds number high enough to "decouple" the dissipation spectrum from the energy spectrum is not easy to obtain in the laboratory.*

In this investigation, four-pole Butterworth high-pass filters were used to extract the fine-structure signals. The filters have sharp enough low frequency cut-off (24 db/octave) to eliminate the large-scale signals from the output.

It is assumed in the analysis of Chapter III that both the overlap of the fine-structure regions and the probability that two separate regions hit two hot-wires simultaneously are negligible. To meet these assumptions it is desirable to have the fraction of space occupied by the fine-structure as small as possible, i.e., to have a filtered signal with small intermittency factor γ . If the intermittency factor is too small, however, the measurements of coincidence functions γ_2 and Π_2 will be difficult and inaccurate. A compromise value of $\gamma = 0.2 \sim 0.4$ was chosen, and the cut-off

*see discussion in Section 4.1.

frequency of the high-pass filter was determined with the aid of the result of the intermittency measurements of band-pass signals described in Section 4.4. The energy spectrum of the filtered signal (Figure 46) has a peak at 3.5 kHz (Kolmogorov-scale frequency 5.9 kHz) which corresponds to a length scale $S_f = 2.26 \times 10^{-2}$ in.

The results of single probe measurements of the high-pass signal are

$$\gamma = 0.3 \quad \Gamma = 750/\text{sec}^* \quad \langle W \rangle = 0.2 \text{ in}$$

A summary of the analytical results of Chapter III is the following:

(1) Single-Probe Analysis (Table 1)

	$\frac{m}{\text{in}}$	γ	$\frac{\langle W \rangle}{\text{in}}$
Spherical Model	$\frac{\pi}{4} N \langle l^2 \rangle$	$\frac{\pi}{6} N \langle l^3 \rangle$	$\frac{2}{3} \frac{\langle l^3 \rangle}{\langle l^2 \rangle}$
Cylindrical Model	$\frac{\pi}{4} L \langle l \rangle$	$\frac{\pi}{4} L \langle l^2 \rangle$	$\frac{\langle l^2 \rangle}{\langle l \rangle}$
Slab Model	$\frac{1}{2} S$	$S \langle l \rangle$	$2 \langle l \rangle$

(2) Two-Probe Analysis (Table 2)

	$\frac{m_2}{m} = \frac{n_2}{n}$	$\frac{\delta_2/\gamma_1}{\text{in}}$
Spherical Model	$\left\{ \begin{array}{l} 1 - \frac{4}{9} \frac{\langle l^3 \rangle^2}{\langle l^2 \rangle^3} \frac{r_1^2}{\langle W \rangle^2} + C_1(r_1) \\ 1 - \frac{8}{3\pi} \frac{\langle l \rangle \langle l^3 \rangle}{\langle l^2 \rangle^2} \frac{r_2}{\langle W \rangle} + C_3(r_2) \end{array} \right.$	$1 - \frac{r_1}{\langle W \rangle} + \frac{1}{2} \frac{r_1^3}{\langle l^3 \rangle} + C_2(r_1)$
Cylindrical Model	$\left\{ \begin{array}{l} 1 - \frac{3}{8} \frac{\langle l^2 \rangle^2}{\langle l^3 \rangle} \langle \frac{1}{l} \rangle \frac{r_1^2}{\langle W \rangle^2} + C_4(r_1) \\ 1 - \frac{2}{\pi} \frac{\langle l^2 \rangle}{\langle l \rangle^2} \frac{r_2}{\langle W \rangle} + C_6(r_2) \end{array} \right.$	$1 - \frac{r_1}{\langle W \rangle} + C_5(r_1)$

*According to Corrsin's theorem,³² for this value of Γ the average surface area of the fine-structure regions inside a cubic inch of space is 6 in².

$$\text{Slab Model } \begin{cases} 1 - C_7(r_1) \\ 1 - C_9(r_2) \end{cases} \quad 1 - \frac{r}{\langle W \rangle} + C_8(r)$$

To compare these results with those of two-probe coincidence measurements, the restriction to $r < \langle l \rangle = o(\langle W \rangle)$ is necessary for the following reasons:

- (1) to make the higher order terms of $r/\langle W \rangle$ in $C_i(r)$ negligible,
- (2) to make terms in $C_i(r)$ involving the integration of $g(l)$ negligible; otherwise a specific functional form of $g(l)$ has to be known before the comparison can be made,
- (3) to minimize the probability that two separate fine-structure regions hit two probes simultaneously.

A reasonable assumption for the behavior of $g(l)$ at small l is

$$g(l) = 0, \quad \text{if } l \leq S_f \quad (4.11)$$

since it is necessary that the smallest dimension of fine-structure regions be at least as big as S_f , the size of eddies with which they are active. Therefore for $r \leq S_f$, those terms in $C_i(r)$ with factors Nr/N , Lr/L , or Sr/S are identically zero, and the coincidence functions may be well approximated by the first two terms of Equations in Table 2, except for the slab model, which gives 1.0.

Let us "normalize" $g(l)$ in the following form:

$$g(l) = \frac{A}{\langle l \rangle} f\left(\frac{l}{\langle l \rangle}\right), \quad (4.12)$$

where A is N , L , or S for spherical, cylindrical, or slab model, respectively. From the conditions that

$$\int_0^{\infty} g(l) dl = A \quad \text{and} \quad (4.13)$$

$$\frac{1}{A} \int_0^{\infty} l g(l) dl = \langle l \rangle$$

we have

$$\int_0^{\infty} f(x) dx = 1 \quad \text{and}$$

$$\int_0^{\infty} x f(x) dx = 1 \quad (4.14)$$

A more elaborate approximation than Equation (4.11) is to use the fact that $f(0) = 0$ and to approximate $f(x)$ with a Taylor series for small x , i.e.,

$$f(x) = \sum_{k=1}^{\infty} \frac{a_k}{k!} x^k \quad (4.15)$$

where $a_k = f^{(k)}(0)$. Equation (4.11) suggests that $a_k = 0$ for the first few k 's. For the purpose of estimating the order of magnitude of the functions $C_i(r)$, a most conservative assumption is $0 < a_1 < 1$. Then, for small l , $g(l) = \frac{a_1 A}{\langle l \rangle^2} l +$ higher order terms, and

$$\langle l^k \rangle_r = \frac{A}{A_r} \frac{a_1}{\langle l \rangle^2} \int_0^r l^{k+1} dl$$

$$\frac{A_r}{A} \langle l^k \rangle_r = \frac{a_1}{k+2} \frac{r^{k+2}}{\langle l \rangle^2}$$

Assuming $\langle l^k \rangle \langle l^m \rangle = o(\langle l \rangle^{k+m}) = o(\langle w \rangle^{k+m})$

we have

$$c_1(r_1) = o\left(\frac{r_1}{\langle w \rangle}\right)^4$$

$$c_2(r) = o\left(\frac{r}{\langle w \rangle}\right)^5$$

$$c_3(r_2) = o\left(\frac{r_2}{\langle w \rangle}\right)^3$$

$$c_4(r_1) = o\left(\frac{r_1}{\langle w \rangle}\right)^3$$

$$c_5(r) = o\left(\frac{r}{\langle w \rangle}\right)^3$$

$$c_6(r_2) = o\left(\frac{r_2}{\langle w \rangle}\right)^3 \quad (4.16)$$

$$c_7(r_1) = o\left(\frac{r_1}{\langle w \rangle}\right)^2$$

$$c_8(r) = o\left(\frac{r}{\langle w \rangle}\right)^3$$

$$c_9(r_2) = o\left(\frac{r_2}{\langle w \rangle}\right)^2$$

Substituting Equations (4.16) into equations in Table 2, we have the following for $r_1/\langle l \rangle < 1$, or $r_1/\langle w \rangle < 1$

Table 3

	$m_2/m = r_2/r_1$	γ_2/γ
Spherical Model	$\begin{cases} 1 - \frac{4}{9} \frac{\langle l^3 \rangle^2}{\langle l^2 \rangle^3} \frac{r_1^2}{\langle w \rangle^2} + o\left(\frac{r_1}{\langle w \rangle}\right)^4 \\ 1 - c_5 \frac{r_2}{\langle w \rangle} + o\left(\frac{r_2}{\langle w \rangle}\right)^3 \end{cases}$	$1 - \frac{r}{\langle w \rangle} + o\left(\frac{r}{\langle w \rangle}\right)^3$
Cylindrical Model	$\begin{cases} 1 - \frac{3}{8} \frac{\langle l^2 \rangle^2}{\langle l \rangle^3} \langle \frac{1}{l} \rangle \frac{r_1^2}{\langle w \rangle^2} + o\left(\frac{r_1}{\langle w \rangle}\right)^3 \\ 1 - c_c \frac{r_2}{\langle w \rangle} + o\left(\frac{r_2}{\langle w \rangle}\right)^3 \end{cases}$	$1 - \frac{r}{\langle w \rangle} + o\left(\frac{r}{\langle w \rangle}\right)^3$
Slab Model	$\begin{cases} 1 - o\left(\frac{r_1}{\langle w \rangle}\right)^2 \\ 1 - o\left(\frac{r_2}{\langle w \rangle}\right)^2 \end{cases}$	$1 - \frac{r}{\langle w \rangle} + o\left(\frac{r}{\langle w \rangle}\right)^3$

where

$$C_S = \frac{8}{3\pi} \frac{\langle l \rangle \langle l^3 \rangle}{\langle l^2 \rangle^2} \quad \text{and} \quad C_C = \frac{2}{\pi} \frac{\langle l^2 \rangle}{\langle l \rangle^2}$$

The results of the measurements of the coincidence functions are shown in Figures 47, 48, 49 and 50. The dependence of γ_2/γ_1 on $r_1/\langle W \rangle$ and $r_2/\langle W \rangle$ for small values of $r/\langle W \rangle$ agree well with the straight line $\gamma_2/\gamma_1 = 1 - r/\langle W \rangle$, which was estimated for all three types of geometry. Although this result does not differentiate among the three geometric categories, it is extremely encouraging for our approach because there are no adjustable constants.

The data on $\frac{\pi_2}{\pi}(r_1)$ show a quadratic departure from 1.0, which agrees qualitatively with all three types of geometry. Since the value $1 - \frac{\pi_2}{\pi}(r_1)$ is so small for small r_1 , and the coefficients of the quadratic terms in the analytical results are unknown, no distinction among the types of geometry can be made here either.

In Figure 50, the data for $\frac{\pi_2}{\pi}(r_2)$ follow the line $\frac{\pi_2}{\pi} = 1 - C \frac{r_2}{\langle W \rangle}$ for small values of r_2 . This eliminates the possibility of the fine-structure regions being slabs, since the slab model predicts a quadratic departure from 1.0. Empirically, $C \approx 0.7$.

In order to differentiate between the spherical and cylindrical models, the constants C_S and C_C in the equations of Table 3 have to be estimated. First we can show that the quantities $\frac{\langle l^3 \rangle \langle l \rangle}{\langle l^2 \rangle^2}$ and $\langle l^2 \rangle / \langle l \rangle^2$ are each greater than or equal to 1.0:

$$\therefore \langle l^k \rangle = \frac{1}{A} \int_0^\infty l^k g(l) dl$$

$$\therefore A^2 (\langle l^3 \rangle \langle l \rangle - \langle l^2 \rangle^2)$$

$$= \int_0^\infty l^3 g(l) dl \int_0^\infty l' g(l') dl' - \int_0^\infty l^2 g(l) dl \int_0^\infty l'^2 g(l') dl'$$

$$= \int_0^\infty \int_0^\infty (l^3 l' - l^2 l'^2) g(l) g(l') dl dl'$$

$$= \frac{1}{2} \int_0^\infty \int_0^\infty (l^3 l' + l l'^3 - 2 l^2 l'^2) g(l) g(l') dl dl'$$

$$= \frac{1}{2} \int_0^\infty \int_0^\infty ll' (l-l')^2 g(l) g(l') dl dl'$$

$$\geq 0 \quad \text{because } g \geq 0 \text{ and } l, l' \geq 0$$

$$\therefore \langle l^3 \rangle \langle l \rangle - \langle l^2 \rangle^2 \geq 0$$

$$\text{and } \frac{\langle l^3 \rangle \langle l \rangle}{\langle l^2 \rangle^2} \geq 1 \quad (4.17)$$

Schwarz's inequality states that

$$\frac{\langle l^2 \rangle}{\langle l \rangle^2} \geq 1 \quad (4.18)$$

Therefore $C_s \geq \frac{8}{3\pi} = 0.85$

(4.19)*

and $C_c \geq \frac{2}{\pi} = 0.64$

The data in Figure 50 are shown in Figure 51 with amplified scale. The experimental result of $C \approx 0.70$ disagrees with the behavior of the spherical model, which requires that the coincidence function fall below the straight line

$$\frac{n_2}{n} = 1 - 0.85 \frac{f_2}{\langle W \rangle} \quad (4.20)$$

On the other hand, the cylindrical model requires that the coincidence function fall below the straight line

$$\frac{n_2}{n} = 1 - 0.64 \frac{f_2}{\langle W \rangle} \quad (4.21)$$

*If we assume a functional form for $g(l)$ such that $f(x)$ in Equation (4.12) has the form

$$f(x) = a x^K \exp(-bx)$$

then

$$C_s = 0.85 \frac{K+3}{K+2}$$

$$C_c = 0.64 \frac{K+2}{K+1}$$

To fit the experimental result of $C = 0.70$, a value of $K = 9$ is required.

If we assume $f(x) = a x^K \exp(-bx^2)$

then $C_c = 0.70$ and $C_s = 0.94$ for $K = 4$.

We see that the measured coincidence rates are consistent with the random cylinder model. Therefore the statistical geometry of the fine-structure regions is more likely to be "rods" distributed randomly in the flow field than slabs or blobs.

Needless to say, the possibility of a mixture of rods and blobs, and the possibility of "elongated blobs," remain. For elaborate measurements will be required to decide among these alternatives.

V. DISCUSSION AND CONCLUSIONS

This investigation shows that in a fully developed turbulent flow, the fine-scale components of velocity localize in relatively small regions which are distributed randomly throughout the flow field, while the large-scale components cover the entire flow field. Two-probe coincidence measurements for the occurrence of fine-structure suggest that the fine-structure regions are likely to have random "rod-like" geometry, as contrasted to two other categories: "blob-like" and "slab-like." If we divide the flow field into two kinds of regions, one with fine-structure and the other without, then Figure 54 may represent a qualitative picture of the flow field. The "diameters" of the rods are large compared with the size of eddies of which they are composed, and is smaller for smaller eddies. The mean "rod separation" is fairly large compared with the rod thickness. The radius of curvature of the rod axis is presumably large compared with rod diameter, otherwise the rods would "look like" blobs.

Since we divide the flow field into two kinds of regions only, the geometry of the fine-structure regions must be related to that of the "no-fine-structure" regions. In fact, we could have done intermittency and coincidence measurements using absence of fine-structure as the significant event. Suppose that, instead of $I(t)$, we had obtained from the intermittency circuit a signal $I'(t)$ such that

$$I'(t) = \begin{cases} 1.0 & \text{when the hot-wire is in the} \\ & \text{no-fine-structure regions,} \\ 0 & \text{otherwise} \end{cases}$$

We could have defined and measured the statistical quantities γ' , η' and the coincidence functions γ_2' , η_2' for $I'(t)$ the same way as γ , η , γ_2 , η_2 for $I(t)$. It can be shown that γ' , η' , γ_2' , and η_2' are determined uniquely by γ , η , γ_2 , and η_2 .

(1) By definition, we have

$$\begin{aligned} I'(t) &= 1 - I(t) \\ \gamma &= \overline{I(t)} \\ \gamma' &= \overline{I'(t)} \quad \eta' = \eta \end{aligned}$$

$$\therefore \gamma' = \overline{1 - I(t)} = 1 - \gamma \quad (5.1)$$

Let $I_1'(t)$ and $I_2'(t)$ be the signals from two hot-wires separated by a distance r . Then

$$\begin{aligned} \gamma_2' &= \overline{I_1'(t) I_2'(t)} \\ &= \overline{[1 - I_1(t)][1 - I_2(t)]} \\ &= 1 - 2\gamma + \gamma_2 \end{aligned}$$

$$\therefore \frac{\gamma_2'}{\gamma'} = 1 - \frac{\gamma}{1-\gamma} \left(1 - \frac{\gamma_2}{\gamma}\right) \quad (5.2)$$

For the experimental results on $\delta_2/\gamma(r)$, it is convenient to define a function $h(\frac{r}{\langle W \rangle})$ such that

$$\frac{\delta_2}{\gamma} \equiv 1 - h\left(\frac{r}{\langle W \rangle}\right) \quad (5.3)$$

and if we define the mean "pulse length" as

$$\langle W' \rangle \equiv \frac{\gamma' \eta}{\eta'} = \frac{(1-\gamma) \eta}{\eta} = \frac{1-\gamma}{\gamma} \langle W \rangle$$

then, by substituting (5.3) into (5.2) and changing the variable,

we have

$$\begin{aligned} \frac{\gamma_2'}{\gamma_1'} \left(\frac{r}{\langle W' \rangle} \right) &= 1 - \frac{\gamma}{1-\gamma} h \left(\frac{\langle W' \rangle}{\langle W \rangle} \frac{r}{\langle W' \rangle} \right) \\ &= 1 - \frac{\gamma}{1-\gamma} h \left(\frac{1-\gamma}{\gamma} \frac{r}{\langle W' \rangle} \right) \end{aligned} \quad (5.4)$$

The experimental results in Figures 47 and 49 suggest that, for $r \ll \langle W \rangle$,

$$h \left(\frac{r}{\langle W \rangle} \right) \approx K \frac{r}{\langle W \rangle}$$

with $K \approx 1.0$.

Therefore, from Equation (5.4), we have

$$\frac{\gamma_2'}{\gamma_1'} \approx 1 - \frac{r}{\langle W' \rangle} \quad \text{for } r \ll \langle W' \rangle \quad (5.5)$$

(2) η_2' is the average number of pulses per unit time of the signal

$$I_1'(t) I_2'(t) = 1 - I_1(t) - I_2(t) + I_1(t) I_2(t)$$

We may determine η_2' by counting the rate of occurrence (per unit time) of time intervals for which

$$I_1'(t) I_2'(t) = 0$$

Whenever $I_1(t)$ is 1.0, $I_1'(t) I_2'(t)$ is 0, regardless of the value of $I_2(t)$; this contributes number η to η_2' , since η is the number of time intervals in which $I_1(t)$ is 1.0. When $I_2(t)$ is 1.0, it will add new time intervals to the event $I_1'(t) I_2'(t) = 0$ only if $I_1(t) = 0$ simultaneously. This contributes a number $\eta - \eta_2$

to n_2' .

$$\therefore n_2' = 2n - n_2$$

$$\frac{n_2'}{n'} = 2 - \frac{n_2}{n} \quad (5.6)$$

For the experimental results on $\frac{n_2}{n}(r)$ it is convenient to define a function $h'(\frac{r}{\langle W \rangle})$ such that

$$\frac{n_2}{n}(\frac{r}{\langle W \rangle}) \equiv 1 - h'(\frac{r}{\langle W \rangle}) \quad (5.7)$$

Then, by substituting (5.7) into (5.6), and changing the variable, we have

$$\frac{n_2'}{n'}(\frac{r}{\langle W' \rangle}) = 1 + h'[\frac{1-\gamma}{\gamma} \frac{r}{\langle W' \rangle}] \quad (5.8)$$

The experimental results in Figures 48 and 50 suggest that, for

$$r \ll \langle W \rangle \quad h'(\frac{r}{\langle W \rangle}) = c' \frac{r^2}{\langle W \rangle^2}, \quad c' > 0$$

$$h'(\frac{r}{\langle W \rangle}) = c \frac{r}{\langle W \rangle}, \quad c \approx 0.70$$

Therefore, from Equation (5.8), we have

$$\frac{n_2'}{n'}(\frac{r}{\langle W' \rangle}) = \begin{cases} 1 + c' (\frac{1-\gamma}{\gamma})^2 \frac{r^2}{\langle W' \rangle^2} \\ 1 + c (\frac{1-\gamma}{\gamma}) \frac{r}{\langle W' \rangle} \end{cases} \quad (5.9)$$

Equations (5.5) and (5.9) could have described the experimental results for small r if two-probe coincidence measurements for the occurrence of no-fine-structure regions had been performed.

Equation (5.5) agrees with the analytical results of Chapter III for all three geometric categories [Equations (3.15), (3.20), (3.23)] , so no distinction can be drawn from this. Equation (5.9), on the other hand, is different from the analytical results for all three geometric categories.

At least two possible conclusions may be drawn: (1) the no-fine-structure regions cannot be described by any of the three geometric categories assumed in Chapter III; we can, however, infer that the fine-structure regions cannot be slabs, since that would imply that the no-fine-structure regions are blobs, whose prediction disagrees with Equation (5.9). (2) the volume fraction of the no-fine-structure regions is too large (0.7 in this experiment) for the analysis of Chapter III to be applicable.

The comparison of experimental and analytical results in Section 4.5 shows clearly that the fine-structure regions are not slabs. But the distinction between cylindrical and spherical models is not so clear cut. In Figure 51, the first quadrant of the $\frac{n_2}{n} - \frac{r_2}{\langle W \rangle}$ plane may be divided into three regions by the two straight lines

$$(a) \quad \frac{n_2}{n} = 1 - 0.64 \frac{r_2}{\langle W \rangle}$$

$$(b) \quad \frac{n_2}{n} = 1 - 0.85 \frac{r_2}{\langle W \rangle}$$

Line (a) is the upper limit for the cylindrical model. This upper limit occurs if all the cylinders have the same diameter, i.e.,

$g(l)$ is a Dirac function. Line (b) is the upper limit for the spherical model. This upper limit occurs if all the spheres have

the same diameter. If the η_2/η data for small r_2 fell above line (a), neither cylindrical nor spherical model would be correct. If the data fell below line (b), both the cylindrical and spherical models would be possible. The data actually fall in the region between lines (a) and (b), in which among the three geometric models only the cylindrical one is possible. But the cylindrical geometry alone does not predict the exact location of the data. The probability function $g(\ell)$ of the cylinder diameters determines the coefficient C_c in the equation of Table 2.

If the probability function $g(\ell)$ were known, then we could check the form of the coincidence function in detail. But we still have neither an experimental procedure for measuring it nor a theoretical basis for estimating it.

With $g(\ell)$ unknown, the experimental results also allow the possibility that the fine-structure regions are a mixture of rods and blobs or have a geometry between rod and blob, i.e., an elongated blob or finite length cylinder.

If the fine-structure regions are a mixture of spheres and cylinders, we can write the intermittency factor and the pulse rate of the signal $I(t)$ as

$$\gamma = \gamma_s + \gamma_c \quad (5.10)$$

$$\text{and} \quad \eta = \eta_s + \eta_c \quad (5.11)$$

where γ_s and γ_c are the fractions of time the detector is in a sphere and in a cylinder, respectively, η_s and η_c are the average

numbers per unit time, of spheres and cylinders hitting the hot-wire.

We may also define the average intersection chord lengths with spheres, with cylinders, and with the mixture as

$$\langle W_s \rangle = \frac{\cup \gamma_s}{n_s}$$

$$\langle W_c \rangle = \frac{\cup \gamma_c}{n_c}$$

$$\langle W \rangle = \frac{\cup \gamma}{n} = \frac{\cup (\gamma_s + \gamma_c)}{n_s + n_c}$$

If we place two detectors in the flow field and neglect the occasions on which they are in two different fine-structure domains, we can write

$$\frac{\gamma_2}{\gamma} = \frac{\gamma_{2s} + \gamma_{2c}}{\gamma_s + \gamma_c} \quad (5.12)$$

At $r \ll \langle W \rangle$, Equations (3.15) and (3.20) reduce to

$$\frac{\gamma_{2s}}{\gamma_s} = 1 - \frac{r}{\langle W_s \rangle} \quad (5.13)$$

$$\frac{\gamma_{2c}}{\gamma_c} = 1 - \frac{r}{\langle W_c \rangle} \quad (5.14)$$

Substituting (5.13) and (5.14) into (5.12), we have

$$\frac{\gamma_2}{\gamma} = 1 - \frac{r}{\langle W \rangle} \quad (5.15)$$

which agrees with experimental results regardless of the ratios γ_s/γ_c and n_s/n_c , the sphere-to-cylinder volume ratio and number ratio, respectively.

For the average number of fine-structure regions, per unit time, hitting both detectors simultaneously, we have

$$\frac{n_2}{n} = \frac{n_{2s} + n_{2c}}{n_s + n_c} \quad (5.16)$$

At $r_2 \ll \langle W \rangle$, Equations (3.17) and (3.21) reduce to

$$\frac{n_{2s}}{n_s} = 1 - C_s \frac{r_2}{\langle W_s \rangle} ; \quad C_s = \frac{8}{3\pi} \frac{\langle l_s \rangle \langle l_s^3 \rangle}{\langle l_s^2 \rangle^2} \quad (5.17)$$

$$\frac{n_{2c}}{n_c} = 1 - C_c \frac{r_2}{\langle W \rangle} ; \quad C_c = \frac{2}{\pi} \frac{\langle l_c^2 \rangle}{\langle l_c \rangle^2} \quad (5.18)$$

Substituting (5.17) and (5.18) into (5.16), we have

$$\frac{n_2}{n} = 1 - C_M \frac{r_2}{\langle W \rangle} \quad (5.19)$$

where

$$C_M = \frac{\gamma \left(\frac{C_s}{\gamma_s} n_s^2 + \frac{C_c}{\gamma_c} n_c^2 \right)}{n^2} \quad (5.20)$$

Write

$$\alpha = \gamma_s / \gamma_c, \quad \beta = n_s / n_c, \quad \rho = \frac{\alpha}{\beta}$$

then

$$C_M = \frac{C_c + \beta \left(C_c \rho + \frac{C_s}{\rho} \right) + C_s \beta^2}{(1 + \beta)^2} \quad (5.21)$$

From Equations (3.3) and (3.7)

$$\alpha = \frac{2 N \langle l_s^3 \rangle}{3 L \langle l_c^2 \rangle}$$

and from Equations (3.2) and (3.6)

$$\beta = \frac{N \langle l_s^2 \rangle}{L \langle l_c \rangle} \quad (5.22)$$

$$\therefore f = \frac{2}{3} \frac{\langle l_s^3 \rangle \langle l_c \rangle}{\langle l_s^2 \rangle \langle l_c^2 \rangle}$$

The coefficients C_s and C_c in Equations (5.17) and (5.18) will attain their minimum values

$$(C_s)_m = 0.85$$

$$(C_c)_m = 0.64$$

if all the spheres are a single size and all the cylinders are a (possibly different) single size, i.e.,

$$l_s = \text{constant}$$

$$l_c = \text{constant}$$

For this case, Equation (5.21) gives

$$C_M = 0.64 \frac{1 + (f + \frac{1.33}{f})\beta + 1.33\beta^2}{(1 + \beta)^2}$$

and, for given β , C_M has a minimum value $(C_M)_m$

$$(C_M)_m = 0.64 \frac{1 + 2.3\beta + 1.33\beta^2}{(1 + \beta)^2} \quad (5.23)$$

$$\text{when } f = 1.15, \quad \text{i.e. } \frac{2}{3} \frac{l_s}{l_c} = 1.15 \quad (5.24)$$

Therefore, restricting to $r_2 \ll \langle W \rangle$, the straight line

$$\frac{n_2}{n} = 1 - (C_M)_m \frac{r_2}{\langle W \rangle} \quad (5.25)$$

is the upper limit for the "mixture" model. $(C_M)_m$ is a monotonically increasing function of β , the number ratio.

Since the experimental data lie on a straight line

$$\frac{n_2}{n} = 1 - 0.7 \frac{r_2}{\langle W \rangle} \quad \text{for } r_2 \ll \langle W \rangle$$

the maximum permissible β will be the one such that $(C_M)_m = 0.7$, then

$$\beta_{\max} = 0.41^* \text{ and}$$

$$\alpha_{\max} = f \beta_{\max} = 0.47$$

That is, the experimental data allow a maximum sphere volume of

$$\frac{0.47}{1+0.47} = 32\% \text{ of the total volume of the fine-structure regions.}$$

This maximum volume ratio occurs when, from Equation (5.24)

$$l_s = 1.73 l_c = \text{constant.} \quad (5.26)$$

If we make the more realistic assumptions that the diameters are distributed over some range, C_M will be greater than $(C_M)_m$ for the same value of β . Therefore the volume fraction of spheres has to be smaller than 32% in order to agree with the experimental results.

*The other solution for β is -0.91, which is physically impossible.

For example, the case that

$$l_s = l_c = \text{constant}$$

gives a volume ratio of 11%.

If we assume that the fine-structure regions are elongated blobs or finite length cylinders, we can expect that the upper limit of the coincidence function $\frac{n_2}{n_1}(r_2)$ (which happens when all the cylinders have the same diameter) will be a straight line between lines (a) and (b) in Figure 51. No analysis has been carried out for this model, but this line of upper limit will tend toward line (b) as the length-to-diameter ratio of the cylinder decreases, and will coincide with line (b) when the length-to-diameter ratio is of the order of unity. Since the experimental data lie closer to line (a) than line (b), it is expected that the length-to-diameter ratio would be at least appreciably greater than one for the special case that all the cylinders are of the same diameter. If the cylinder diameters distribute over some range, the length-to-diameter ratio will be even higher.

This experimental investigation of the statistical geometry of fine-structure regions has provided only a starting point for future work. A more detailed investigation by coincidence measurements with three or even more probes would be interesting. It would also be interesting to detect the geometry in a turbulent shear region, where the rod-like geometry would have a preferred direction.

The fine-scale components of temperature fluctuations have been observed to be intermittent in the wind over ocean.⁶ Studies of the

statistical geometry of the fine-structure regions of such scalar fields are certainly worthwhile. Efforts might be made to correlate the fine-structure intermittency between scalar field and velocity field.

Finally, a theoretical attempt might be tried to relate the geometry by seeking the dominant signs* of the short-time-average principal strain rates following a fluid element in which the fine-structure exists. Betchov³⁷ has discussed the dominant signs of the average principal rates in an isotropic turbulent flow. But his results cannot be applied here since he talked about (ensemble) average fixed in space and time while the average we need is the "short" time average following a particular fine-structure region.

*see discussion at the beginning of Chapter III.

APPENDIX

A1. On the Application of Taylor's Hypothesis

In this experiment, Taylor's hypothesis³⁷ (or "approximation") has been applied on two occasions: (a) interpretation of hot-wire signal high frequency spectrum as turbulence high wave-number spectrum, i.e., fine-structure signal, (b) interpretation of a time-delayed signal as though it were the non-delayed signal from a hot-wire at a location directly downstream of the point of interest.

Taylor's hypothesis is based on the assumption that the spatial pattern of turbulent motion is convected past a fixed point by the mean flow without any essential change. Considering application to the fine-scale structure, we must consider not only the "self-convection" effect, but also the fact that the small eddies are carried about "within" the energy-containing (large) eddies, and thus are convected by the fluctuating velocity field of large-scale structure as well as by the mean velocity field. If u' and L are the velocity and length scales of the energy-containing eddies, the unsteadiness of convection velocity will be of the order of u' and the non-uniformity of the convection velocity across a small eddy of size $\frac{1}{k}$ will be $\frac{u'}{Lk}$.*

*In a mean-shear region there is, of course, a non-uniformity of convection velocity due to the mean velocity gradient, viz.,
 $\frac{1}{k} \frac{\partial U}{\partial y}$, but the data reported here are recorded at positions where $\frac{\partial U}{\partial y} = 0$.

In a grid-generated turbulence, the mean velocity U is uniform and u'/U (≈ 0.02) is small, hence the unsteadiness and non-uniformity of convection velocity are negligible; this justifies the interpretation of frequency spectrum as wave-number spectrum. Good agreements between theory and experiment in grid-generated turbulence have been reported by Taylor³⁷ and others.

The first extensive and good experimental check of Taylor's hypothesis was made by Farre, Gaviglio, and Dumas³⁸ in a turbulent boundary layer.

The applicability of (a) at large turbulence levels as on the axis of a round jet or in other shear flows has been discussed by Lin,³⁹ Fisher and Davies,⁴⁰ and Lumley.⁴¹ A correction form for the wave-number spectrum has been suggested by Lumley.

Application (b) of Taylor's hypothesis permits us to interpret a time varying signal from a hot-wire fixed in space as an instantaneous function of spatial coordinate along the path of mean flow. The intermittency of a high frequency signal in the time domain may be interpreted as an indication of localization of the fine-structure in the space domain. The time intervals when the high frequency signal is zero correspond to the hot-wire sweeping through spatial regions in which the fine-scale components of motion do not exist.

In the two-probe coincidence measurements for the occurrence of the fine-structure, the flow disturbance due to one hot-wire prohibits the placement of the second one in its wake. The Taylor's hypothesis, however, allows us to measure the coincidence functions

$\frac{\gamma_2}{\gamma_1}(r_1)$ and $\frac{n_2}{n_1}(r_1)$ with a single wire plus a signal delay line. The delayed signal from the hot-wire is interpreted as the signal from a fictitious hot-wire at a location directly downstream, with separation

$$r_1 = U \tau \quad (\text{A1.1})$$

where U is the mean velocity and τ is the delay time.

If t_s is the time scale of the fine-structure whose presence we seek, then the approximation is good only for $\tau \ll t_s$. In the investigation of the geometry of fine-structure regions, the fine-structure signal is a high-pass signal from a grid-generated turbulence with $R_\lambda = 110$. This high-pass signal has a spectrum peaked at 3.5 kHz (see Figure 46), which corresponds to wave-number $k_1 = 17.3/\text{cm}$. The eddies of this wave-number have a characteristic length $1/k_1$ and a characteristic velocity $(k_1 E_{11})^{1/2}$, where $E_{11}(k_1)$ is the one-dimensional energy spectrum. Therefore, a characteristic inertial time scale (essentially the Onsager time scale), is

$$t_s = [k_1^3 E_{11}(k_1)]^{-1/2} \quad (\text{A1.2})^*$$

which is about 84 ms at $k_1 = 17.3/\text{cm}$. A more conservative estimate of time scale is the Kolmogorov time scale, actually the order

*Strictly speaking, the "three-dimensional spectrum" and the wave-number magnitude $|k|$ are more appropriate here, but their values are not available here.

of magnitude of r.m.s. turbulent strain rate and vorticity:

$$t_k = \left(\frac{\langle \epsilon \rangle}{\nu} \right)^{-1/2} \quad (A1.3)$$

which is 7.9 ms for the grid-generated turbulence with $R_\lambda = 110$.

As shown in Figures 47 and 48, both $\sigma_{2/\gamma}(\tau_1)$ and $\tau_{2/\eta}(\tau_1)$ approach asymptotic values within $\tau_1/\langle w \rangle \leq 4$, which corresponds to $\tau = 1.6$ ms, only one-fifth of the smaller time scale, t_k . This comparison of t_k with τ assures that the simulation of two hot-wires separated in the mean flow direction by a single hot-wire plus its delayed signal is satisfactory, at least for the range $\tau_1/\langle w \rangle < 1$, where the data are compared with analytical results.

In fact, the signal which was delayed is the on-off binary signal which distinguishes the fine-structure regions from the rest of the fluid. As long as the shape change of the fine-structure domains is small during the time interval τ , the delayed signal is still a good approximation, even if the fine-scale components inside the domains change appreciably. This probably means that $\tau \ll t_s$ is sufficient.

As an experimental check on the validity of application (b), the space-time correlation of high-pass signals and of on-off binary signals from two hot-wires separated in both τ_1 and τ_2 directions were measured. The τ_2 -separation keeps the downstream wire out of the wake of the upstream one. For two hot-wires fixed in

space, the correlation was measured as a function of τ , the delay time of the signal from the upstream wire. τ_m was plotted as function of r_i in Figures 52 and 53, where τ_m is the delay time at which the correlation is maximum. The figures show that the data agree quite well with a straight line of slope 0.5 in/ms, which is the mean speed of the flow field. This is in agreement with Taylor's hypothesis.

A2. Some Properties of Log-Normal Probability Distribution

In order to take into account the variation of turbulent energy dissipation rate, Oboukhov¹³ and Kolmogorov¹⁴ proposed a modified version of the original universal similarity hypothesis.^{8,9} The modification involved the assumption that the logarithm of the energy dissipation rate has a normal distribution. Gurvich and Yaglom¹⁹ presented a theoretical treatment leading to the prediction that any non-negative quantity (e.g., energy dissipation rate) governed by fine-scale components has a log-normal probability distribution.

The prediction of log-normality has been tested experimentally by several investigators,^{19,5,42,44} but no unique way of curve fitting has been used up to now. Furthermore, no detailed discussion of the behavior of a log-normal density function has been given. In this section, the method used in Section 4.3 to try to fit the data with log-normal distribution functions is discussed in detail, and some moments of the log-normal density are compared with direct measurements. The behavior of a log-normal density function is dis-

cussed, in particular, its implication for intermittency.

By log-normal probability distribution, it is meant that the logarithm of a non-negative random variable ε has a normal probability distribution, i.e.,

$$P_{\ln \varepsilon}(y) = \frac{1}{\sqrt{2\pi} \beta} \exp \left[- \frac{(y-m)^2}{2\beta^2} \right] \quad (\text{A2.1})$$

where $m = \langle \ln \varepsilon \rangle$ and

$$\beta^2 = \langle (\ln \varepsilon - m)^2 \rangle$$

We can compute arbitrary moments:

Write $y = \ln \varepsilon$; then $\frac{dy}{d\varepsilon} = \frac{1}{\varepsilon}$ and

$$P(\varepsilon) = P_{\ln \varepsilon}(y) \frac{dy}{d\varepsilon} \quad (\text{A2.2})$$

$$= \frac{1}{\sqrt{2\pi} \beta \varepsilon} \exp \left[- \frac{(\ln \varepsilon - m)^2}{2\beta^2} \right]$$

The K^{th} moment of ε

$$\langle \varepsilon^K \rangle = \int_0^{\infty} \frac{1}{\sqrt{2\pi} \beta \varepsilon} \varepsilon^K \exp \left[- \frac{(\ln \varepsilon - m)^2}{2\beta^2} \right] d\varepsilon \quad (\text{A2.3})$$

$$= \exp \left(mk + \frac{1}{2} k^2 \beta^2 \right)$$

In particular,

$$\langle \varepsilon \rangle = \exp \left(m + \frac{1}{2} \beta^2 \right)$$

Write $\varepsilon_n = \varepsilon / \langle \varepsilon \rangle$;

then the p.d.f. of this normalized variable is

$$\begin{aligned}
 P_n(\varepsilon_n) &\equiv \langle \varepsilon \rangle P(\varepsilon_n < \varepsilon) \\
 &= \frac{1}{\sqrt{2\pi} \beta \varepsilon_n} \exp \left[- \frac{(\ln \varepsilon_n + \frac{\beta^2}{2})^2}{2\beta^2} \right] \quad (A2.4)
 \end{aligned}$$

The distribution function is

$$\begin{aligned}
 \text{Prob}(\varepsilon_n \leq a^2) &= \int_0^{a^2} \frac{1}{\sqrt{2\pi} \beta \varepsilon_n} \exp \left[- \frac{(\ln \varepsilon_n + \frac{\beta^2}{2})^2}{2\beta^2} \right] d\varepsilon_n \\
 &= \int_{-\infty}^{\ln a^2} \frac{1}{\sqrt{2\pi} \beta} \exp \left[- \frac{(y_n + \frac{\beta^2}{2})^2}{2\beta^2} \right] dy_n \quad (A2.5)
 \end{aligned}$$

where

$$y_n = \ln \varepsilon_n$$

If Equation (A2.5) is plotted on normal probability paper with $\ln a^2$ as amplitude variable, we will have a family of straight lines because we have a normal distribution for $\ln \varepsilon_n$, with parameter β . This normal distribution has a mean of $-\beta^2/2$ and a variance of β . A way to see how closely a set of probability distribution data may be approximated by a log-normal distribution is to try to fit these data with a straight line on a normal probability paper with amplitude plotted in a logarithmic scale. There are two parameters to be adjusted in choosing the best fitted straight lines:

- (a) β , which determines the slope of the straight line, if there is a straight line;

$$2\beta = (\ln a_{0.84}^2) - (\ln a_{0.16}^2)$$

where $a_{0.16}^2$ and $a_{0.84}^2$ are the values such that

$$\text{Prob.}(\epsilon_n \leq a_{0.16}^2) = 0.16$$

$$\text{Prob.}(\epsilon_n \leq a_{0.84}^2) = 0.84$$

(b) $\langle \epsilon \rangle$, which determines the location of the straight line.

Different choices of $\langle \epsilon \rangle$ only add a constant to $\ln \epsilon_n$ and shift the straight line parallel to itself.

In the actual curve fitting for Chapter IV, β was chosen so that the straight line would fit as many data points as possible. The mean $\langle e^2 \rangle$ of the non-negative random variable was used as $\langle \epsilon \rangle$ to non-dimensionalize the variable e^2 for convenience; no attempt was made to determine the best $\langle \epsilon \rangle$, since different choices of $\langle \epsilon \rangle$ only translate the straight line horizontally. If the data were truly log-normally distributed, then

$$\langle e^2 \rangle = \langle \epsilon \rangle$$

and $\langle e^2 \rangle$ would be the best choice for $\langle \epsilon \rangle$; the straight line would be one of the family with $\ln \epsilon_n$ having a mean of $-\frac{\beta^2}{2}$ and a variance of β .

In Figure 42, the straight line corresponds to a log-normal distribution with $\beta = 2.1$. Then, according to Equation (A2.5),

$$\ln \frac{e_m^2}{\langle \epsilon \rangle} = \ln a_{0.5}^2 = -\beta^2/2 = -2.2$$

where e_m^2 is the median of e^2 . But the high-pass signal e^2 is non-dimensionalized by $\langle e^2 \rangle$ in Figure 42, and the straight line has

$$\ln \frac{e_m^2}{\langle e^2 \rangle} = \ln a_{0.5}^2 = -0.16$$

$$\therefore \ln \frac{\langle \mathcal{E} \rangle}{\langle e^2 \rangle} = 0.6$$

(A2.6)

$$\text{or } \langle \mathcal{E} \rangle = 1.82 \langle e^2 \rangle$$

Similarly, the probability distribution of the band-pass signal with $\Delta f/f_m = 0.52$ and $f_m = 6.3$ kHz in Figure 41 shows

$$\langle \mathcal{E} \rangle = 1.68 \langle e^2 \rangle \quad (\text{A2.7})$$

We observe that both of these ratios $\langle \mathcal{E} \rangle / \langle e^2 \rangle$ depart appreciably from 1.0, the value required for a log-normal distribution.

Furthermore, for the probability distribution of a log-normal random variable e^2 , then the flatness factor of e would be

$$\begin{aligned} F &\equiv \frac{\langle e^4 \rangle}{\langle e^2 \rangle^2} = \frac{\langle \mathcal{E}^2 \rangle}{\langle \mathcal{E} \rangle^2} \\ &= \frac{\exp(2m + 2\beta^2)}{\exp(2m + \beta^2)} = \exp(\beta^2) \end{aligned} \quad (\text{A2.8})$$

The band-pass signal data in Figure 41 would give F equal to about 80 ~ 100 ($f_m = 6.3$ kHz). The direct measurement of the flatness factor of this band-pass signal gave a value of 16.*

*A similar comparison between the moments computed from log-normal density and direct measurements was published⁴³ after this Appendix was written.

The fact that the log-normal distribution predicts much higher values for the second and fourth moments of a high frequency signal $e(t)$ cannot be explained by the departure from the log-normal form at small values of e^2 . Since the higher order moments weigh more heavily on the large amplitude range, the figures and the numbers suggest that the large values of e^2 are less probable than would be predicted by a log-normal distribution. This deviation at large amplitude was also reported by Stewart, Wilson, and Burling.^{42,43}

To see how a log-normally distributed random variable may appear intermittent, some features of the probability density function (A2.4) have to be examined. The peak may be located by setting

$$\frac{dP_n(\mathcal{E}_n)}{d\mathcal{E}_n} = 0 \quad (\text{A2.9})$$

Substituting Equation (A2.4), we have

$$(\mathcal{E}_n)_{\max.} = \exp\left(-\frac{3}{2}\beta^2\right) \quad (\text{A2.10})$$

and

$$P_n[(\mathcal{E}_n)_{\max.}] = \frac{1}{\sqrt{2\pi}} \beta \exp(\beta^2) \quad (\text{A2.11})$$

where $(\mathcal{E}_n)_{\max}$ is the location at which $P_n(\mathcal{E})$ is maximum. Equations (A2.10) and (A2.11) indicate that the peak of the probability density function will increase and its \mathcal{E} -location will move toward zero as β increases.

A more detailed examination of the variation of $P_n(\mathcal{E}_n)$ with

β may be achieved by checking the sign of $\frac{\partial P_n}{\partial \beta}$. From Equation (A2.4)

$$\frac{\partial P_n}{\partial \beta} = \frac{1}{\sqrt{2\pi}} \frac{1}{\epsilon_n} \exp \left[-\frac{(\ln \epsilon_n + \frac{\beta^2}{2})^2}{2\beta^2} \right] \left[-\frac{1}{\beta^2} - \frac{1}{4} + \frac{(\ln \epsilon_n)^2}{\beta^4} \right]$$

$$\therefore \frac{\partial P_n}{\partial \beta} > 0 \quad \text{if}$$

$$(\ln \epsilon_n)^2 > \frac{\beta^4}{4} + \beta^2, \quad \text{i.e. if}$$

$$\text{and } \frac{\partial P_n}{\partial \beta} < 0 \quad \begin{cases} \epsilon_n > \exp(\sqrt{\beta^4/4 + \beta^2}) \quad \text{or} \\ \epsilon_n < \exp(-\sqrt{\beta^4/4 + \beta^2}) \end{cases}$$

$$\text{if } \exp(-\sqrt{\beta^4/4 + \beta^2}) < \epsilon_n < \exp(\sqrt{\beta^4/4 + \beta^2})$$

Therefore, as β increases, the probability density at very large and very small values of ϵ_n will increase and that at intermediate values of ϵ_n will decrease. Some log-normal probability density curves with various β are shown in Figure 55 and they suggest that the random variable will appear intermittent if β is large enough.

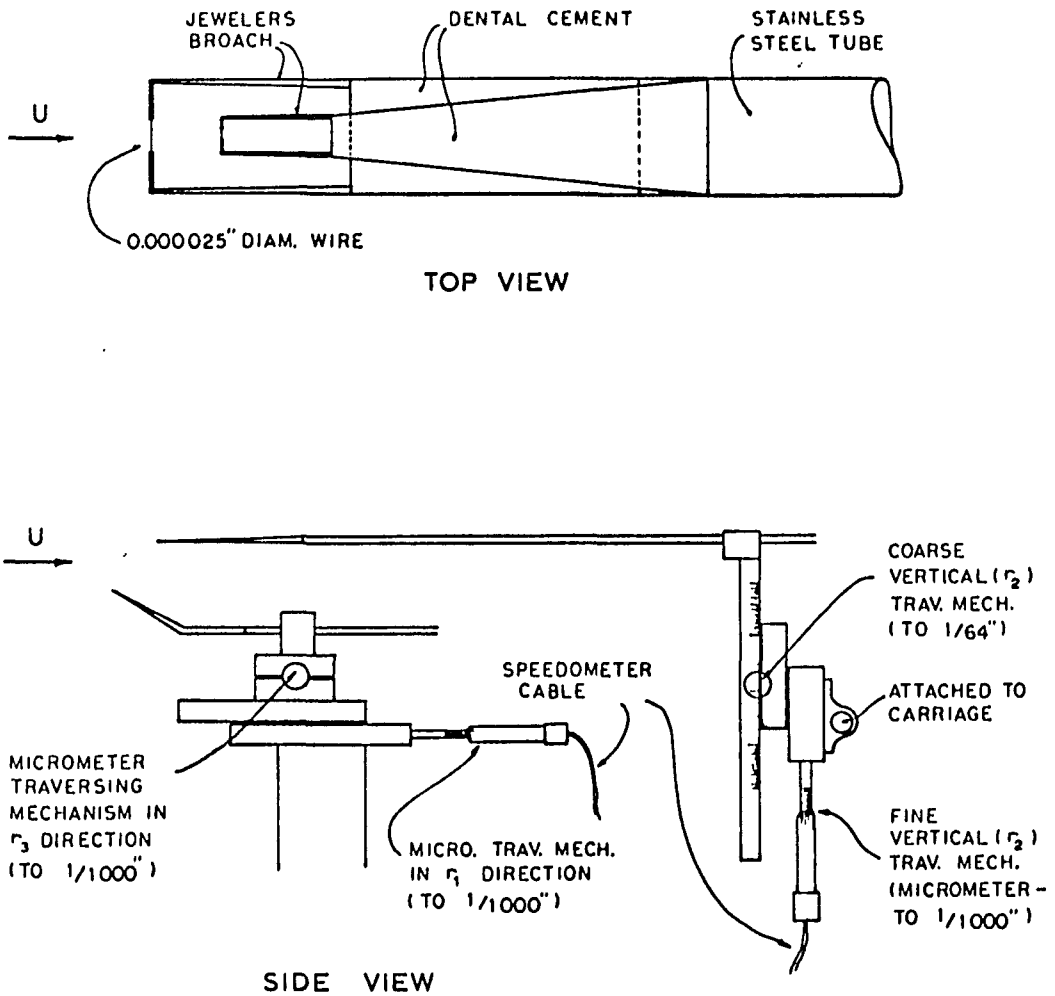


Figure 1. Sketch of Hot-wire Probe Configuration and Traversing Mechanism

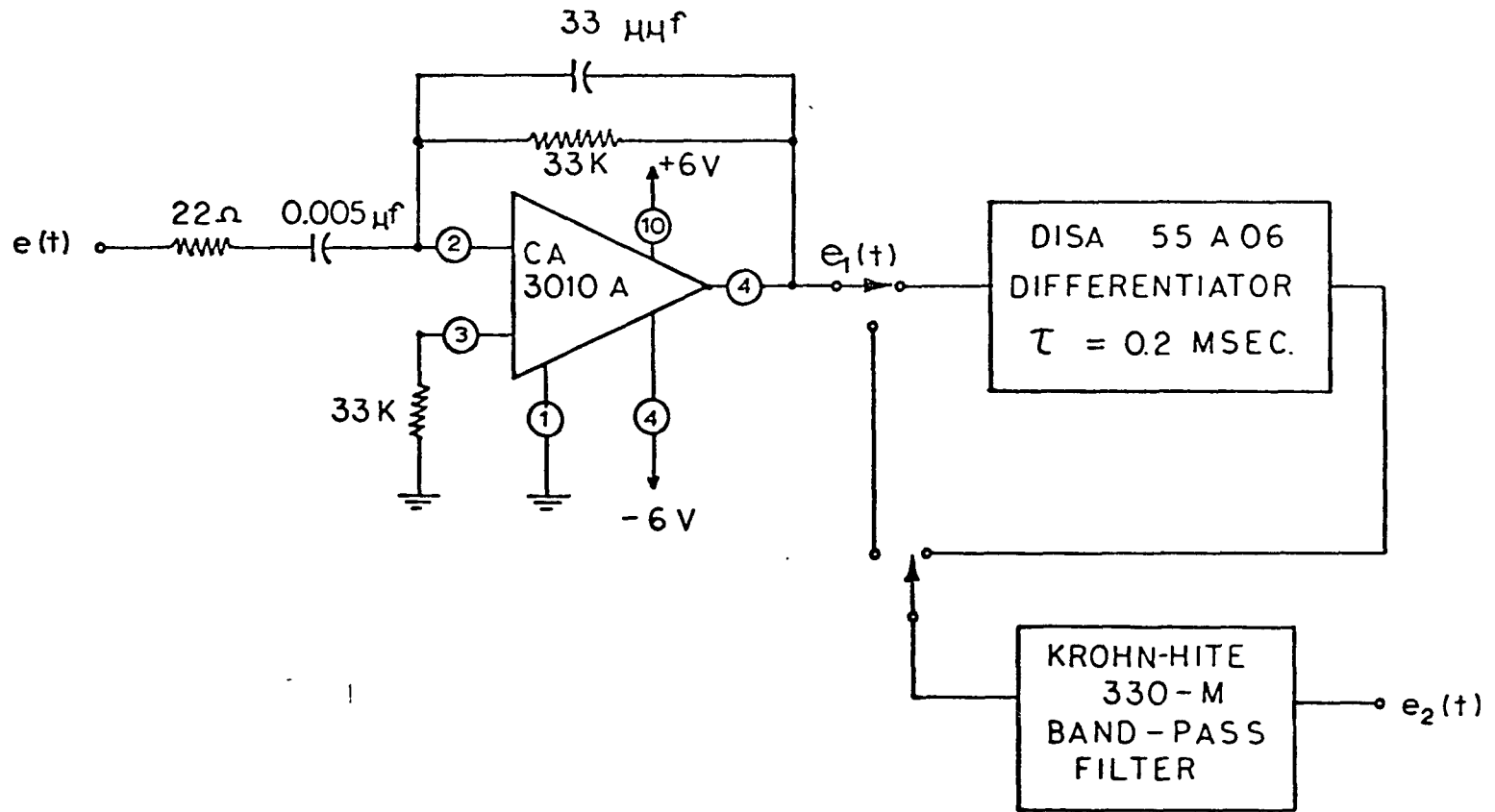


Figure 2. Differentiation Circuit

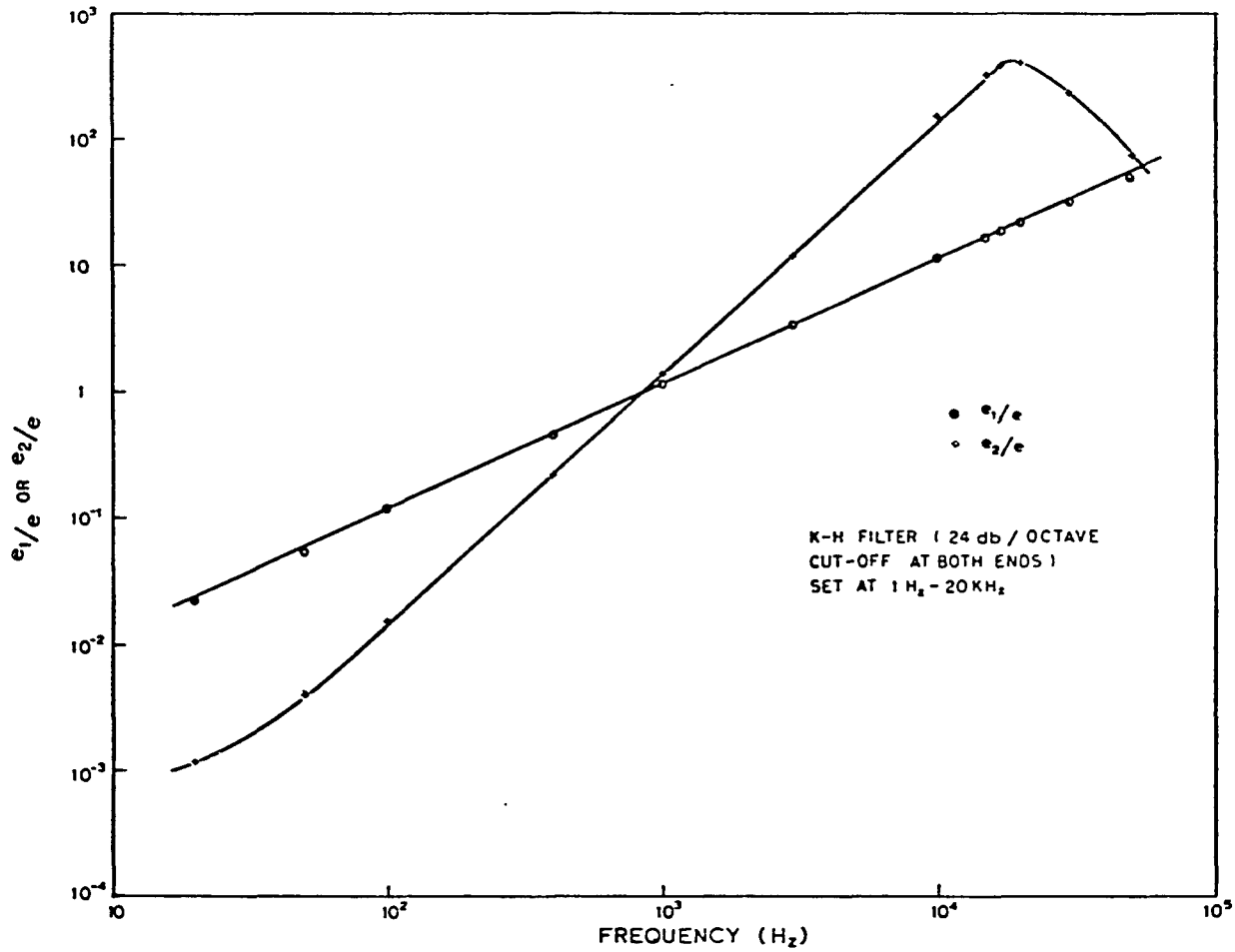


Figure 3. Frequency Response of Differentiation Circuit

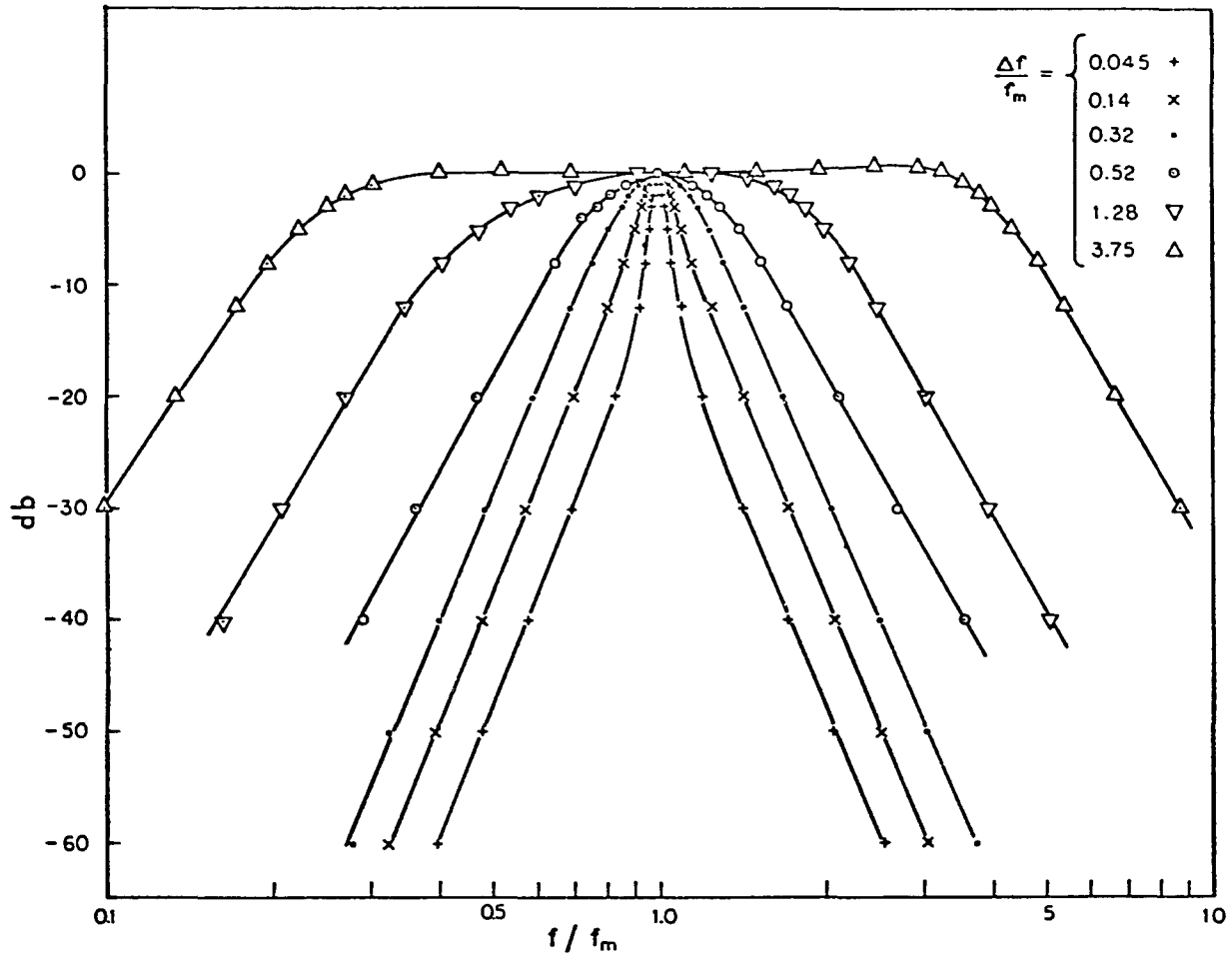


Figure 4. Frequency Response of Band-pass Filter

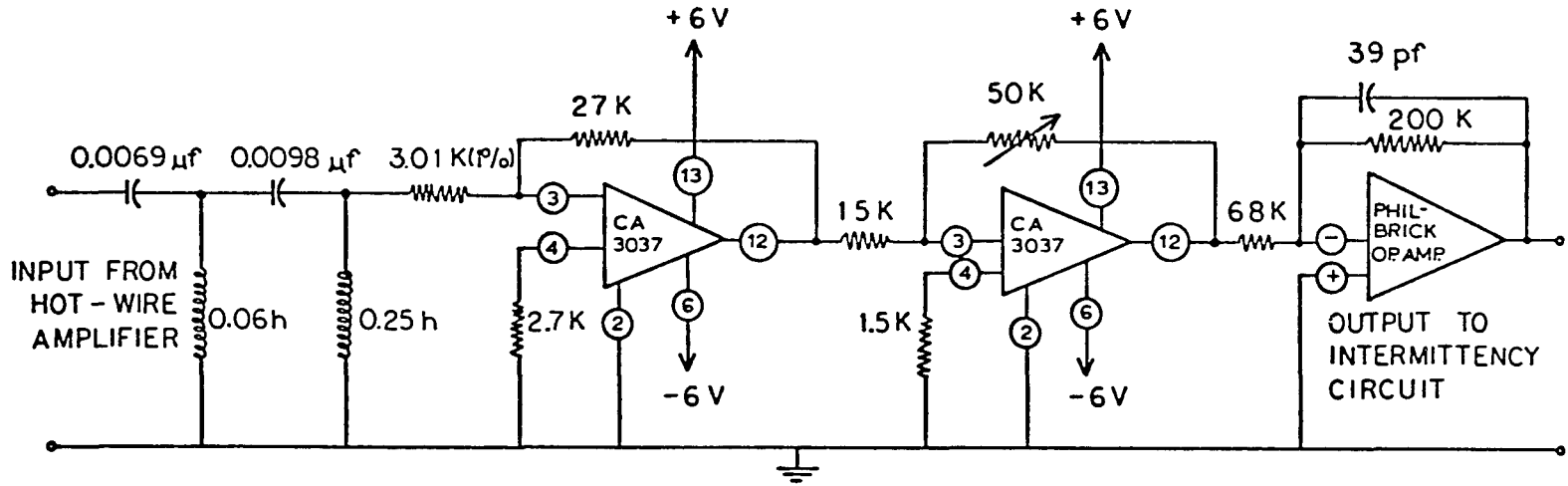


Figure 5. High-pass Butterworth Filter with Amplifiers

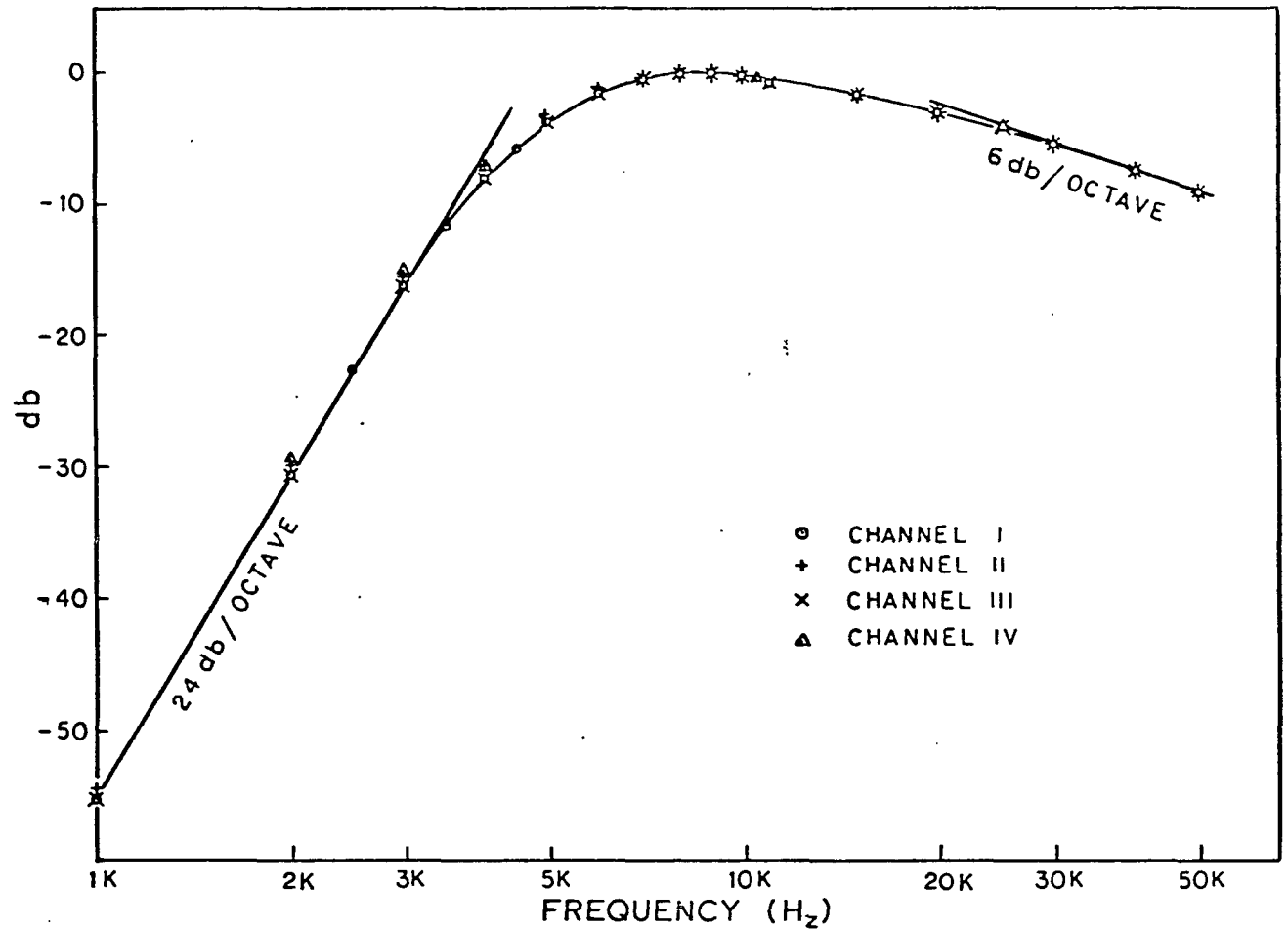


Figure 6. Frequency Response of High-pass Butterworth Filter with Amplifiers

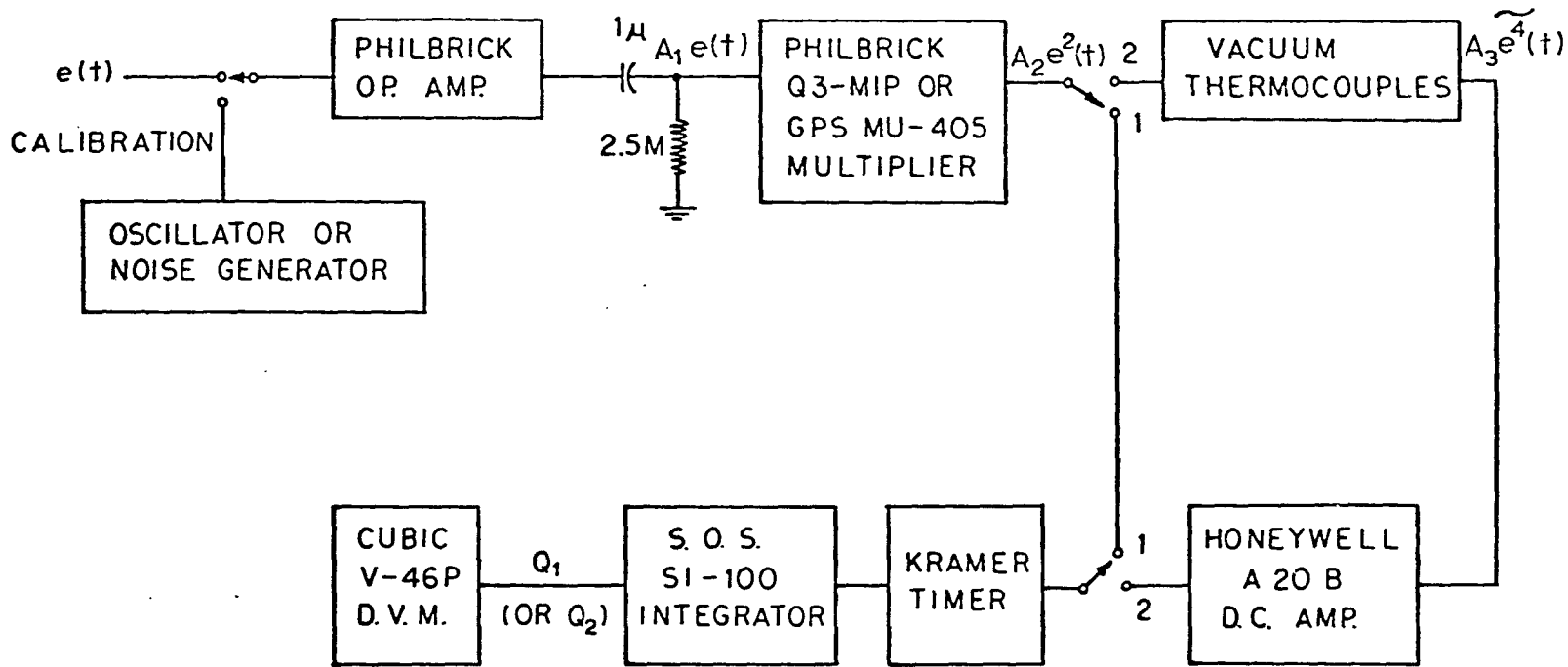


Figure 7. Block Diagram of the Flatness Factor Measuring Devices

(\sim means short time average, which is due to the small time constant of vacuum thermocouples)

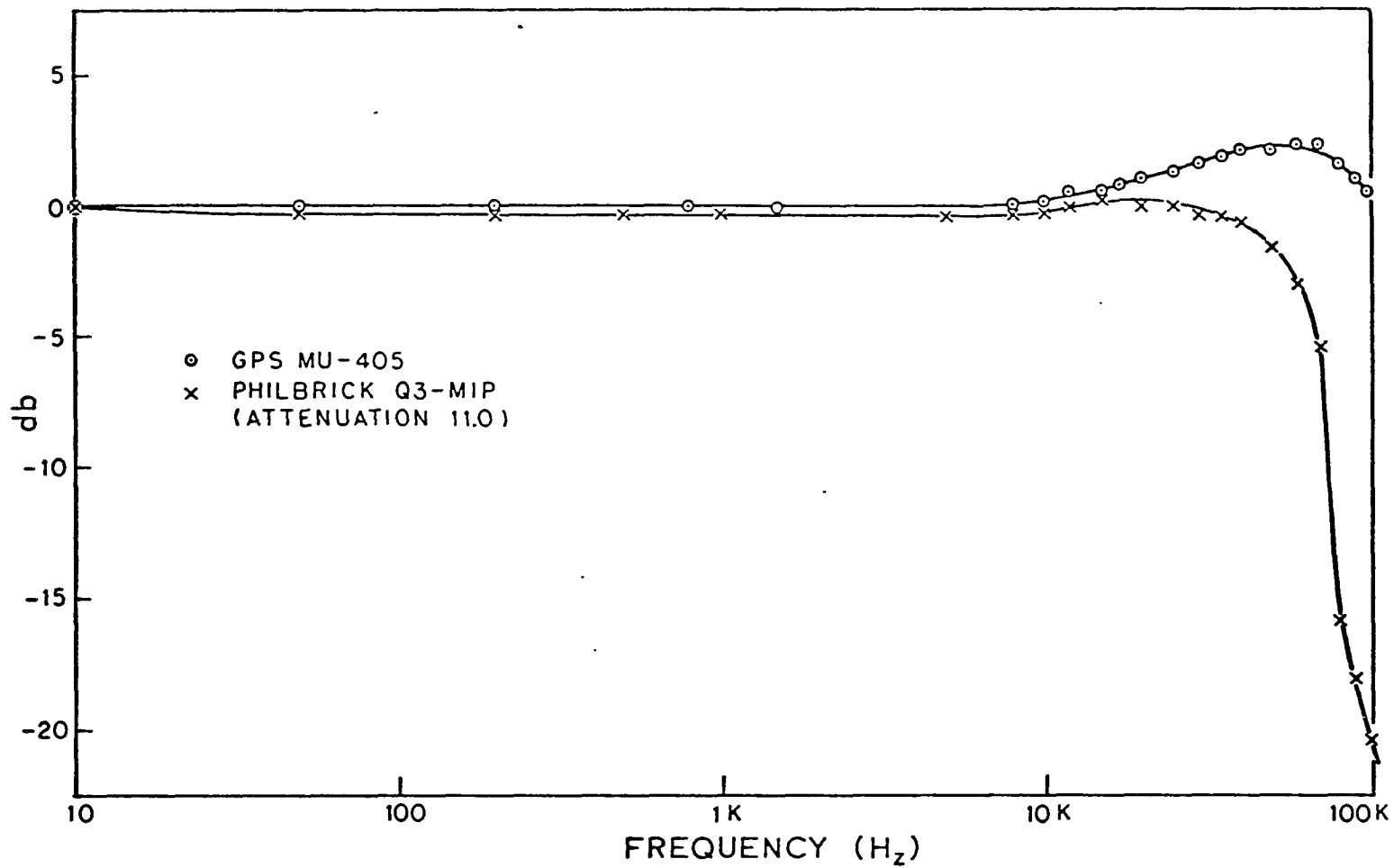


Figure 8. Frequency Response of Multipliers Used as Squaring Circuits

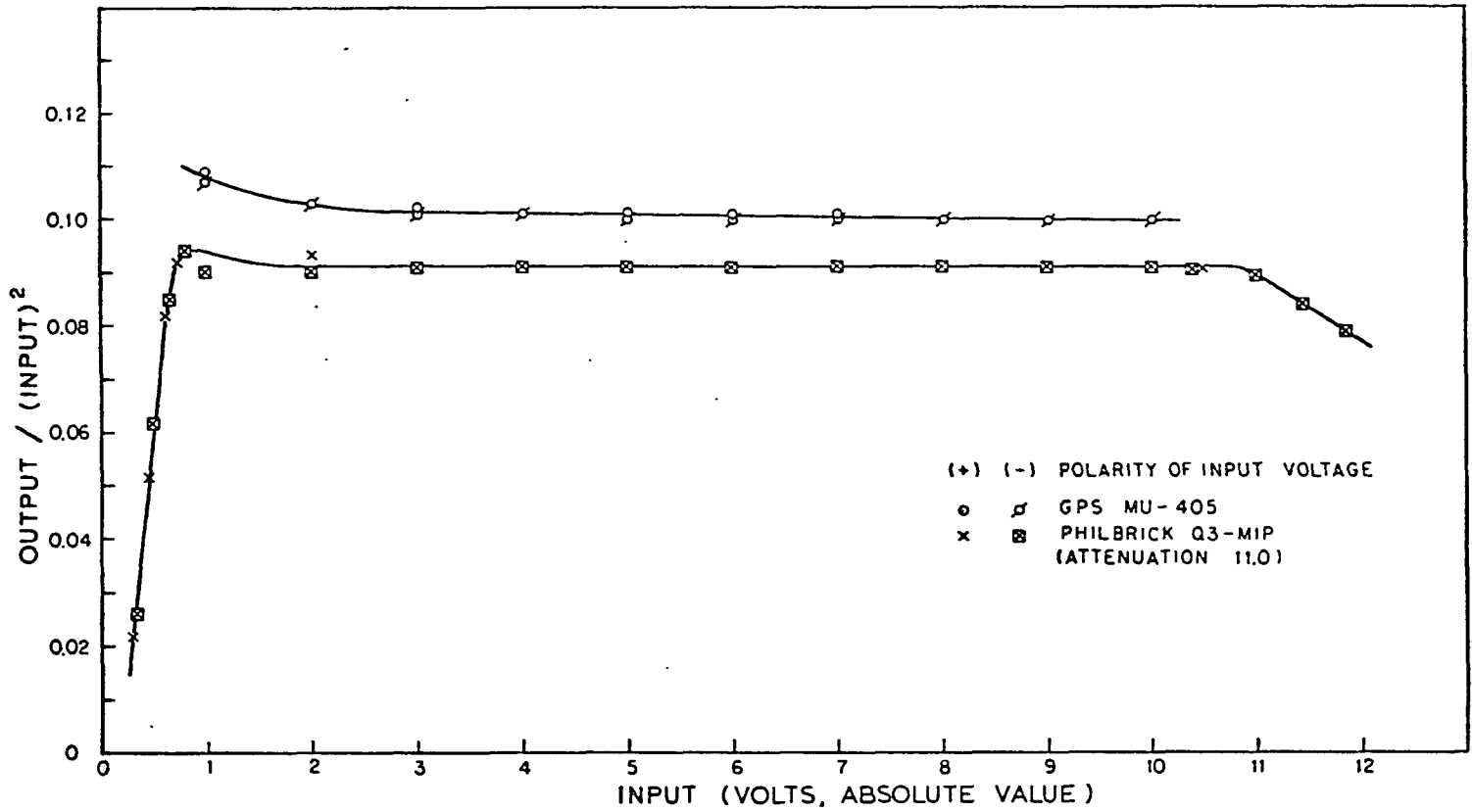


Figure 9. Static Transfer Characteristics of Multipliers Used as Squaring Circuits

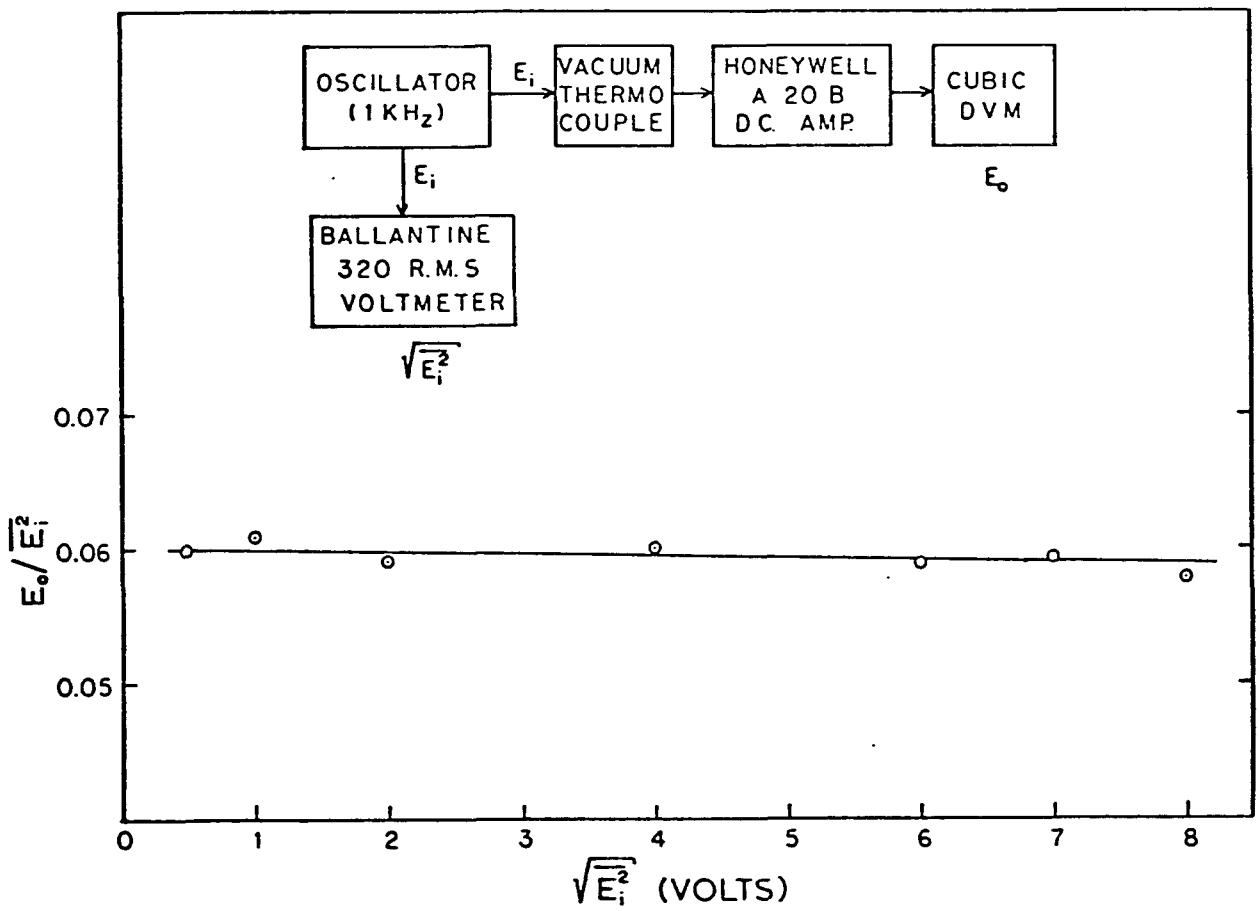


Figure 10. Transfer Characteristics of Vacuum Thermocouples

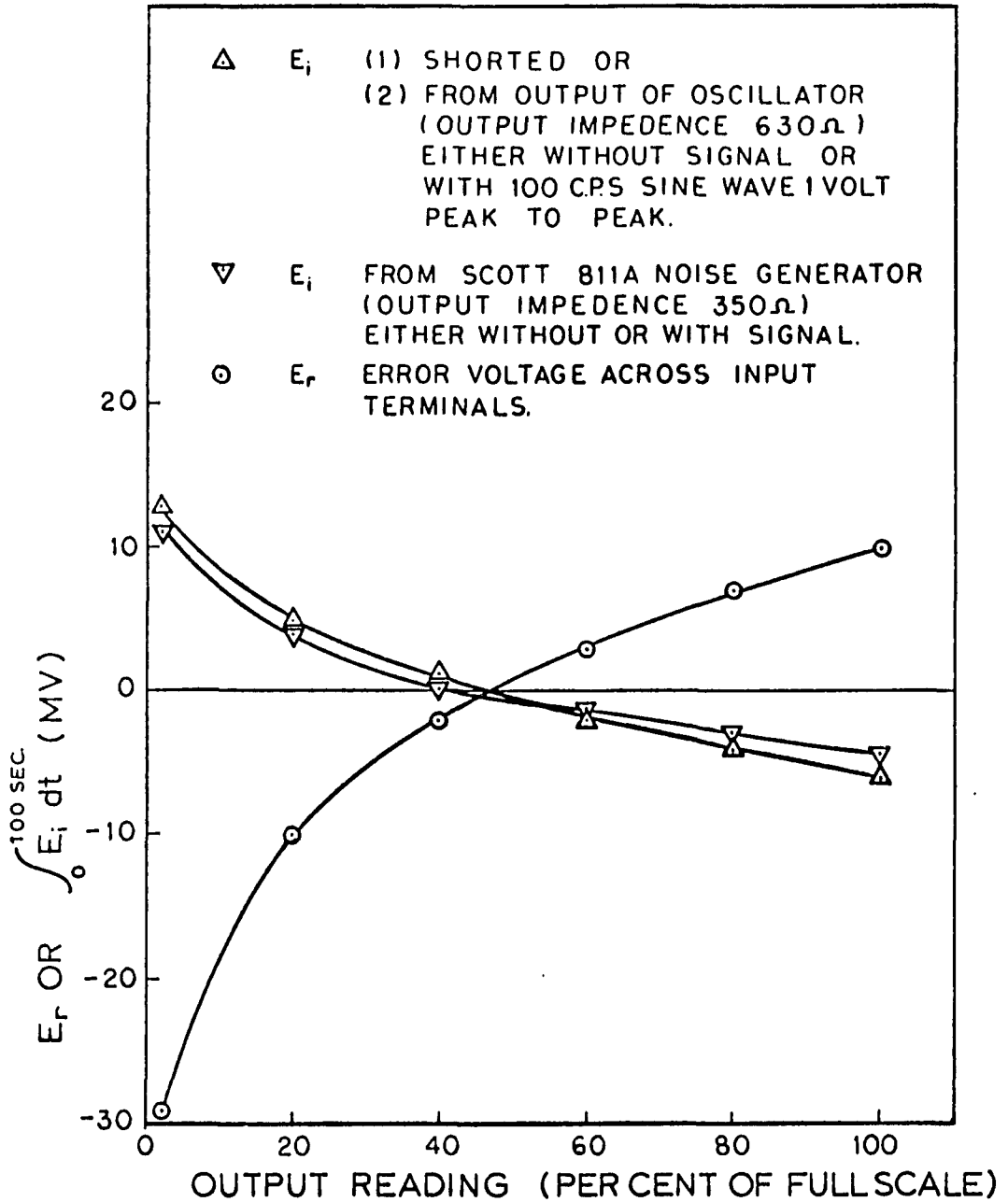


Figure 11. Drift Characteristics of Integrator

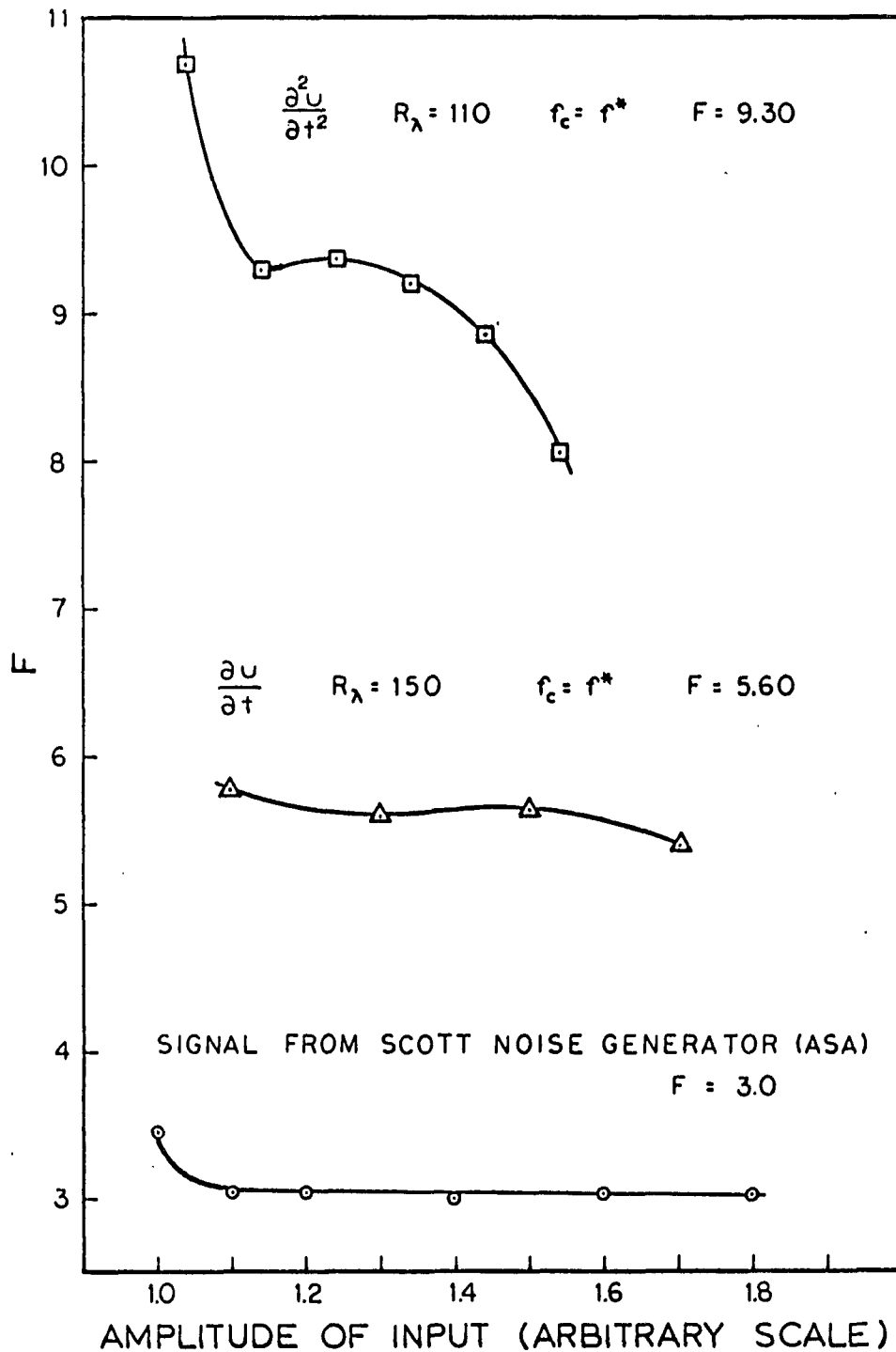


Figure 12. Apparent Flatness Factor vs. Amplitude of Input.

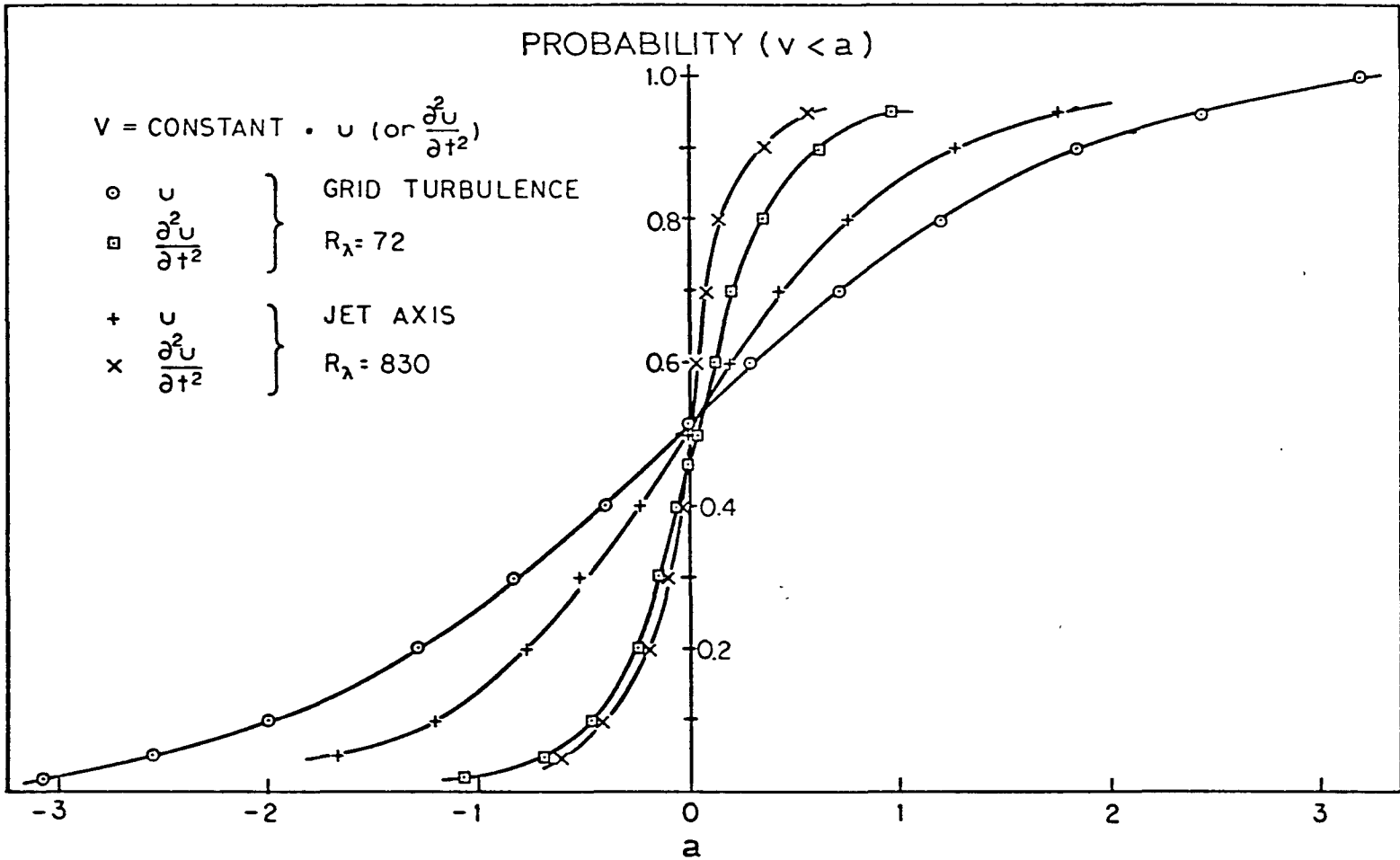


Figure 13. Some Samples of the Data of Probability Distribution

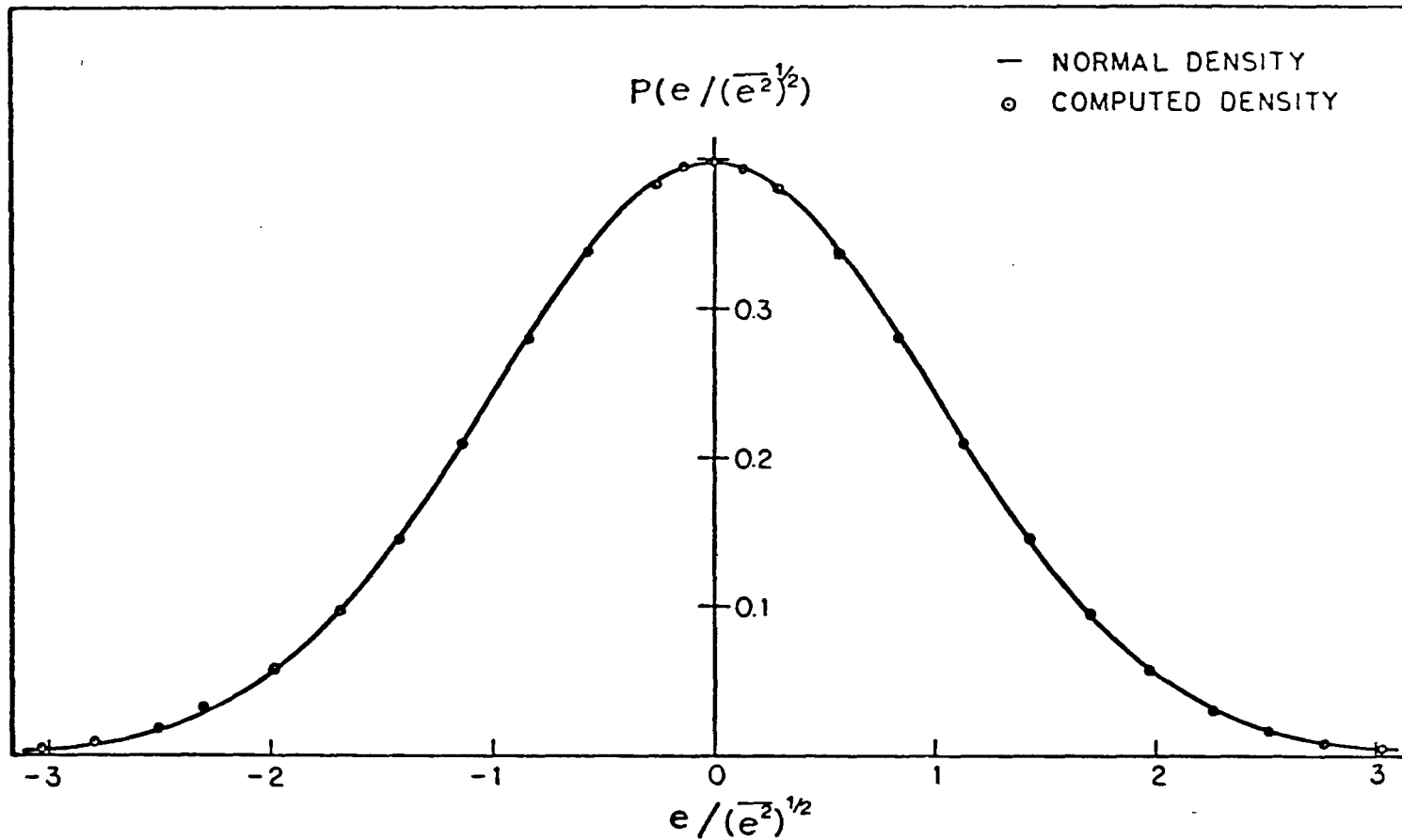


Figure 14. Computed Probability Density from Normal Distribution

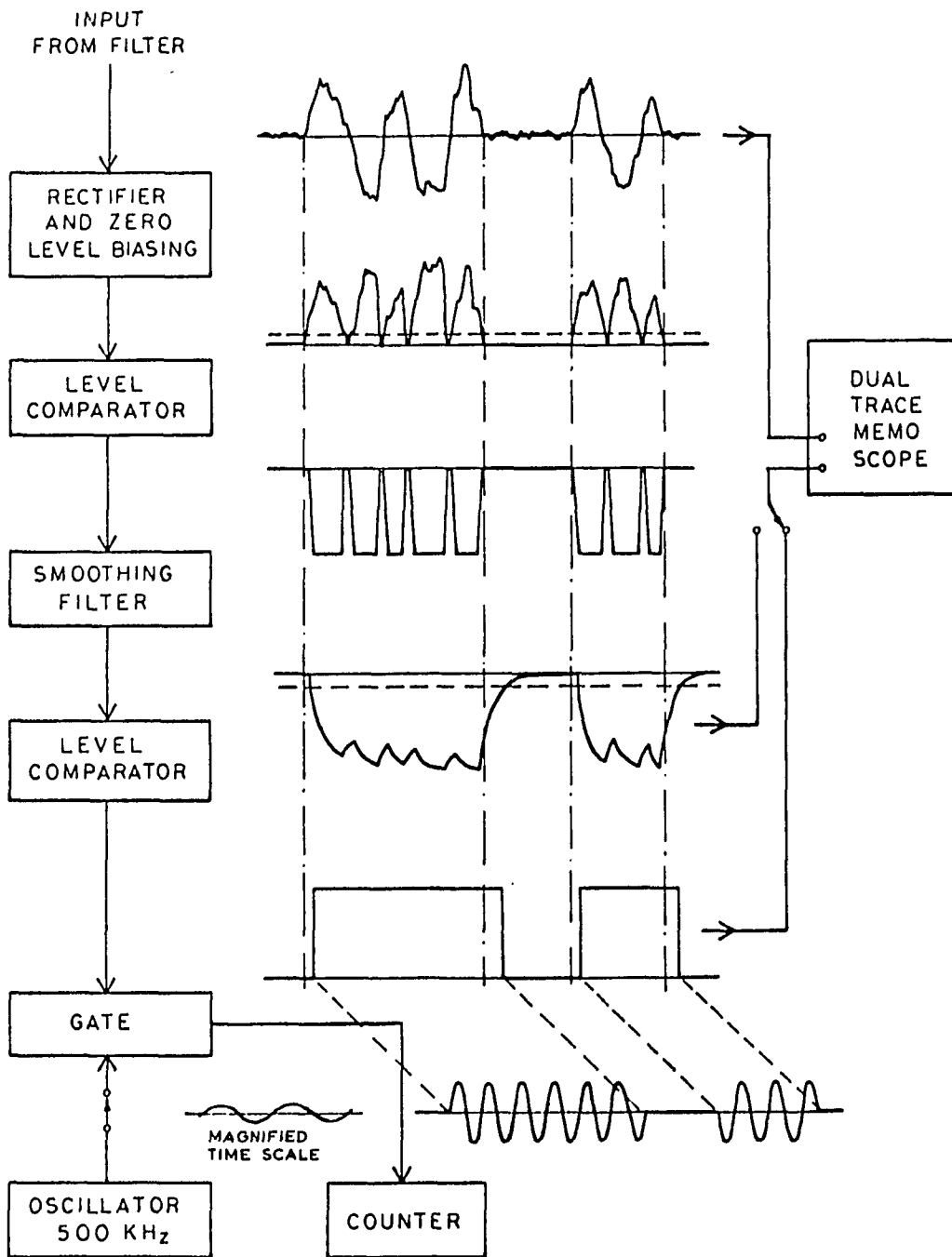


Figure 15. Block Diagram of Intermittency-Measuring Device and Operations on a Hypothetical Signal

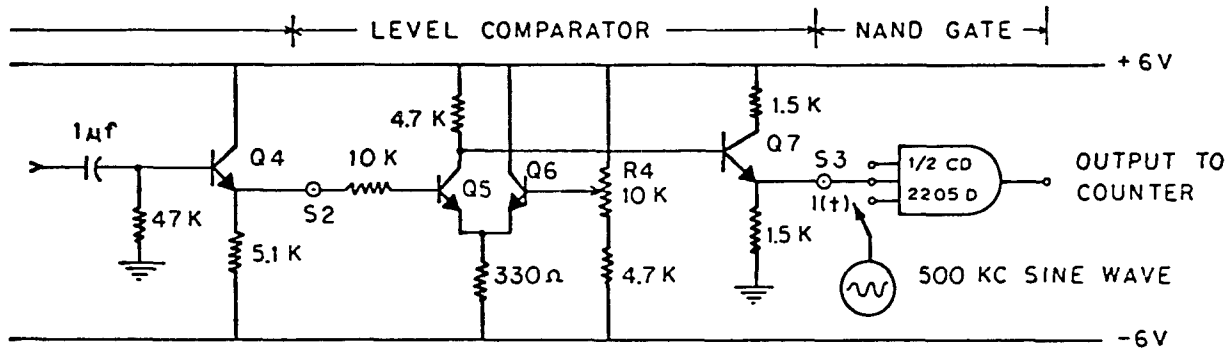
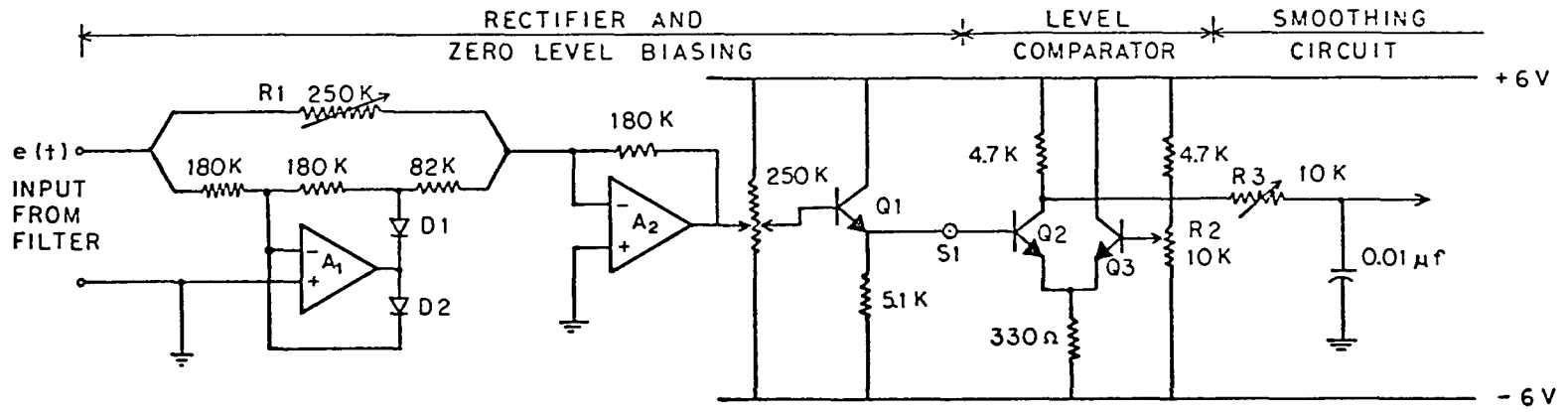


Figure 16. Circuit Diagram of Intermittency-Measuring Device

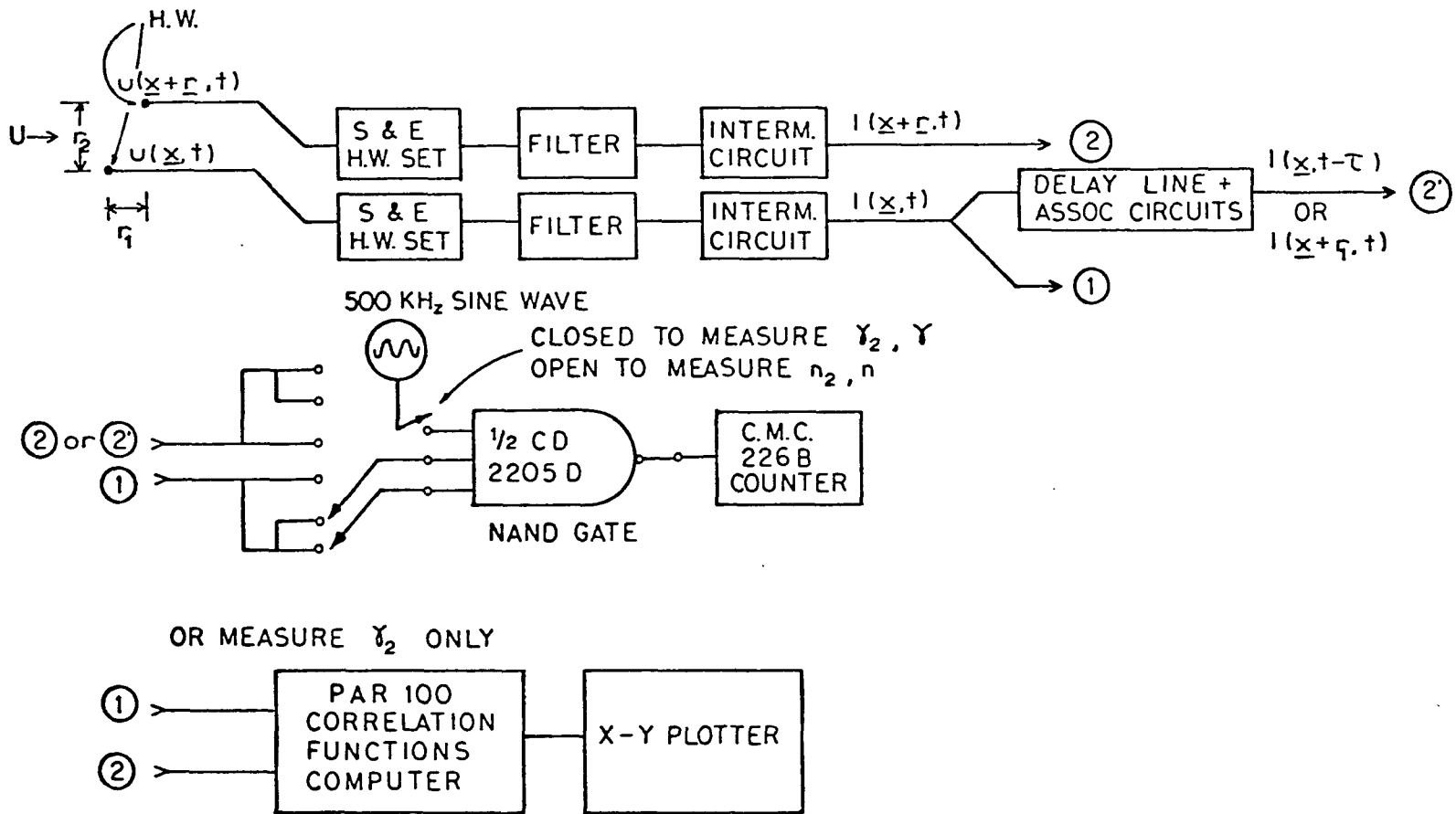


Figure 17. Block Diagram of Instruments to Measure n_2 and γ_2

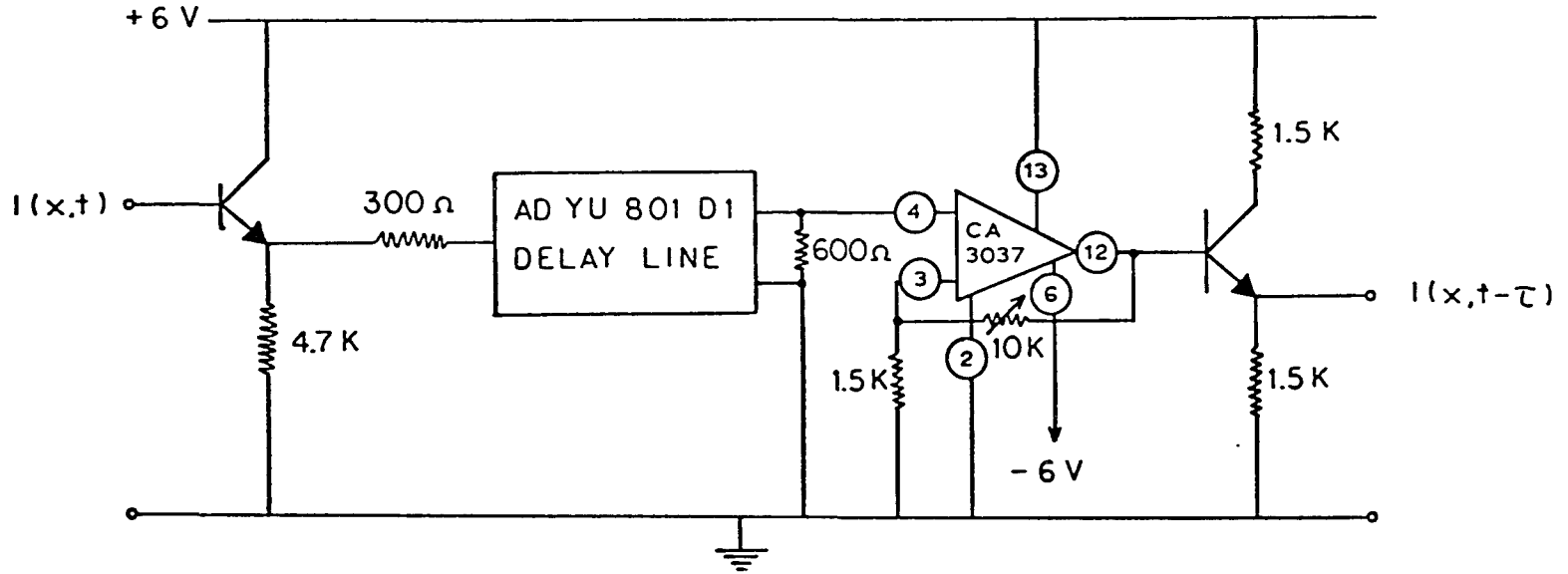


Figure 18. Delay Line and Associated Circuits

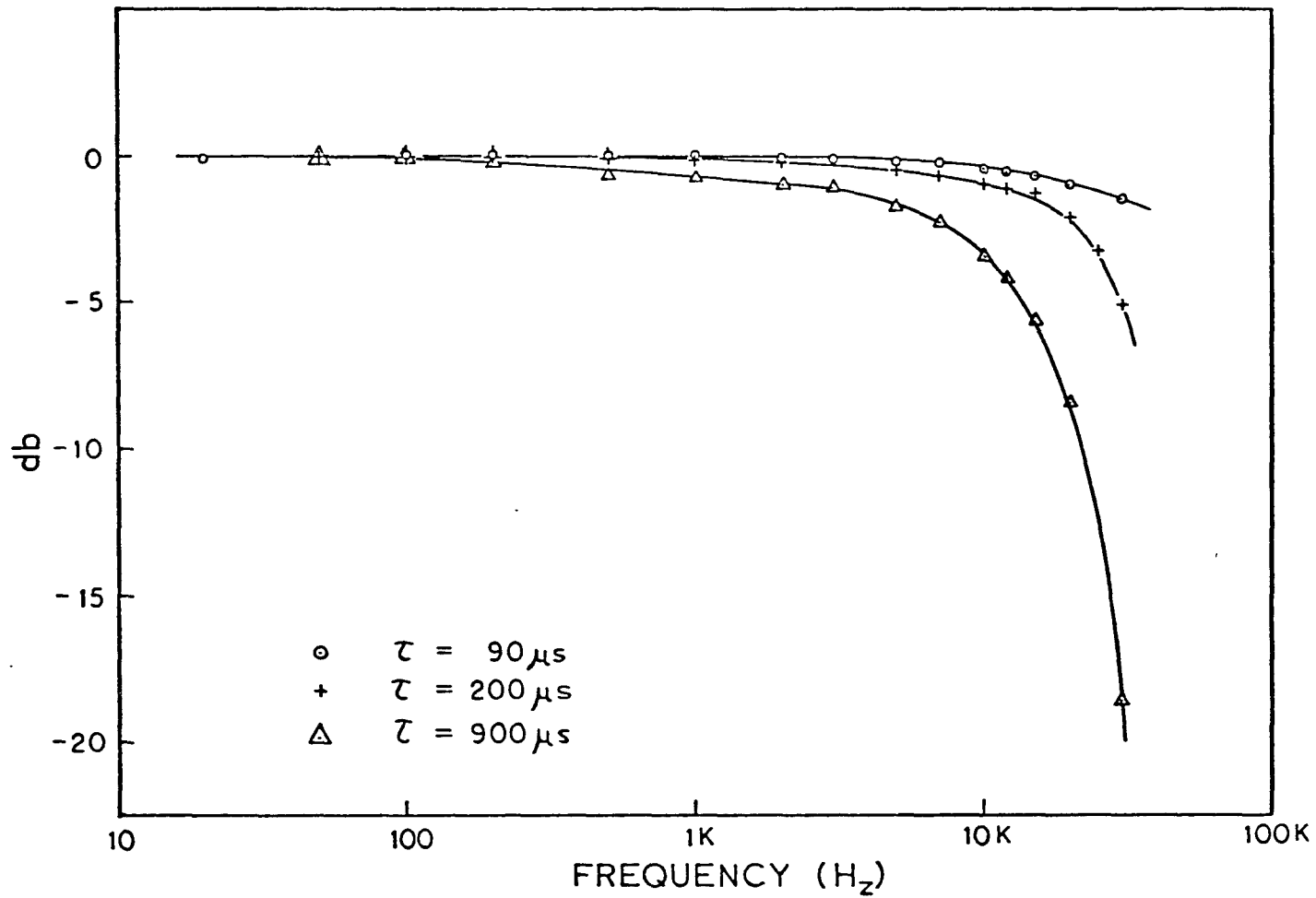


Figure 19. Frequency Response of Delay Line

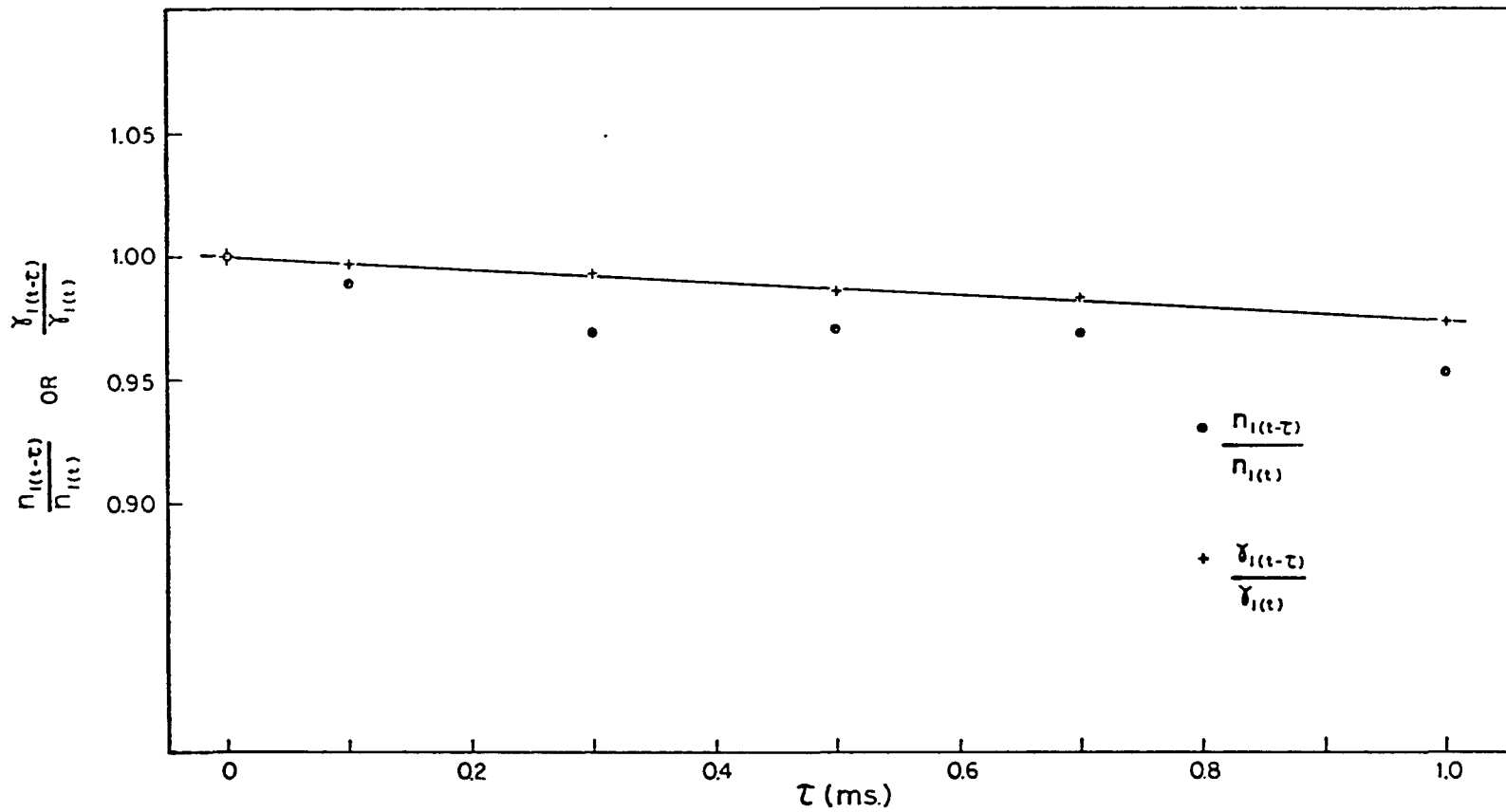
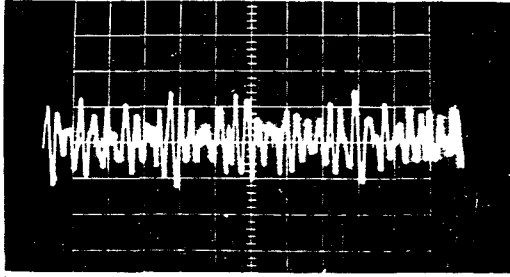
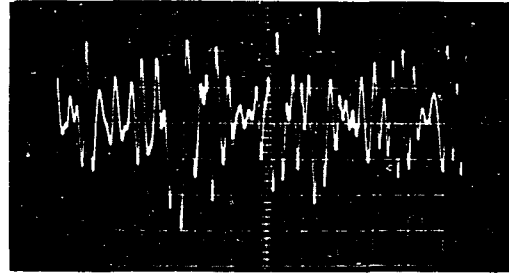


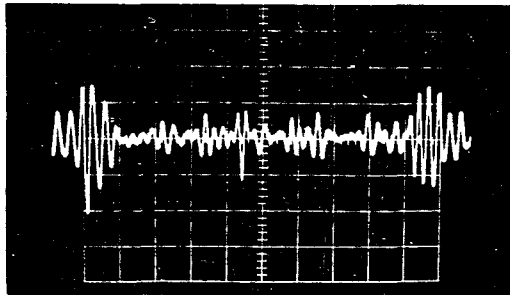
Figure 20. A Performance Check of Delay Line Circuit



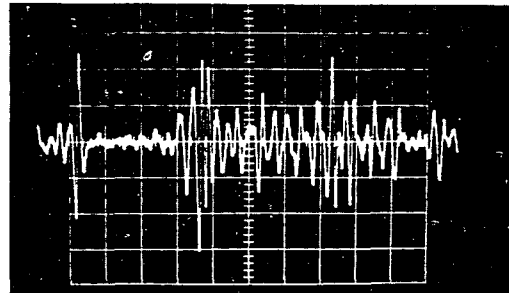
$f_m = 200 \text{ Hz}, \Delta f/f_m = 0.52$
 Horizontal Scale 20 ms/division



$f_m = 1 \text{ kHz}, \Delta f/f_m = 0.52$
 Horizontal Scale 4 ms/division

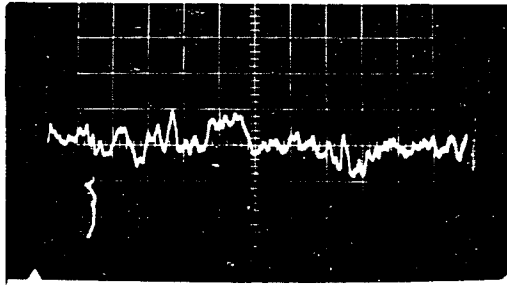


$f_m = 6 \text{ kHz}, \Delta f/f_m = 0.52$
 Horizontal Scale 1 ms/division



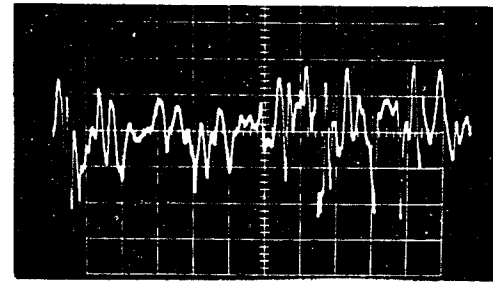
High-Pass Signal, $f_c = 5 \text{ kHz}$
 Horizontal Scale 1 ms/division

Figure 21. Oscillograms of Band-Pass and High-Pass Signals from a Hot-Wire in a Grid-Generated Turbulence $R_\lambda = 110, f^* = 5.9 \text{ kHz}$



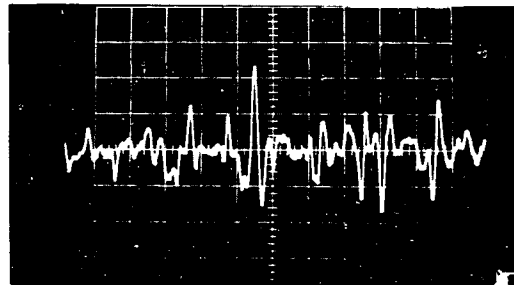
$$u(t)$$

Horizontal Scale 4 ms/division



$$\frac{\partial u}{\partial t}$$

Horizontal Scale 2 ms/division



$$\frac{\partial^2 u}{\partial t^2}$$

Horizontal Scale 1 ms/division

Figure 22. Oscillograms of Velocity Fluctuation and Time Derivatives in a Grid-Generated Turbulence $R_\lambda = 110$, $f^* = 5.9$ kHz

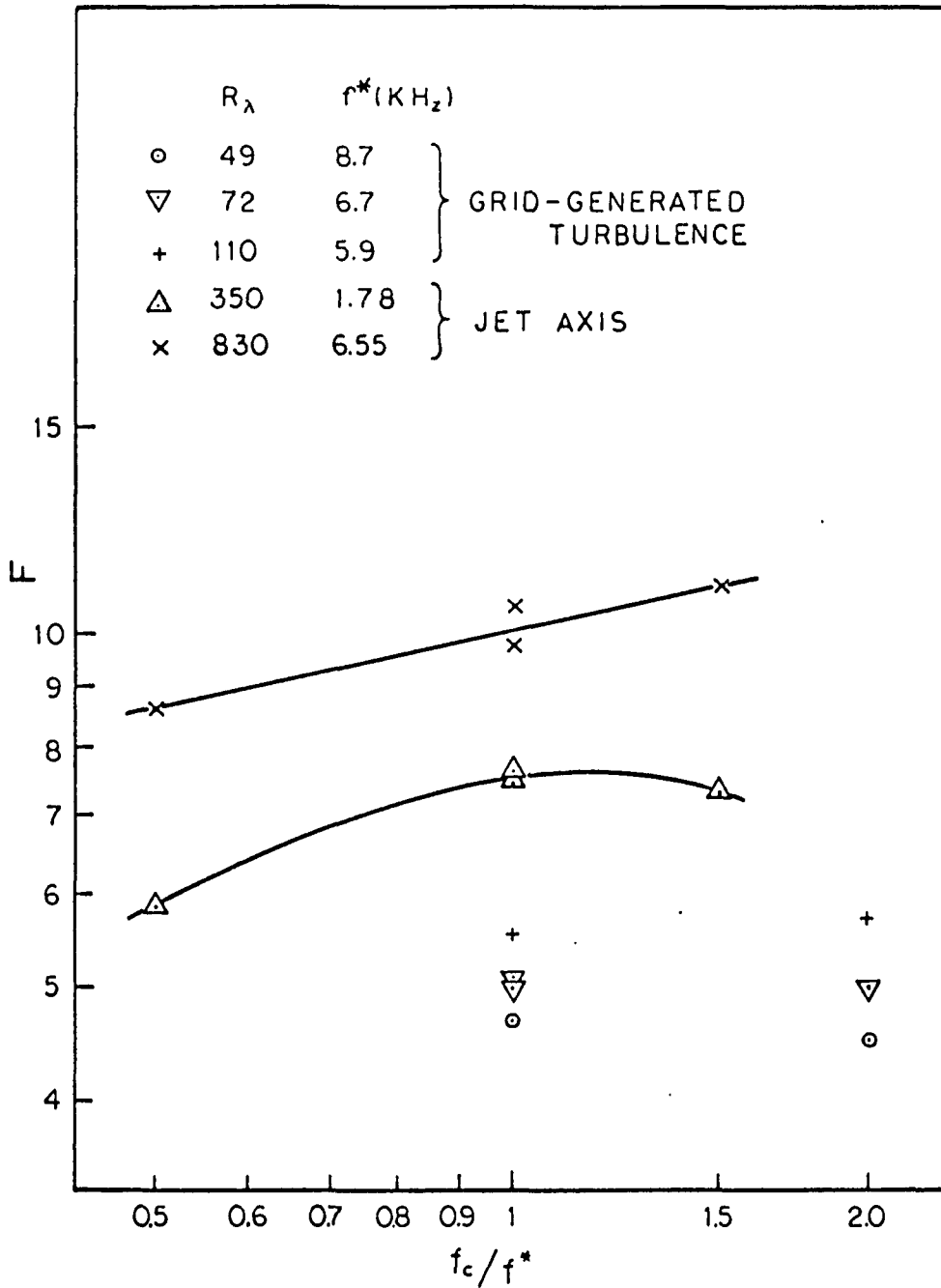


Figure 23. Flatness Factor of $\frac{\partial u}{\partial t}$ as a Function of High Cut-off Frequency $\frac{f_c}{f^*}$.

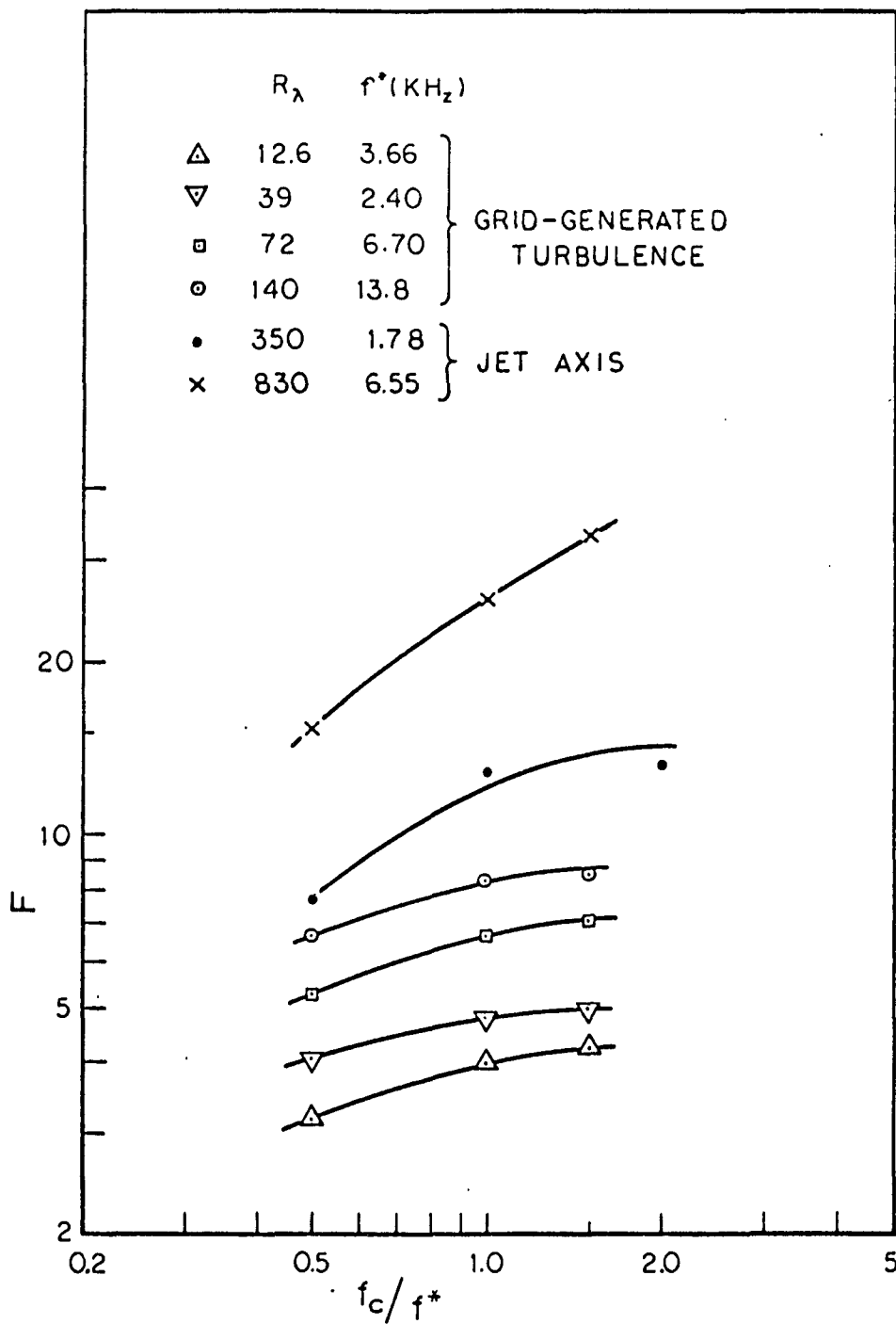


Figure 24. Flatness Factor of $\frac{\partial^2 u}{\partial t^2}$ as a
Function of High Cut-off Frequency $\frac{f_c}{f^*}$.

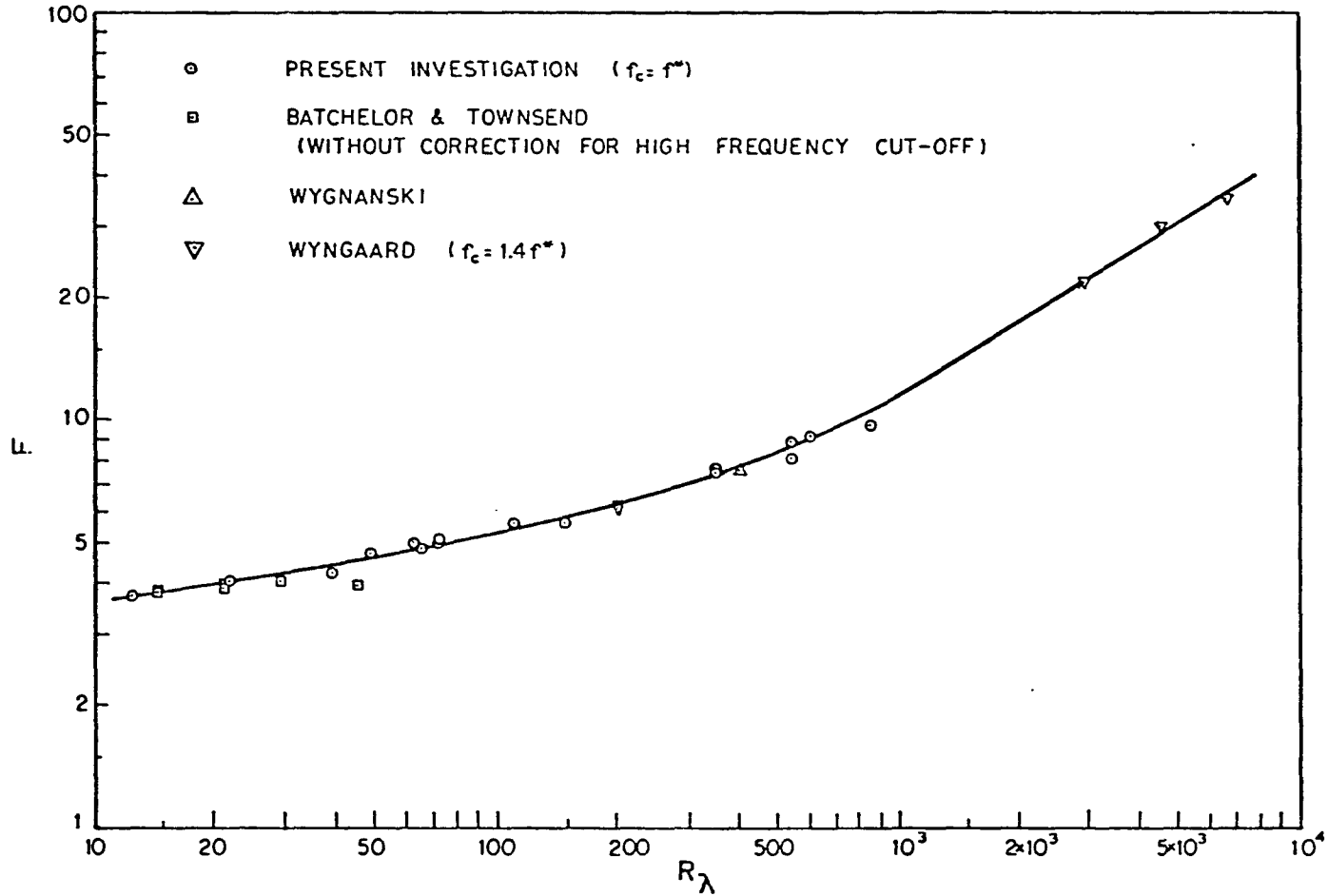


Figure 25. Flatness Factor of $\frac{\partial u}{\partial t}$ as a Function of Reynolds Number R_λ , with $f_c = f^*$.

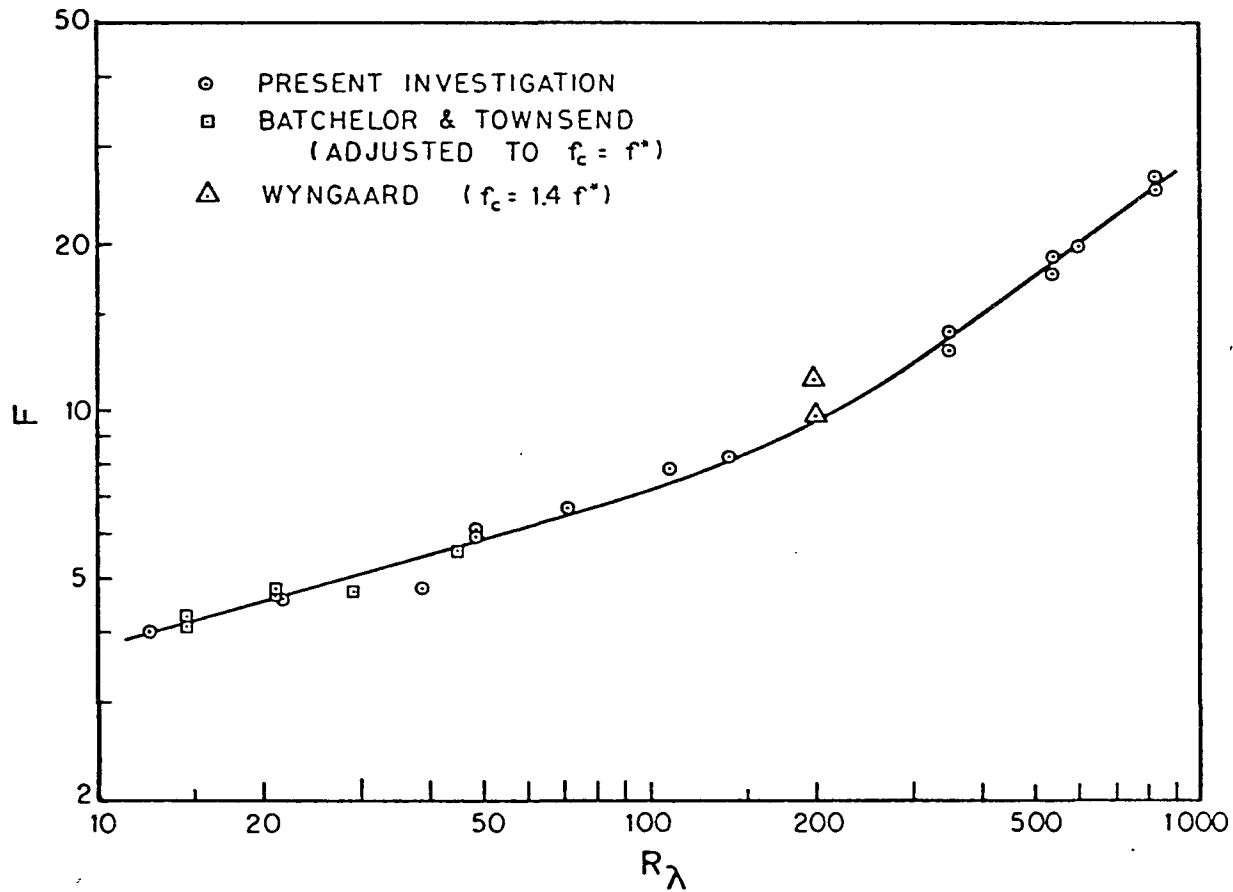


Figure 26. Flatness Factor of $\frac{\partial^2 u}{\partial t^2}$ as a Function of Reynolds Number R_λ , with $f_c = f^*$.

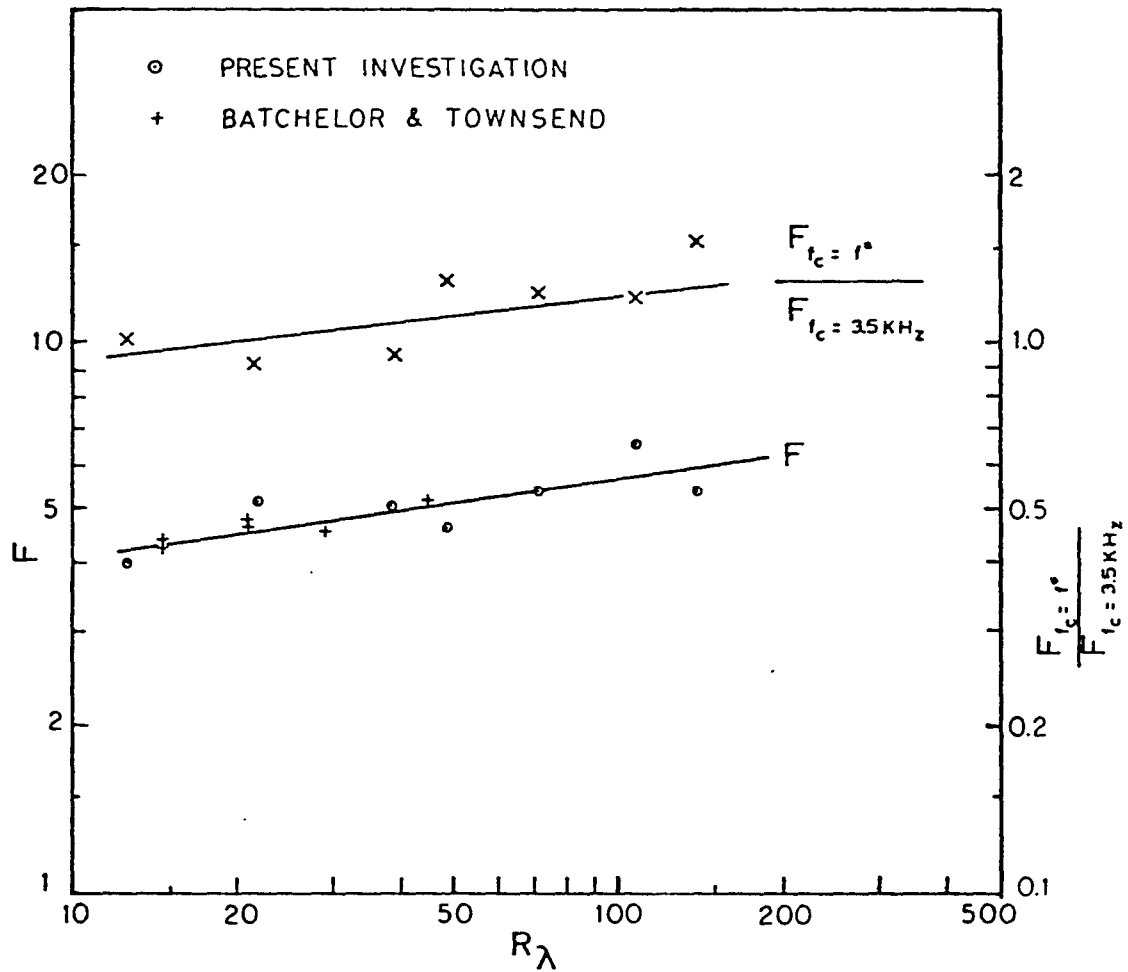


Figure 27. Flatness Factor of $\frac{\partial^2 u}{\partial t^2}$ as a Function of Reynolds Number R_λ , with $f_c = 3.5 \text{ kHz}$.

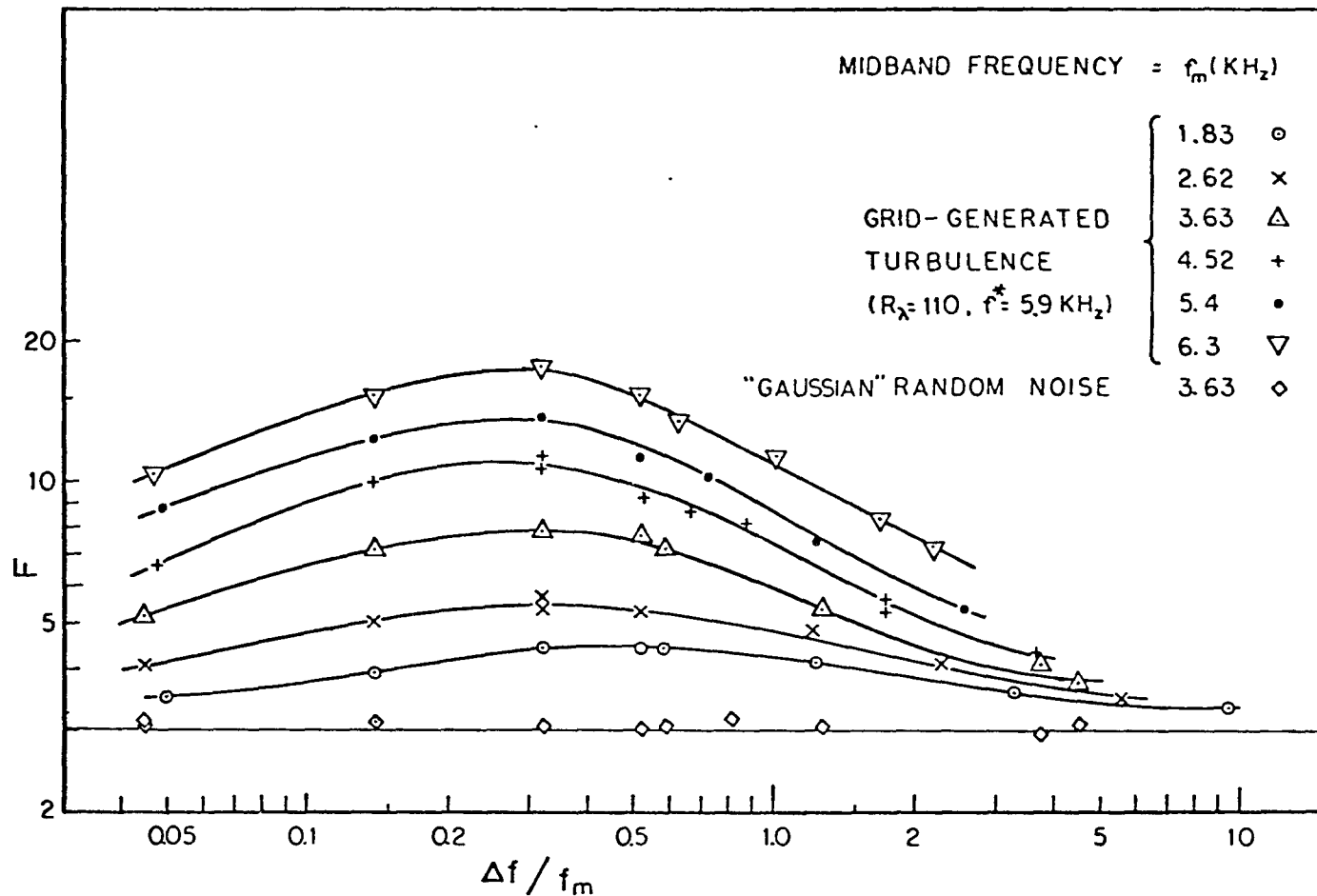


Figure 28. Flatness Factor of Band-pass Signal as a Function of Bandwidth. $\Delta f / f_m$,
Log-log Scale

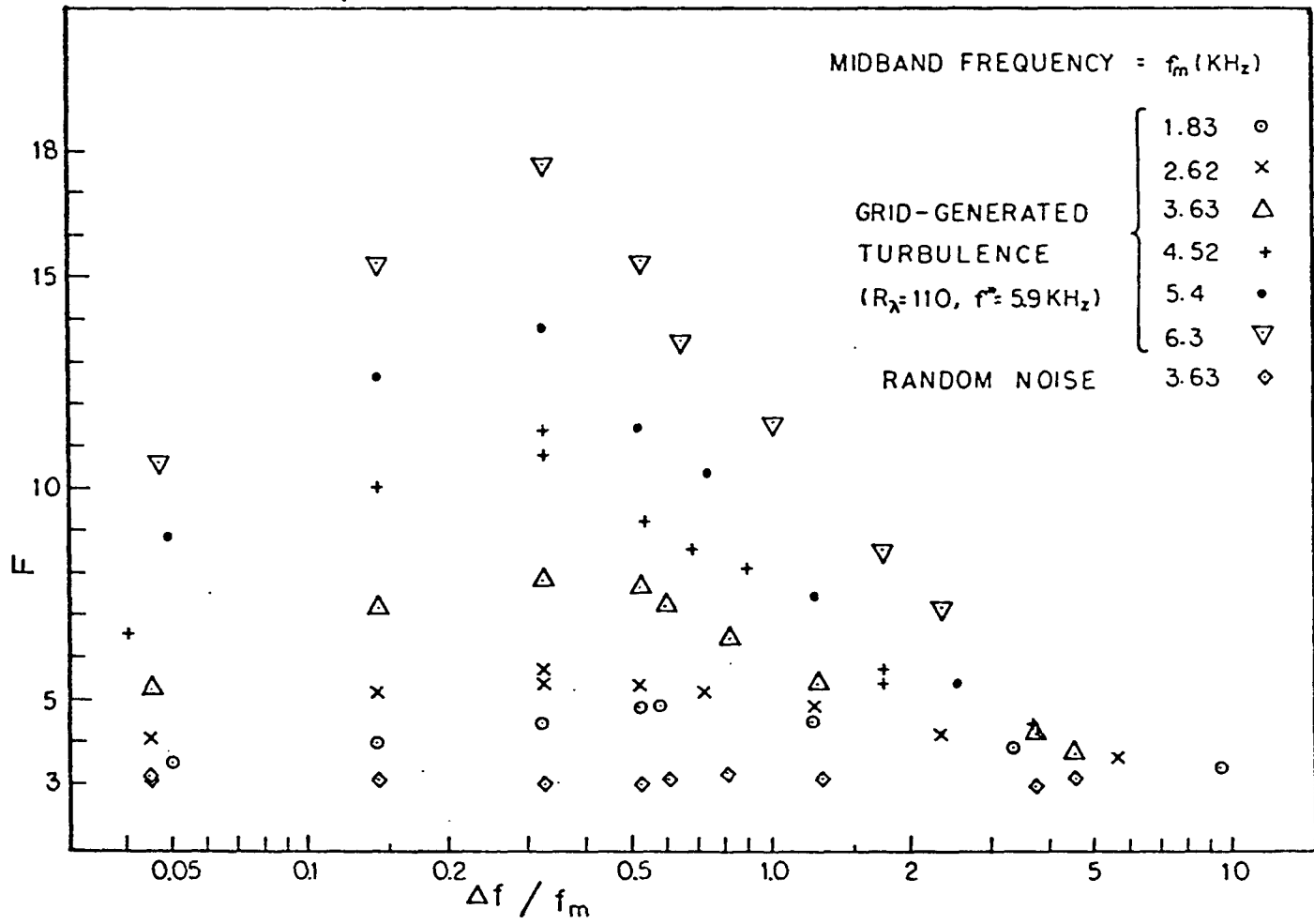


Figure 29. Flatness Factor of Band-pass Signal as a Function of Bandwidth $\Delta f / f_m$, Semi-log Scale

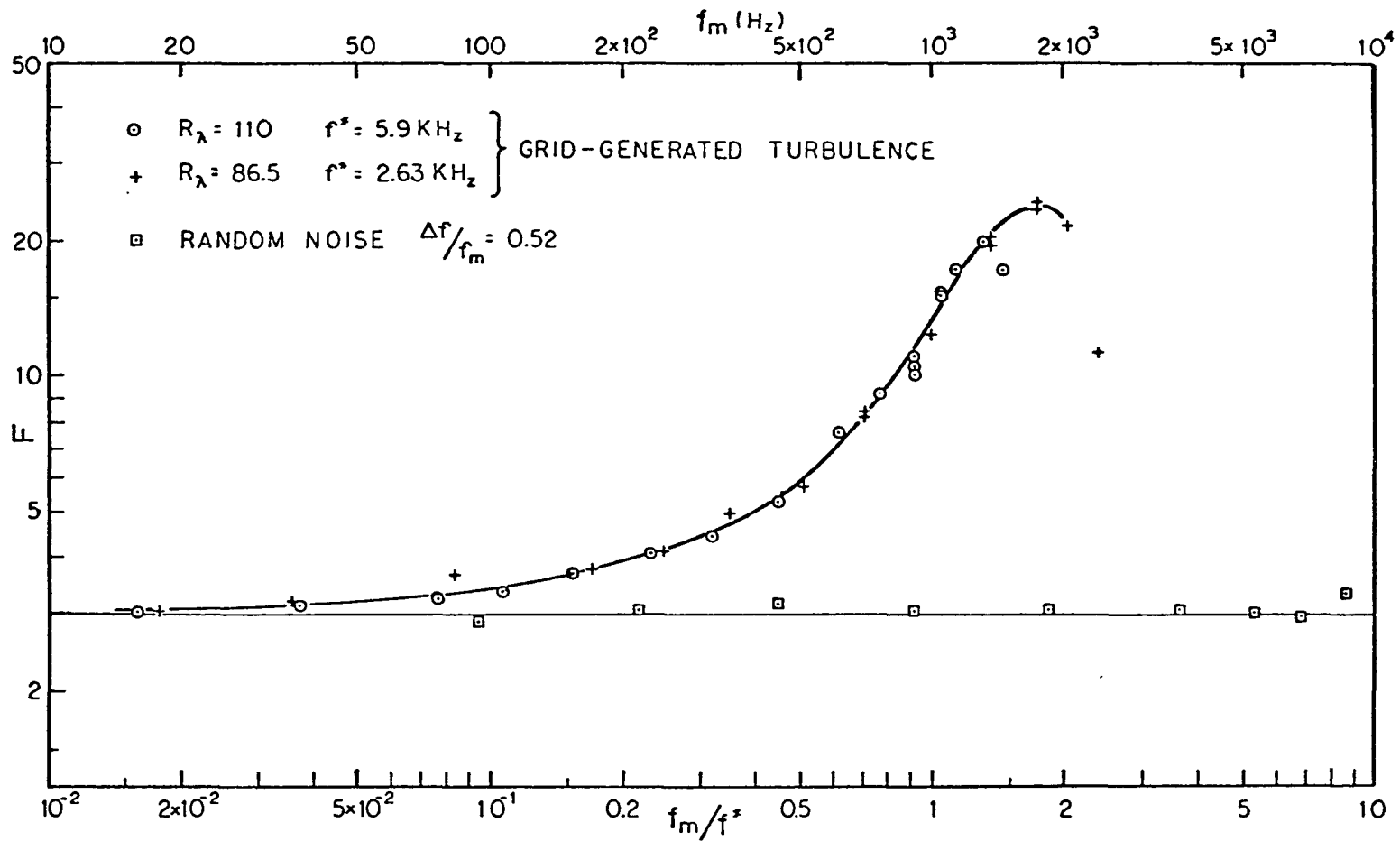


Figure 30. Flatness Factor of Band-pass Signal as a Function of Midband Frequency f_m/f^* .

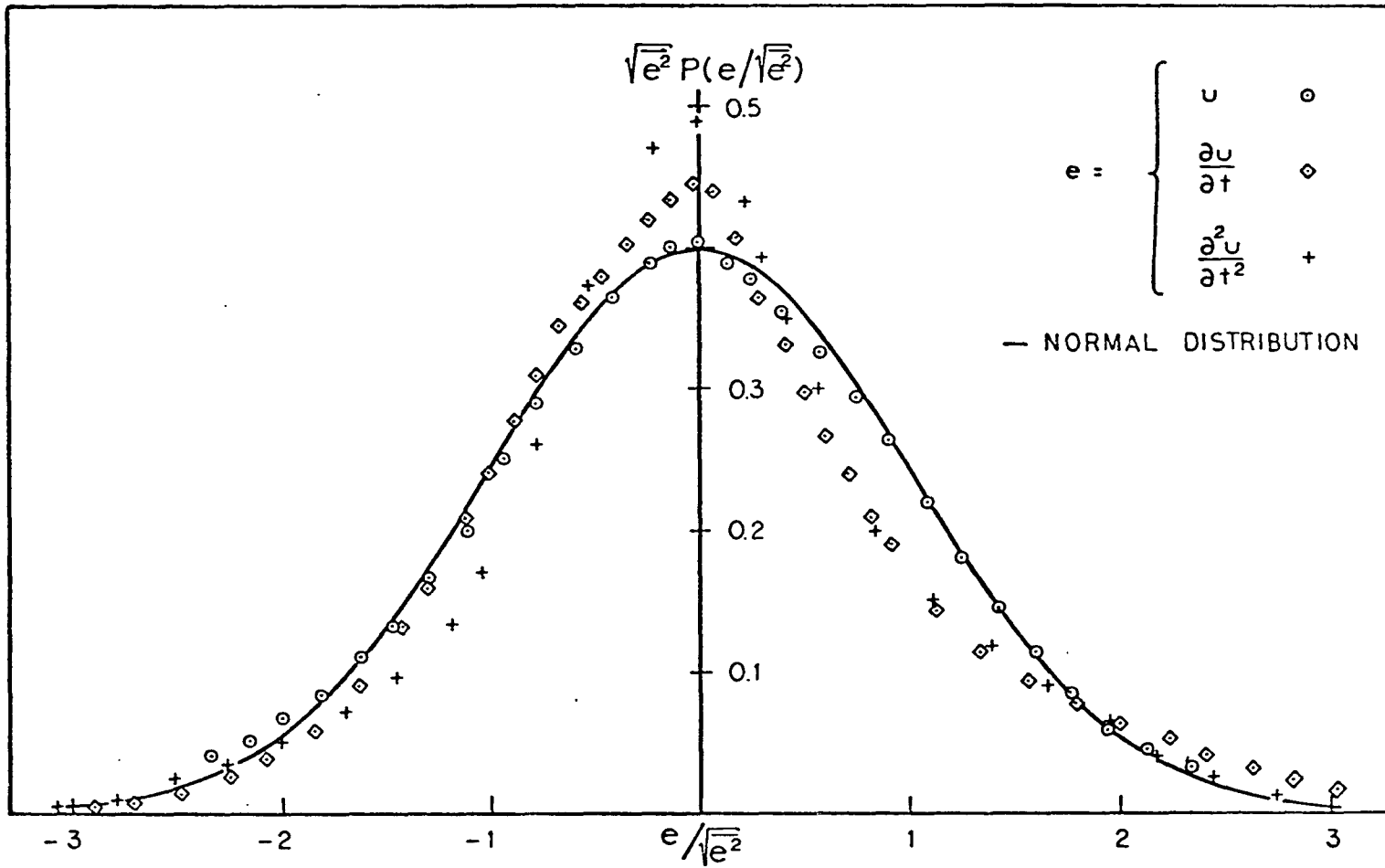


Figure 31. Probability Densities of u , $\partial u/\partial t$, and $\partial^2 u/\partial t^2$ in a Grid-generated Turbulence, $R_\lambda = 72$.

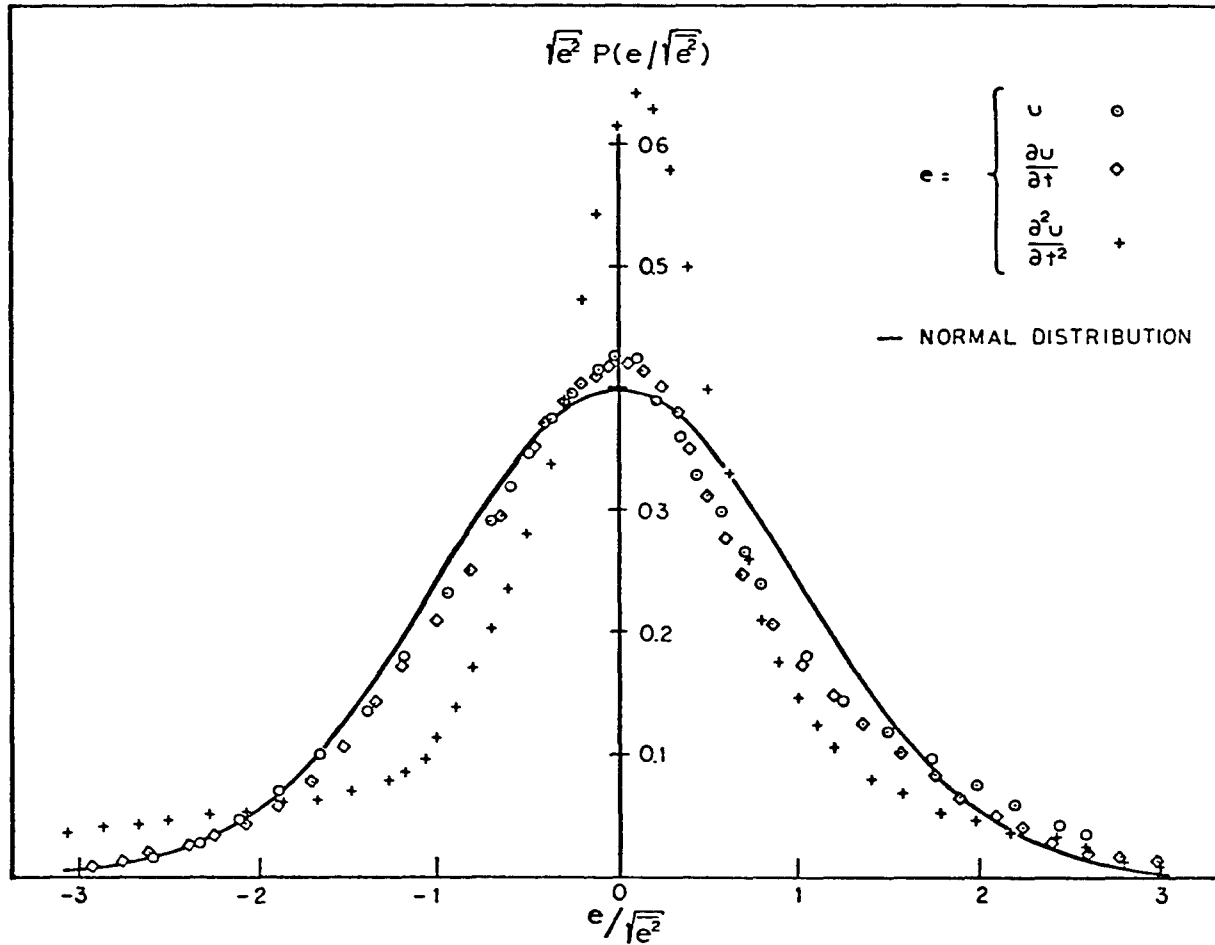


Figure 32. Probability Densities of u , $\frac{\partial u}{\partial t}$, and $\frac{\partial^2 u}{\partial t^2}$
 on the Axis of a Round Jet, $R_\lambda = 830$.

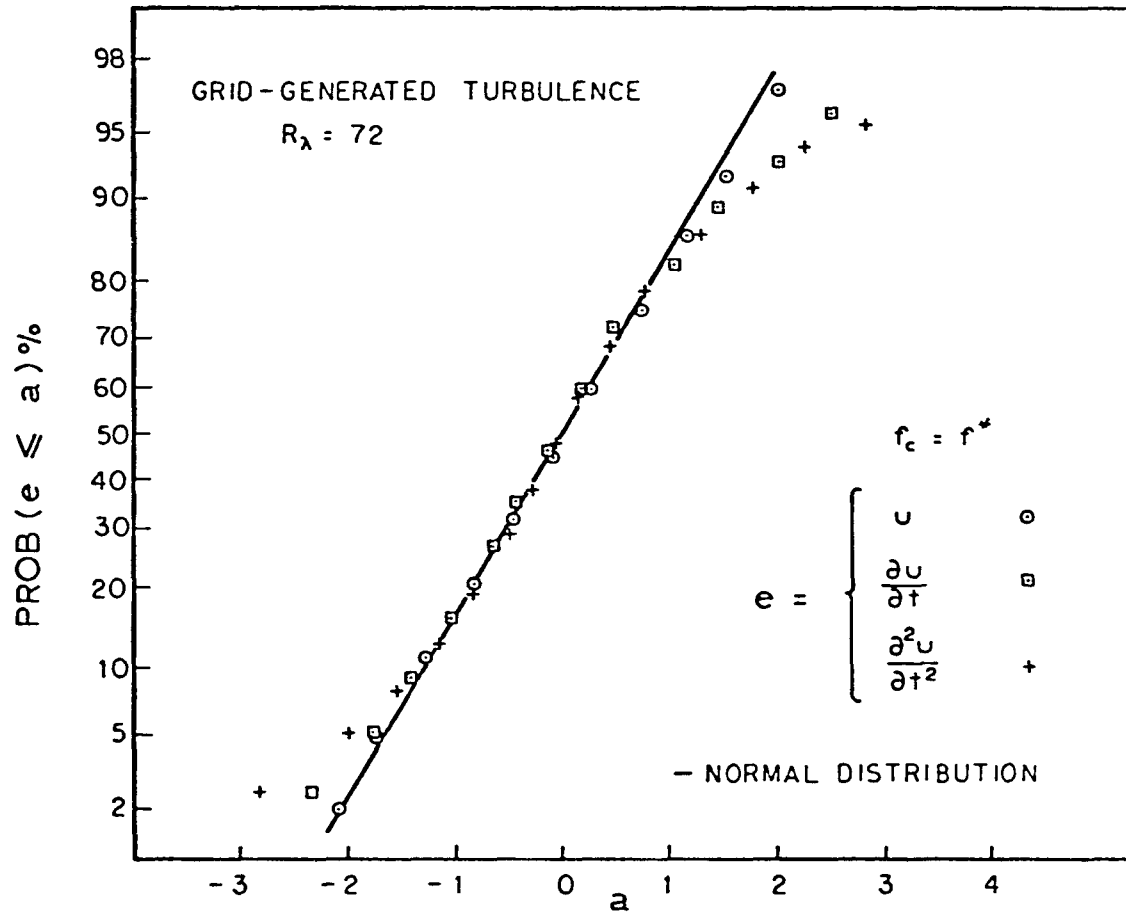


Figure 33. Probability Distributions of u , $\frac{\partial u}{\partial t}$ and $\frac{\partial^2 u}{\partial t^2}$ in a Grid-generated Turbulence, $R_\lambda = 72$.

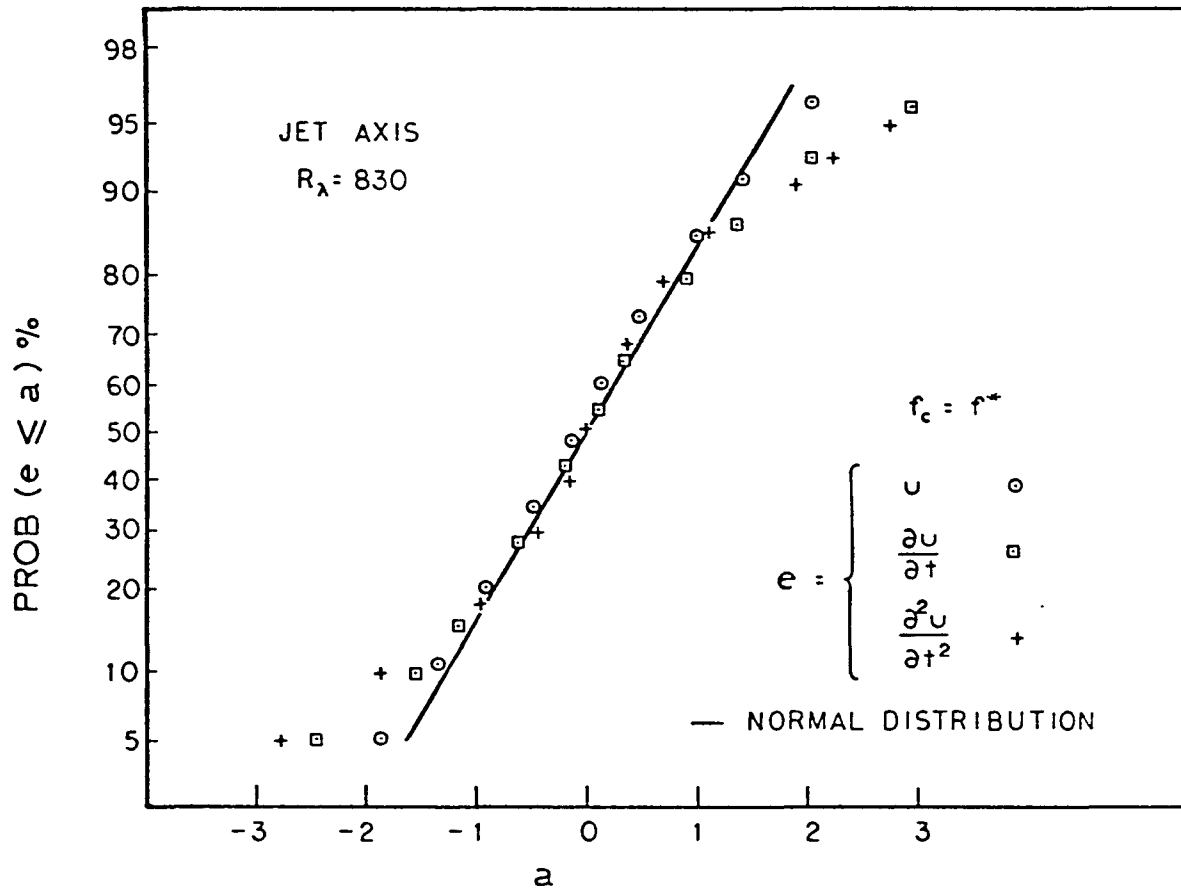


Figure 34. Probability Distributions of u , $\frac{\partial u}{\partial t}$, and $\frac{\partial^2 u}{\partial t^2}$ on the Axis of a Round Jet, $R_\lambda = 830$.

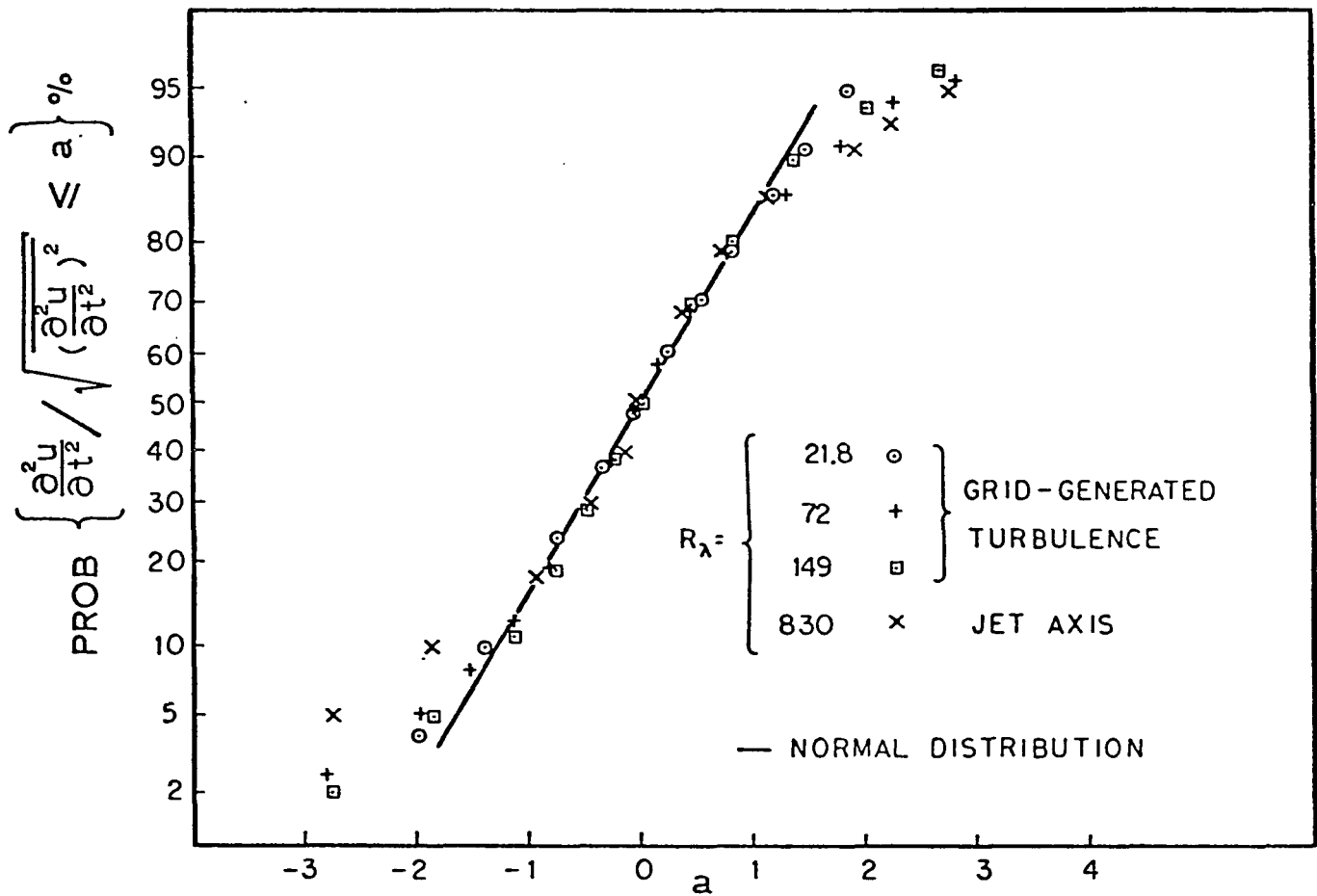


Figure 35. Probability Distributions of $\frac{\partial^2 u}{\partial t^2}$ at Various Reynolds Numbers.

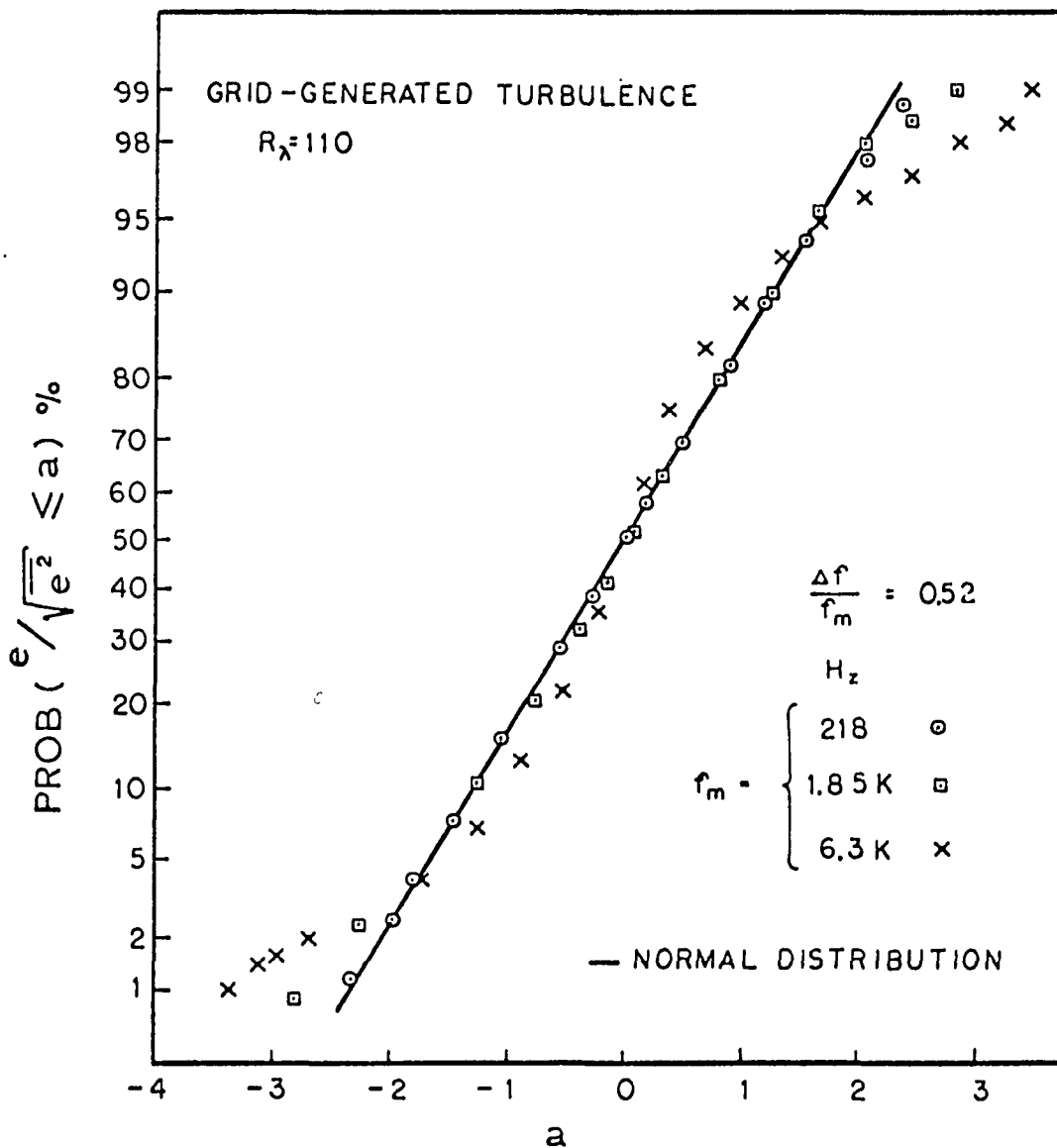


Figure 36. Probability Distribution of Band-pass Signal from a Grid-Generated Turbulence, $R_\lambda = 110$.

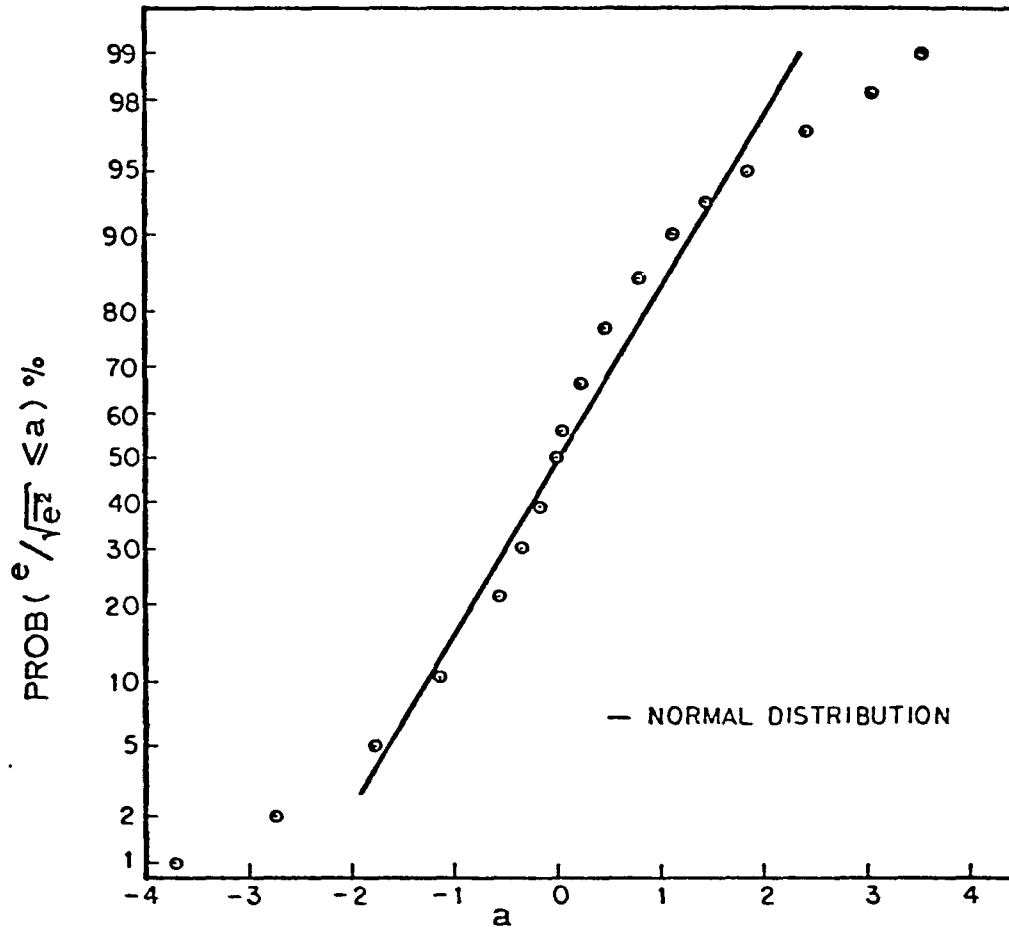


Figure 37. Probability Distribution of High-pass Signal from a Butterworth Filter, $f_c = 5\text{kHz}$

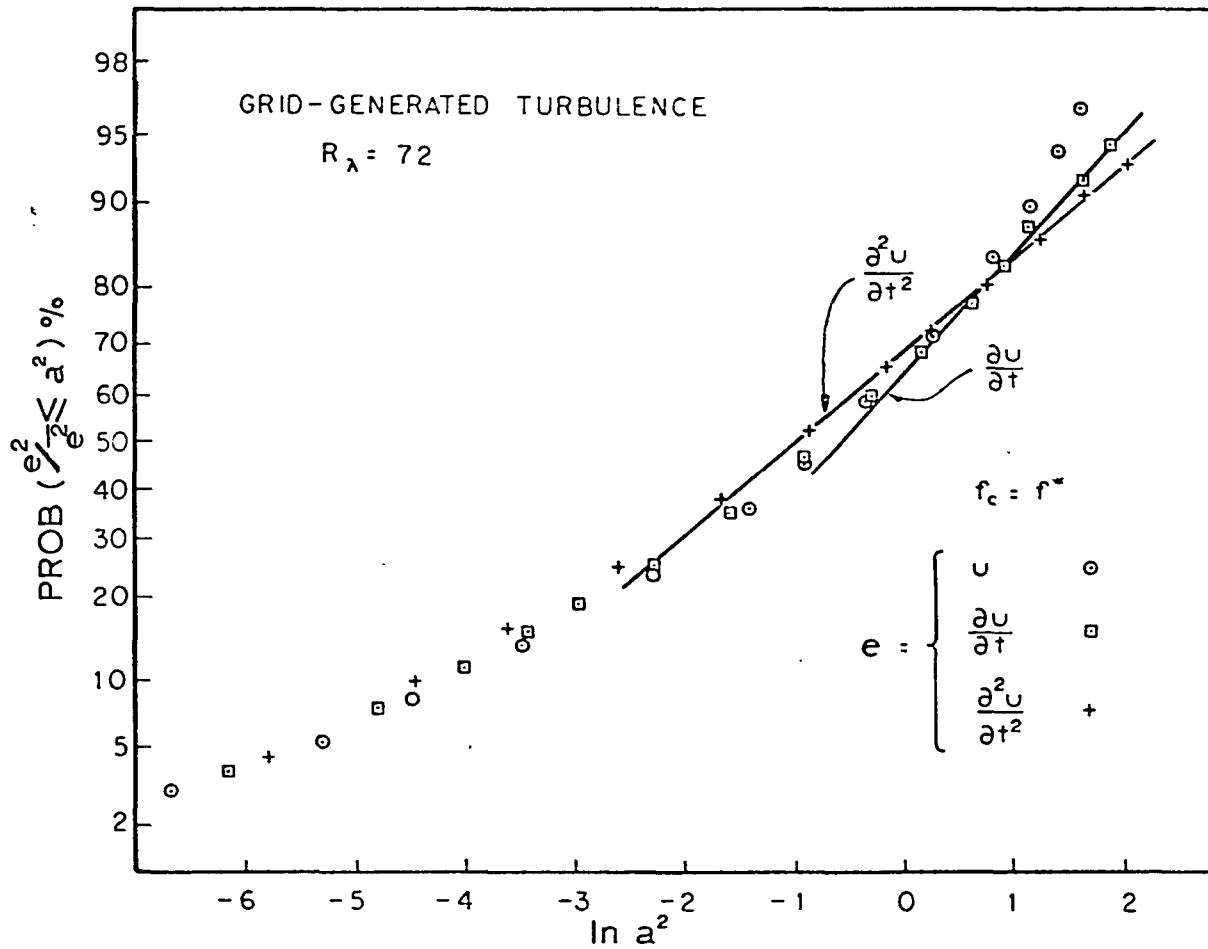


Figure 38. Probability Distributions of u^2 , $(\partial u / \partial t)^2$, and $(\partial^2 u / \partial t^2)^2$ in a Grid-generated Turbulence, $R_\lambda = 72$ (straight lines correspond to Log-normal Distributions).

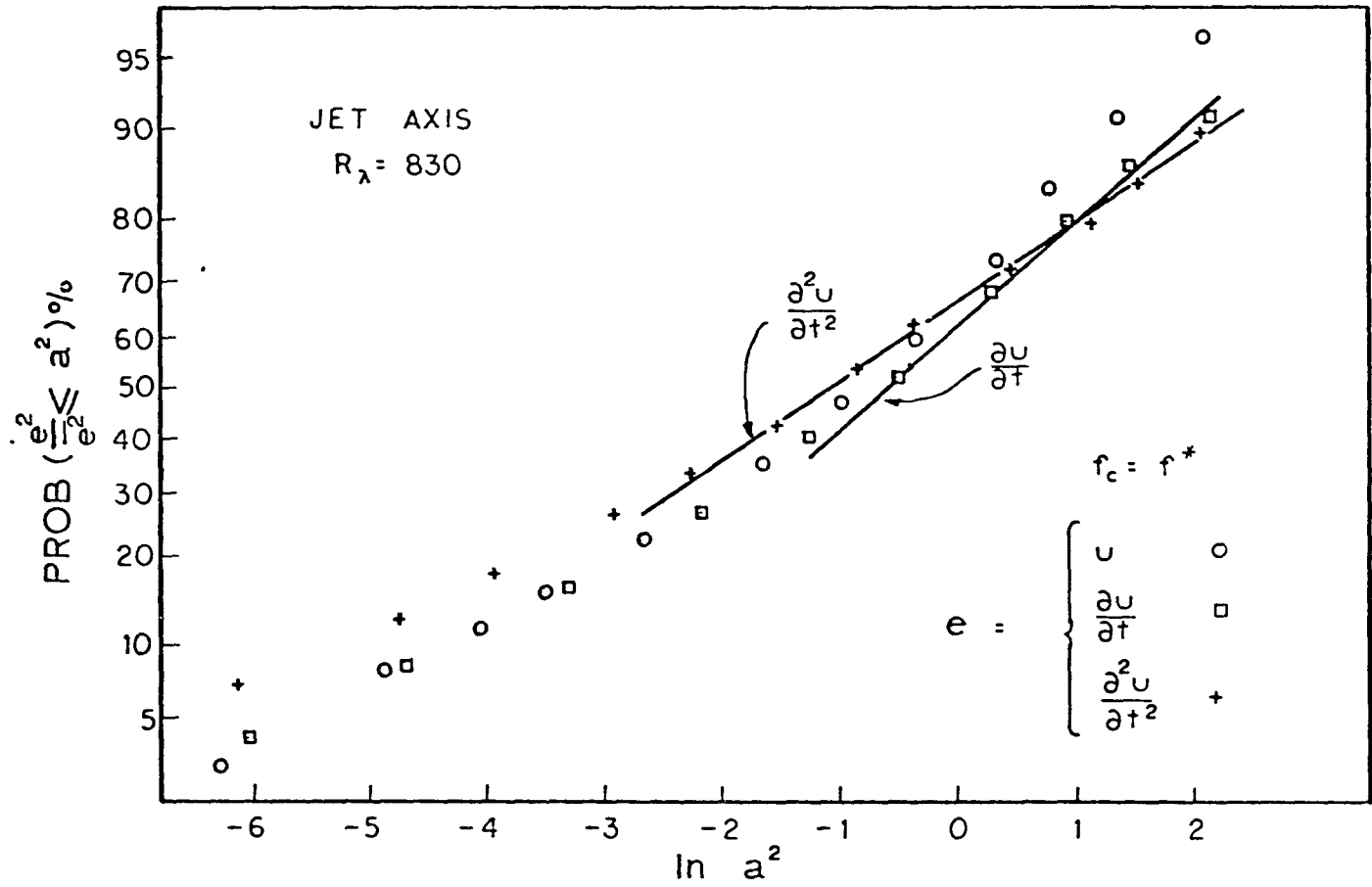


Figure 39. Probability Distributions of u^2 , $(\partial u / \partial t)^2$, and $(\partial^2 u / \partial t^2)^2$ on the Axis of a Round Jet, $R_\lambda = 830$ (straight lines correspond to log-normal distributions).

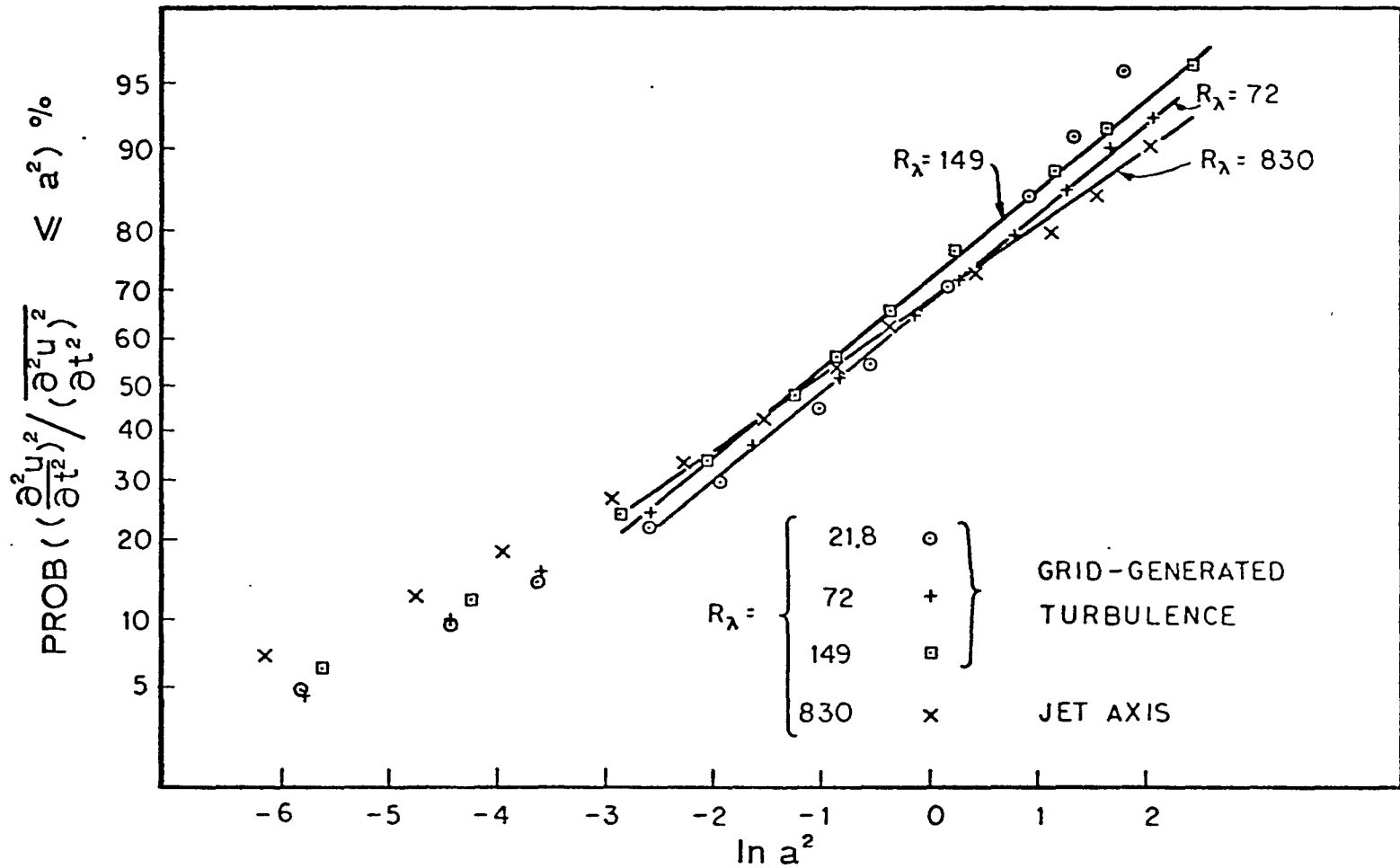


Figure 40. Probability Distributions of $(\frac{\partial^2 u}{\partial t^2})^2$ at Various Reynolds Number.
 (straight lines correspond to log-normal distributions)

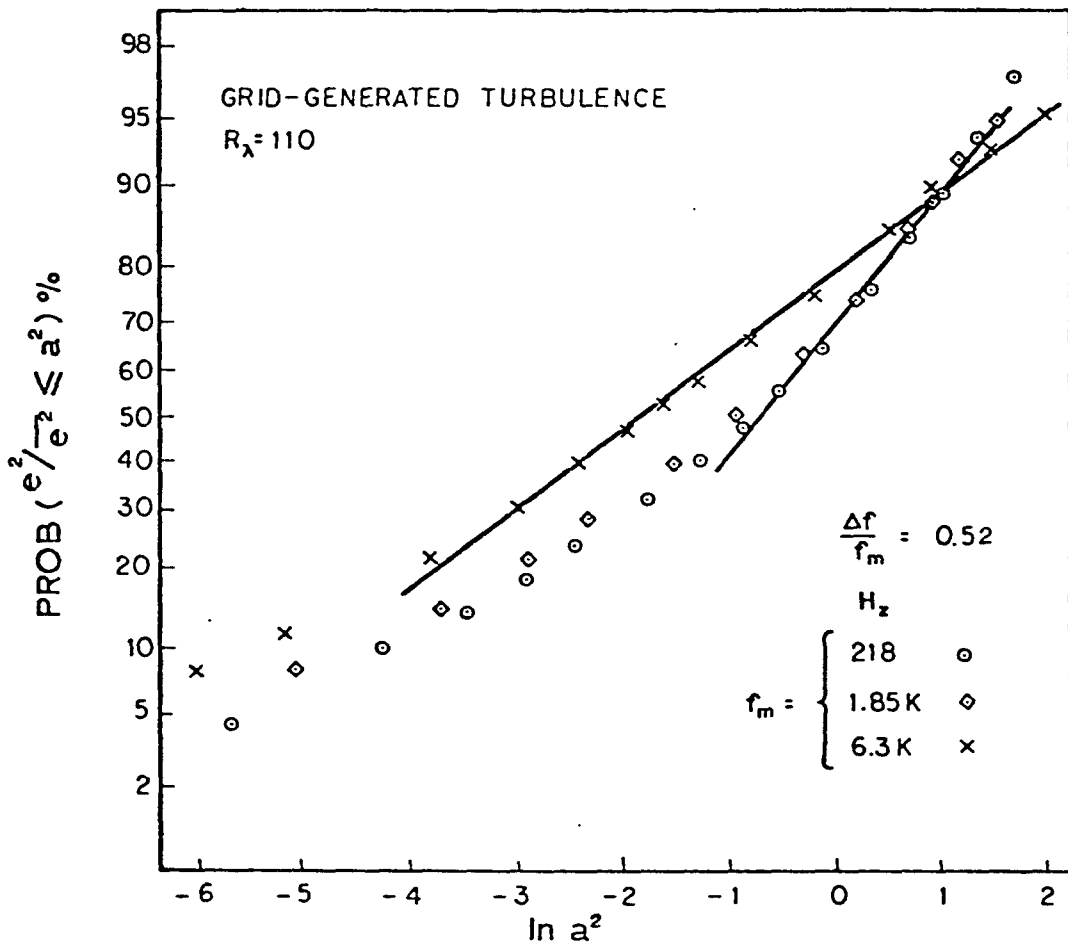


Figure 41. Probability Distributions of the Square of Band-pass Signal from a Grid-generated Turbulence, $R_\lambda = 110$. (straight lines correspond to log-normal distributions)

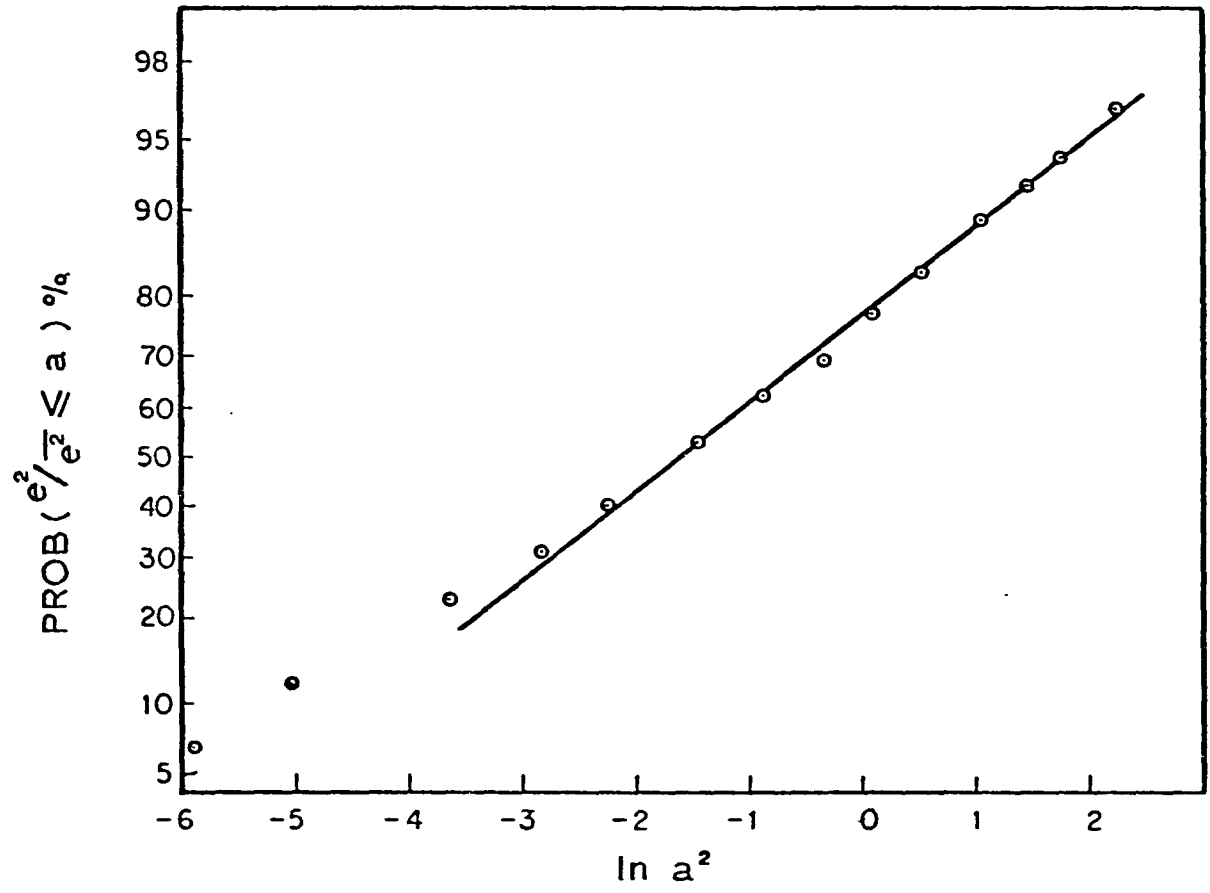


Figure 42. Probability Distribution of the Square of High-pass Signal from a Butterworth Filter, $f_c = 5\text{kHz}$. (straight lines correspond to log-normal distributions)

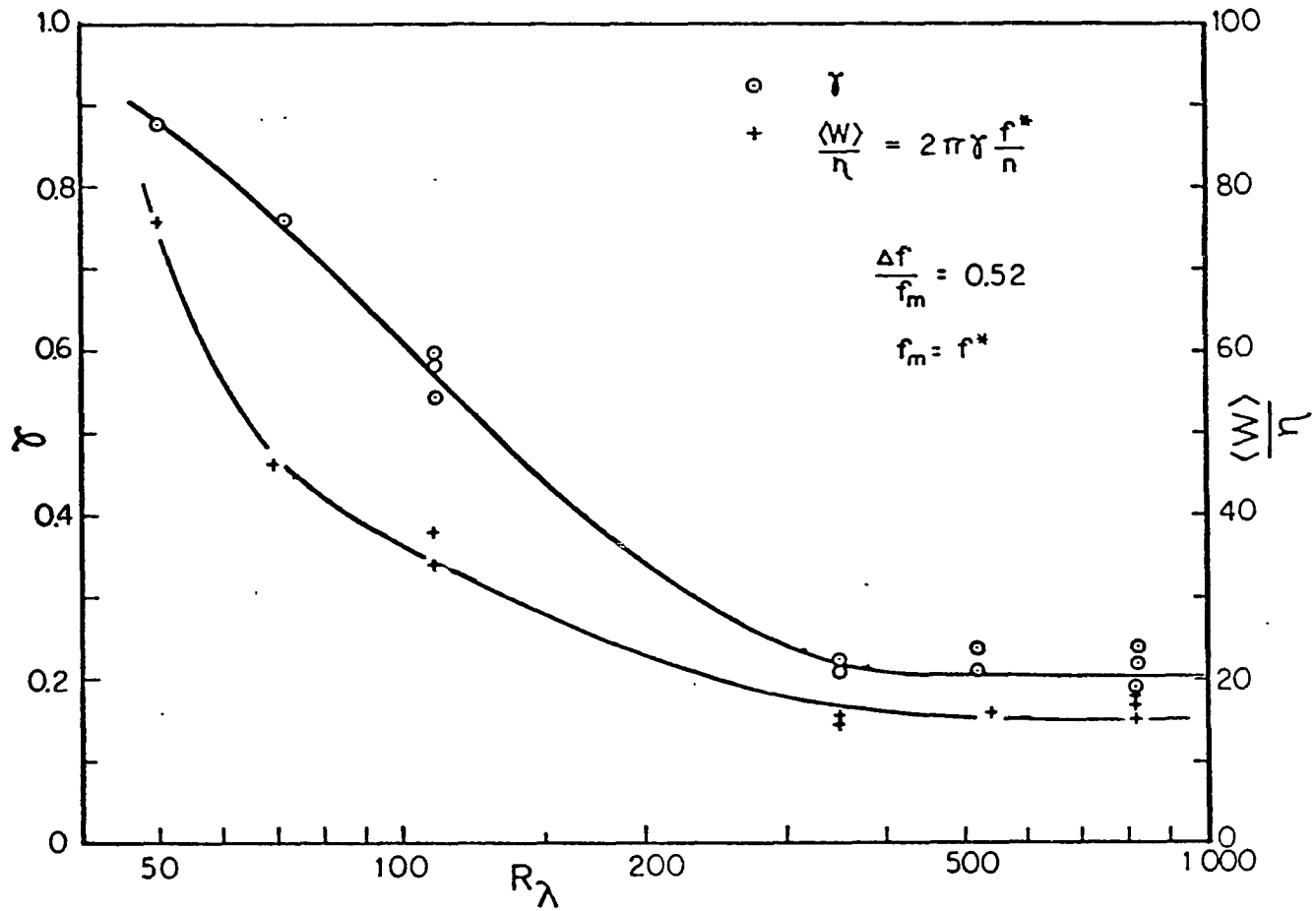


Figure 43. Intermittency Characteristics of Band-pass Signals as Functions of Reynolds Number, R_λ .

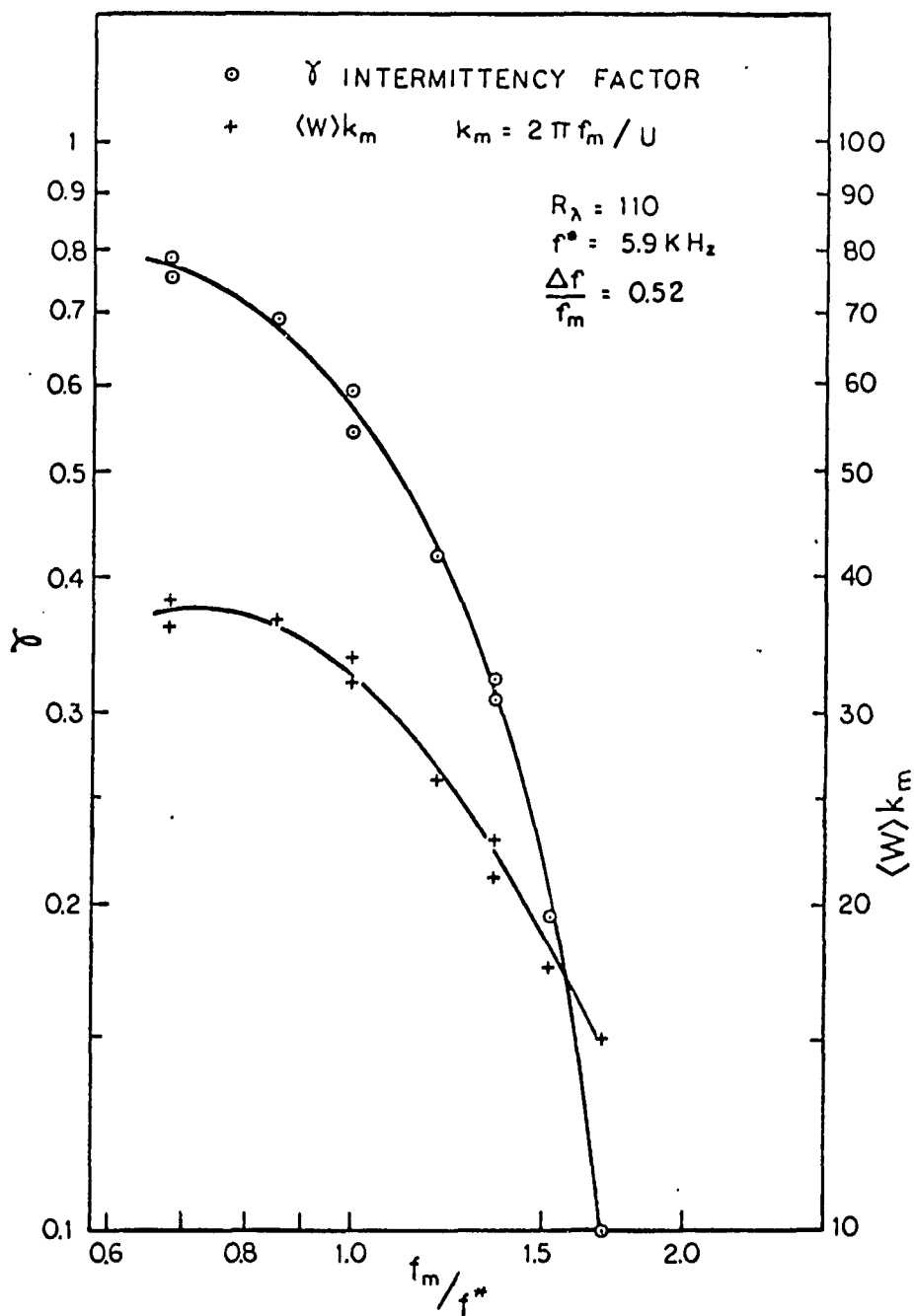


Figure 44(a). Intermittency Characteristics of Band-pass Signals as Functions of Midband Frequency f_m , in Log-log Scale, Grid-generated Turbulence, $R_\lambda = 110$.

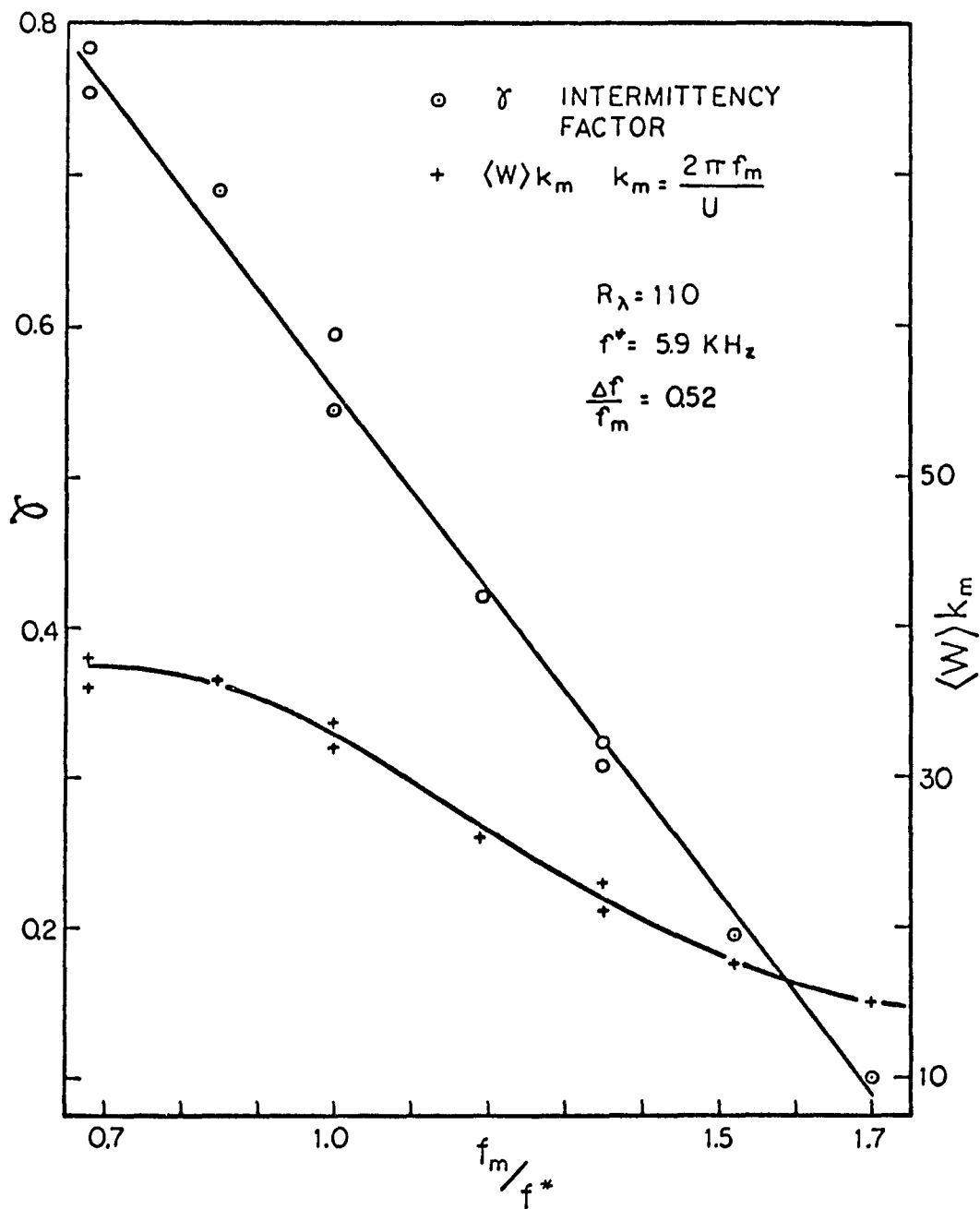


Figure 44(b). Intermittency Characteristics of Band-pass Signals as Functions of Midband Frequency f_m , in Linear Scale, Grid-generated Turbulence, $R_\lambda = 110$.

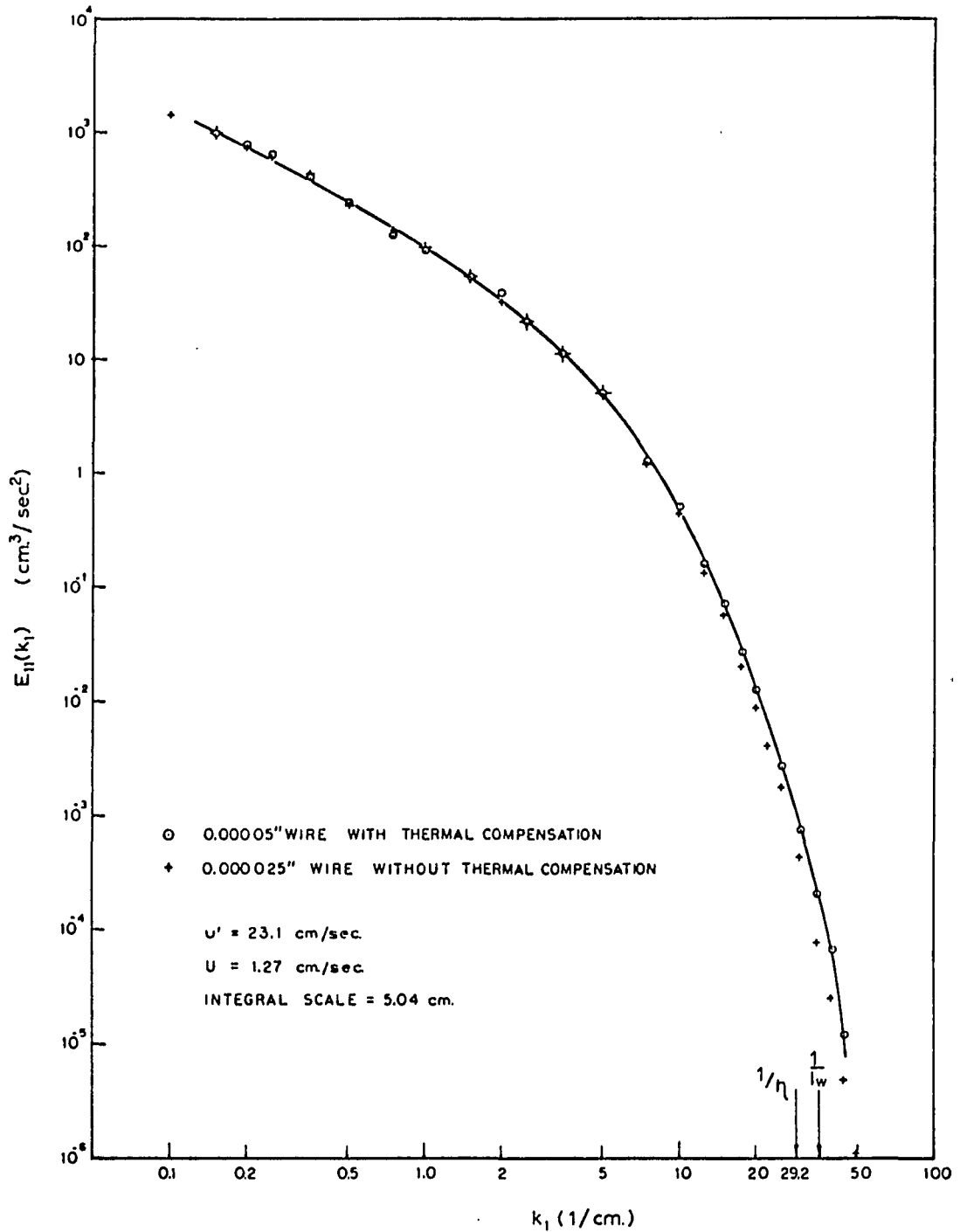


Figure 45. One-dimensional Energy Spectrum of a Grid-generated Turbulence, $R_\lambda = 110$.

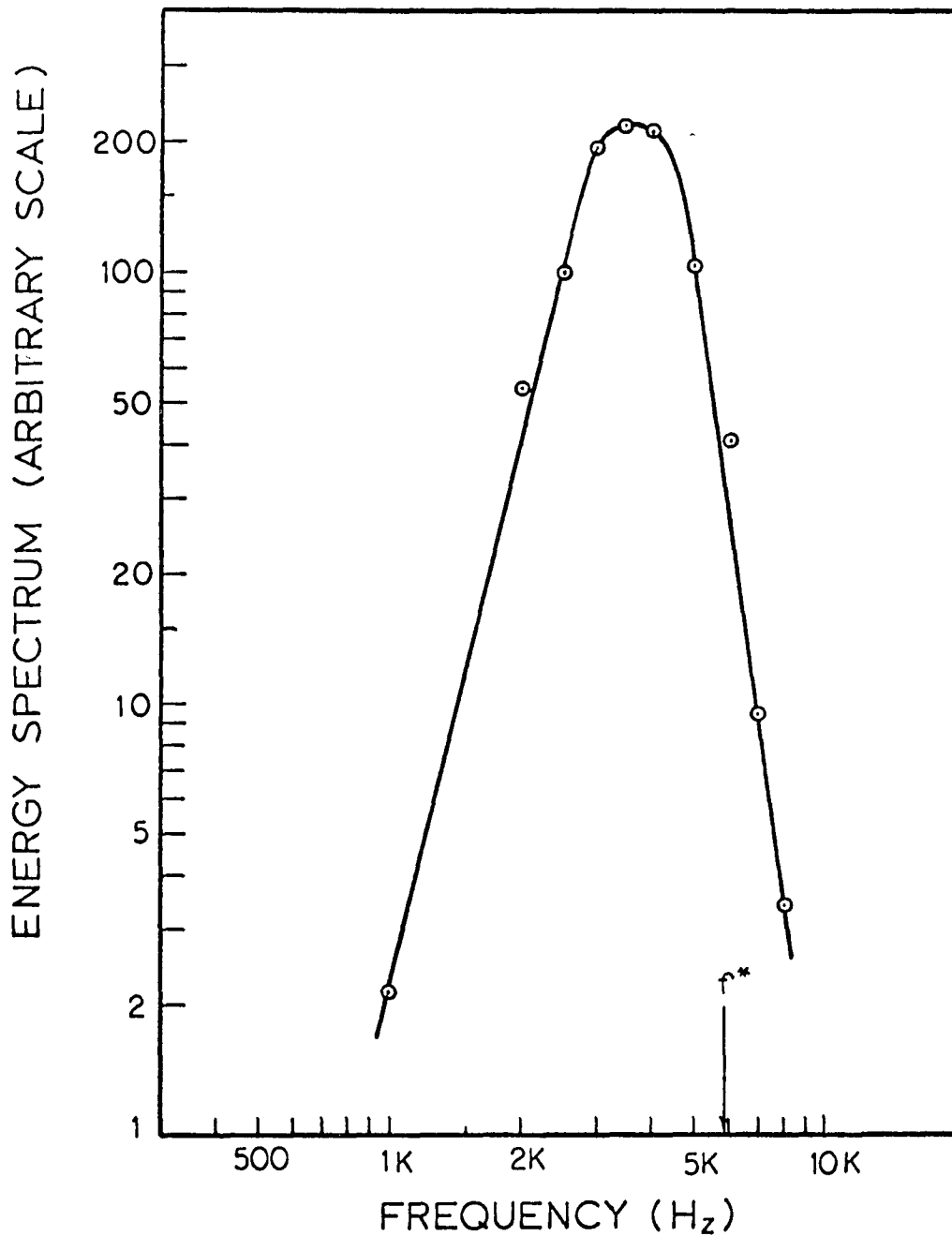


Figure 46. Energy Spectrum of the Signal from High-pass Butterworth Filter

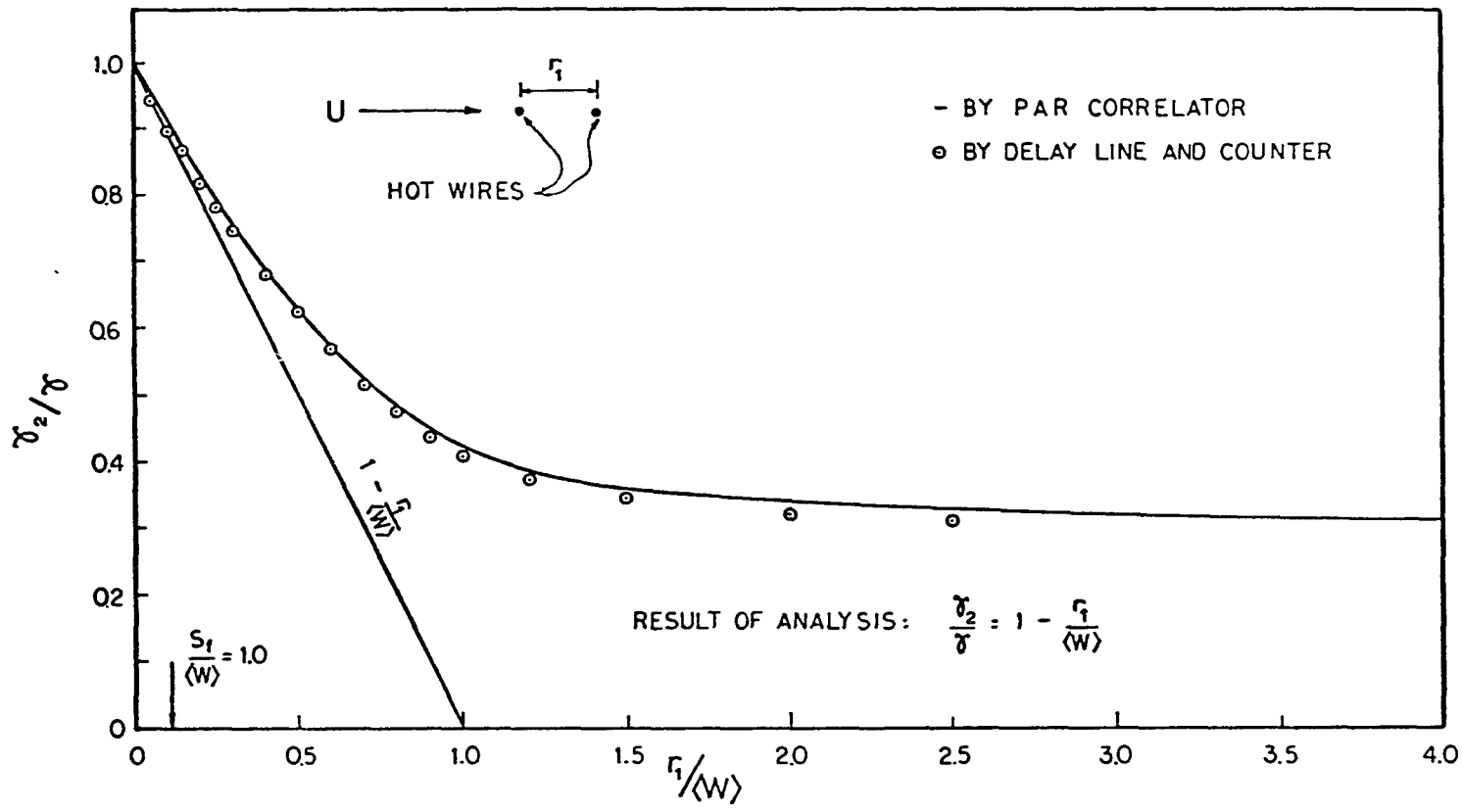


Figure 47. σ_2/σ AS A FUNCTION OF $r_1/\langle W \rangle$

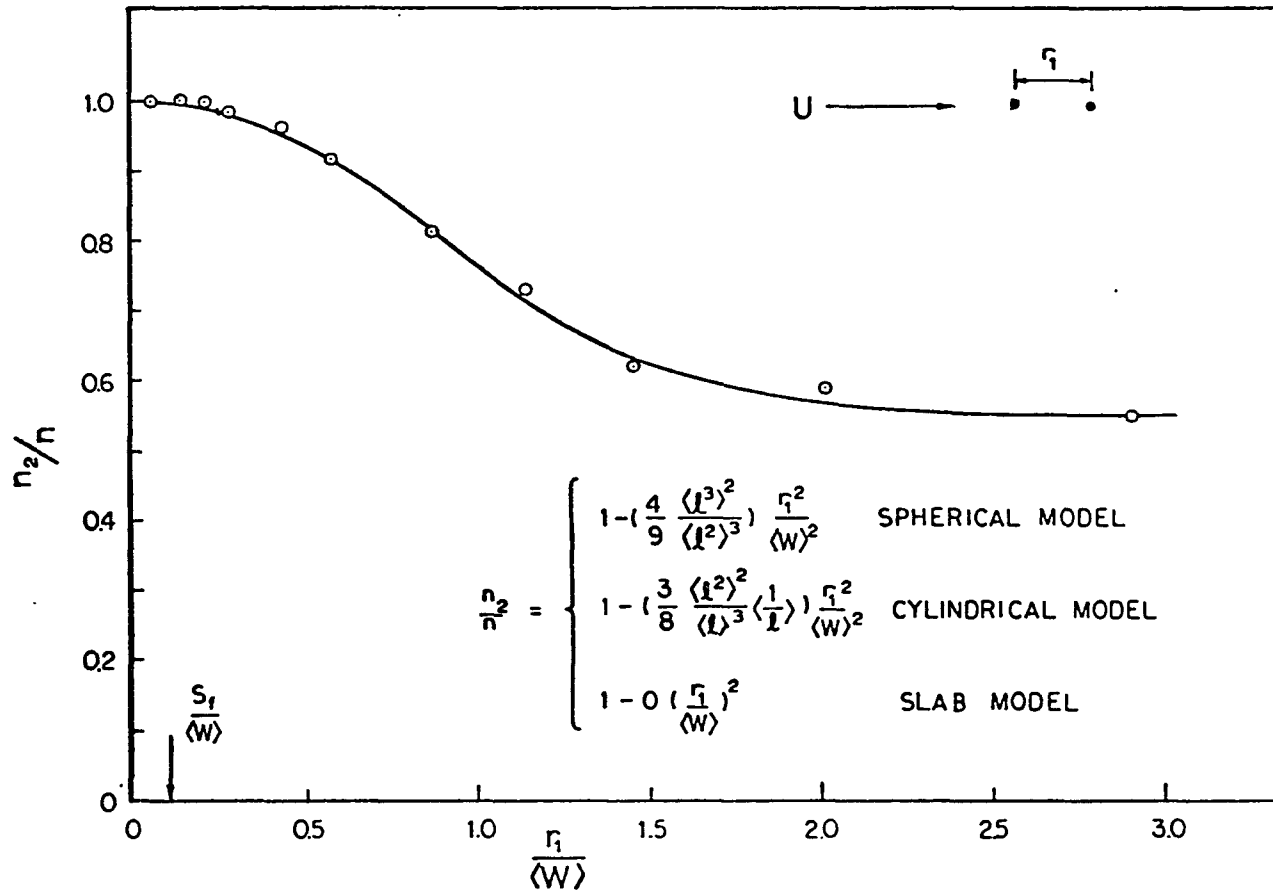


Figure 48. n_2/n AS A FUNCTION OF $r_1/\langle W \rangle$

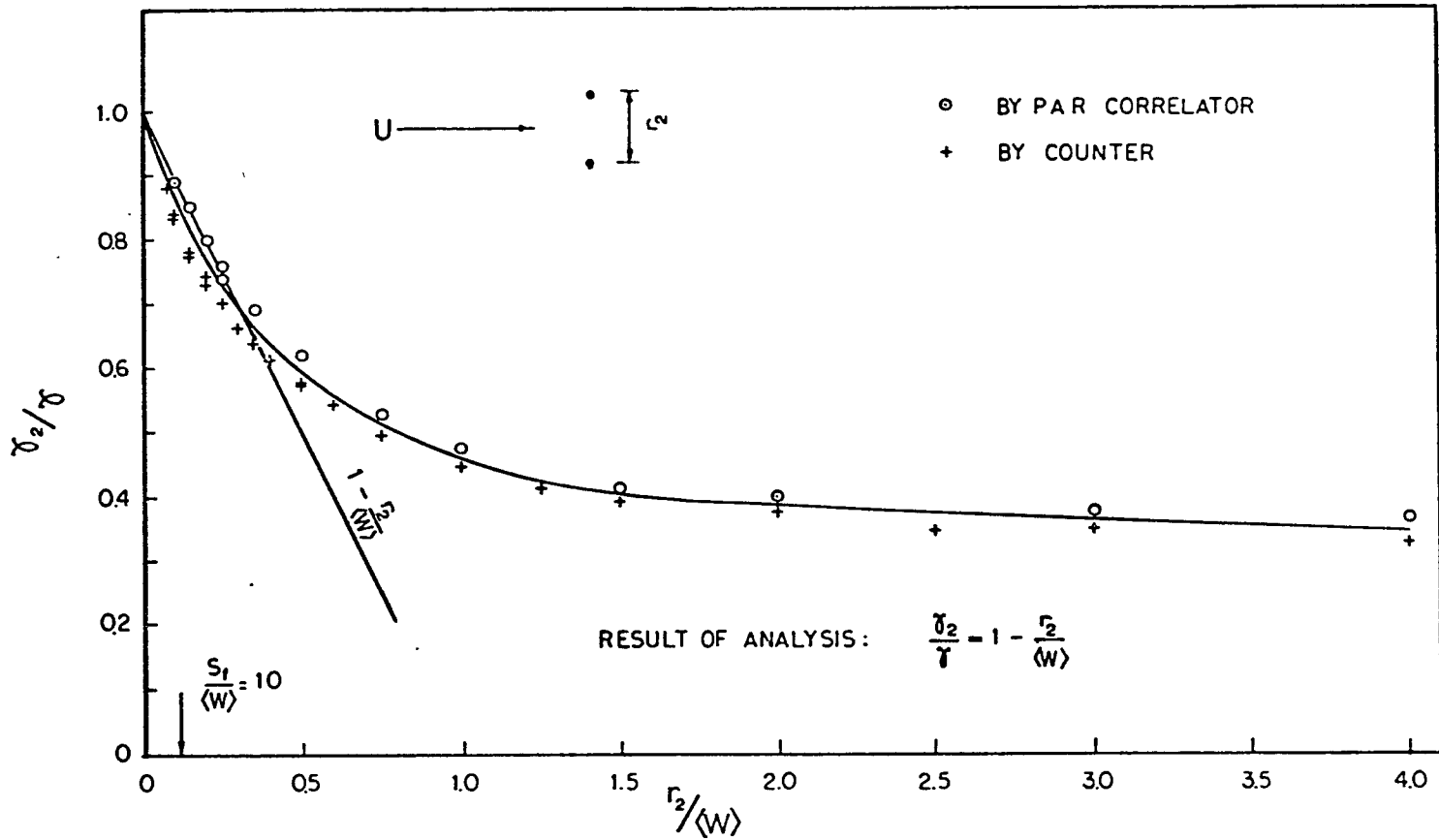


Figure 49. g_2/γ AS A FUNCTION OF $r_2/\langle W \rangle$

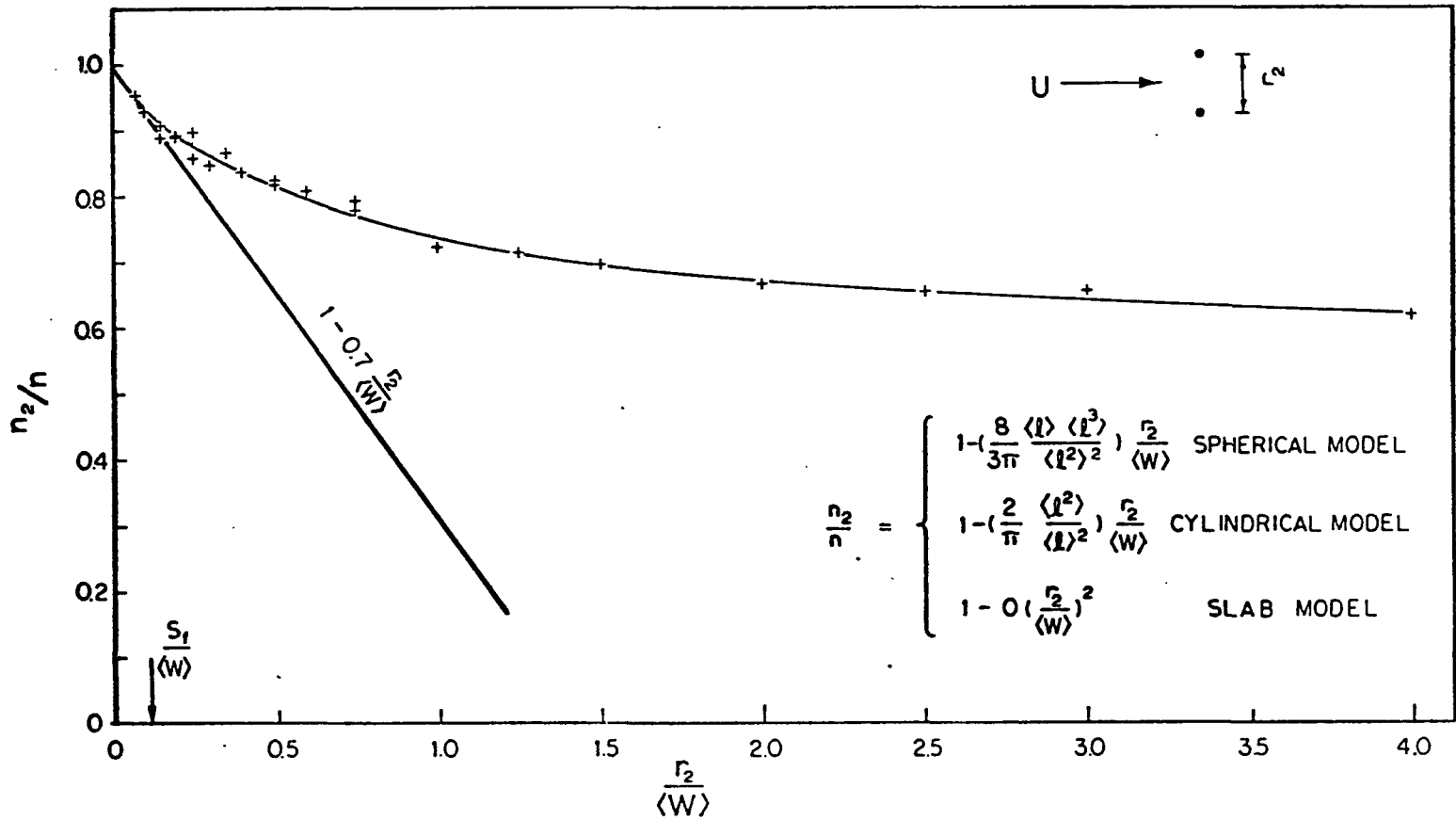


Figure 50. n_2/n AS A FUNCTION OF $r_2/\langle W \rangle$

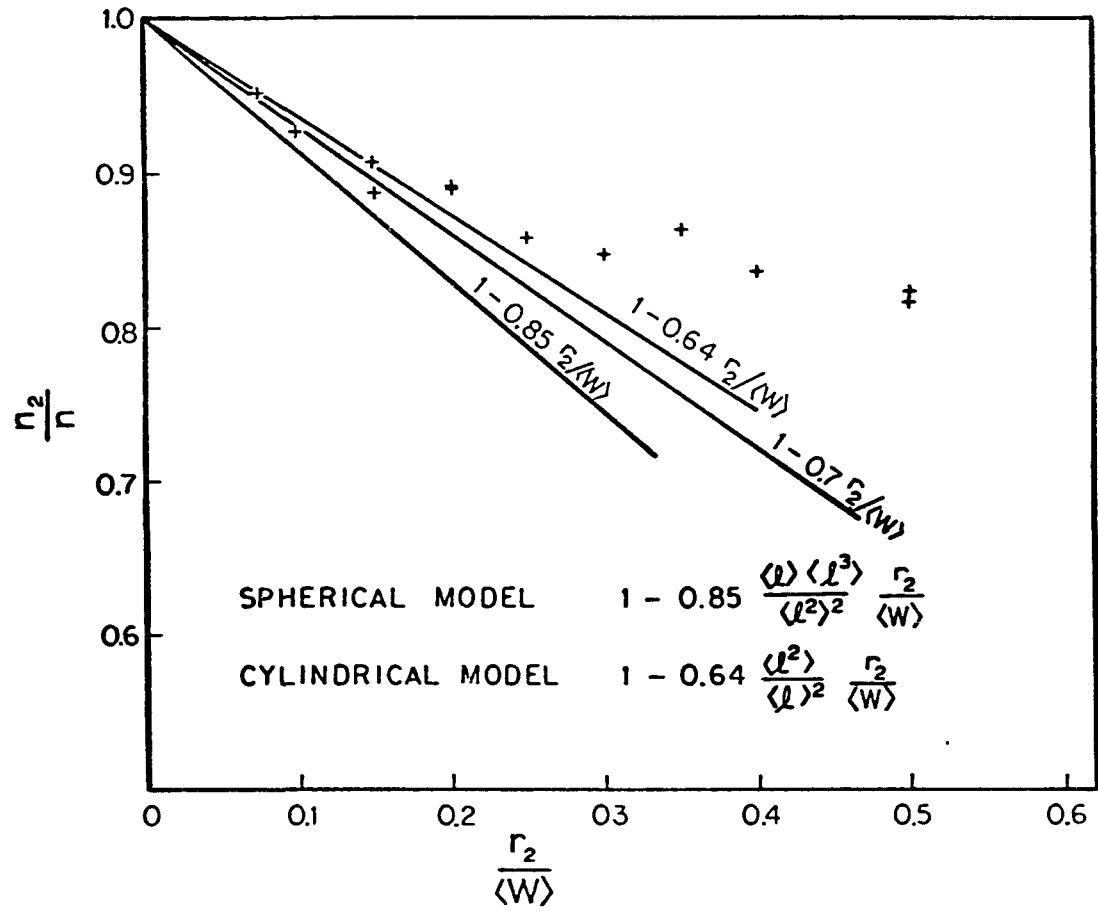


Figure 51. $\frac{\pi_2}{\pi}$ as a Function of $\frac{r_2}{\langle W \rangle}$, with amplified scale.

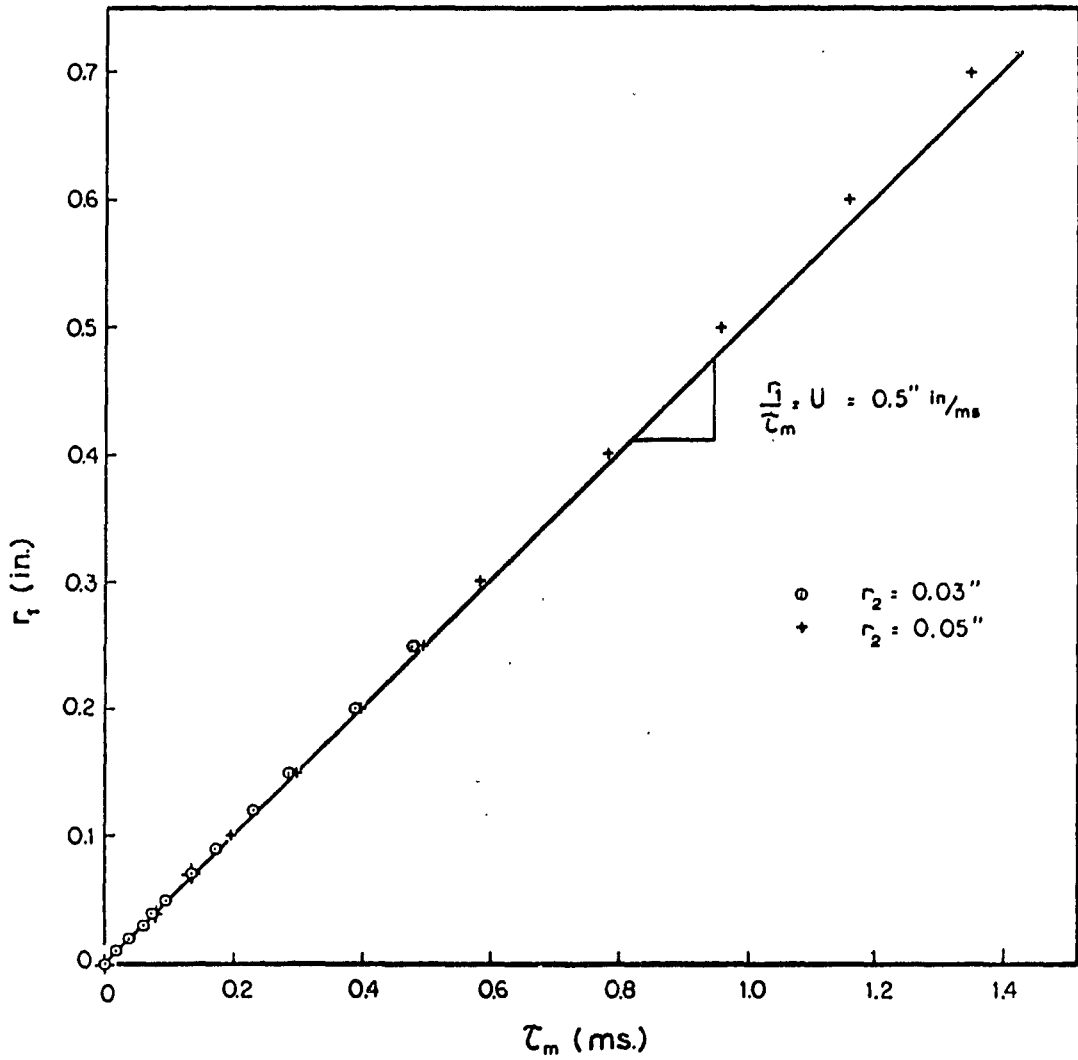


Figure 52. τ_m as Function of r_f for Maximum $R_f(r, r_2, \theta, \tau)$.

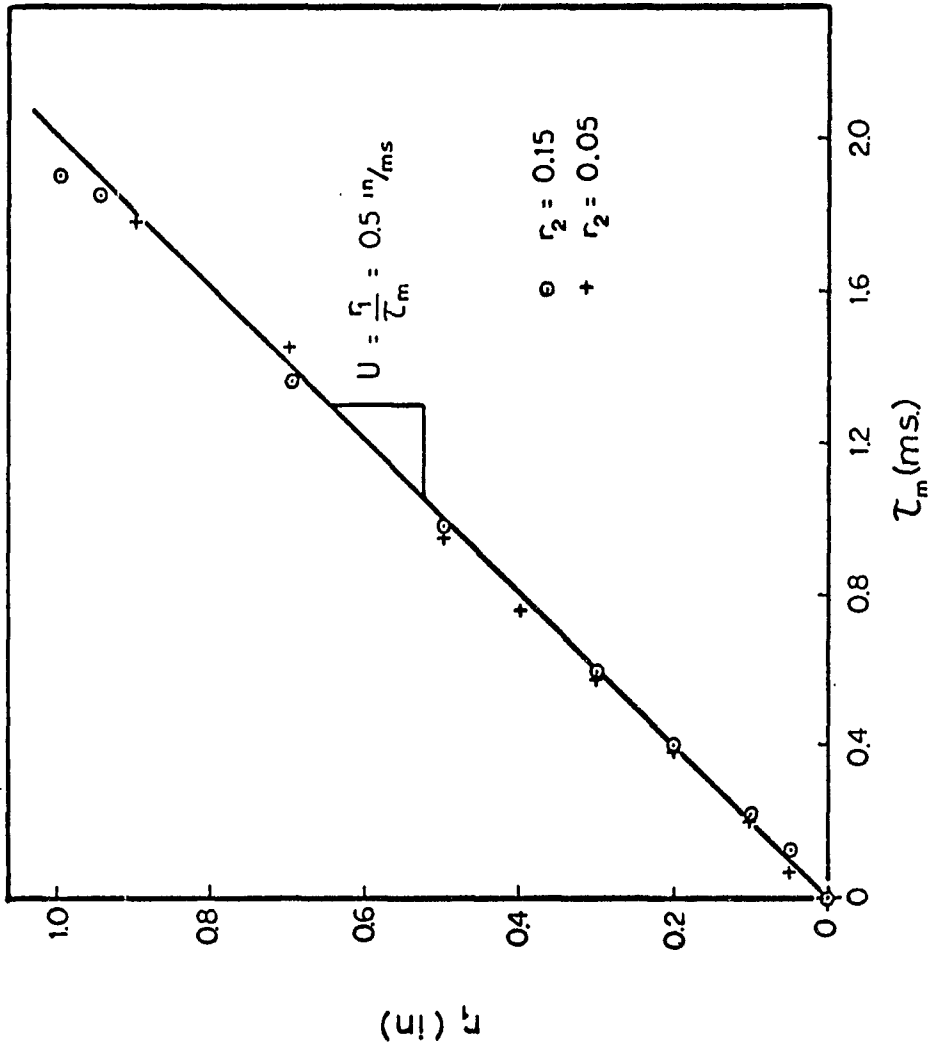
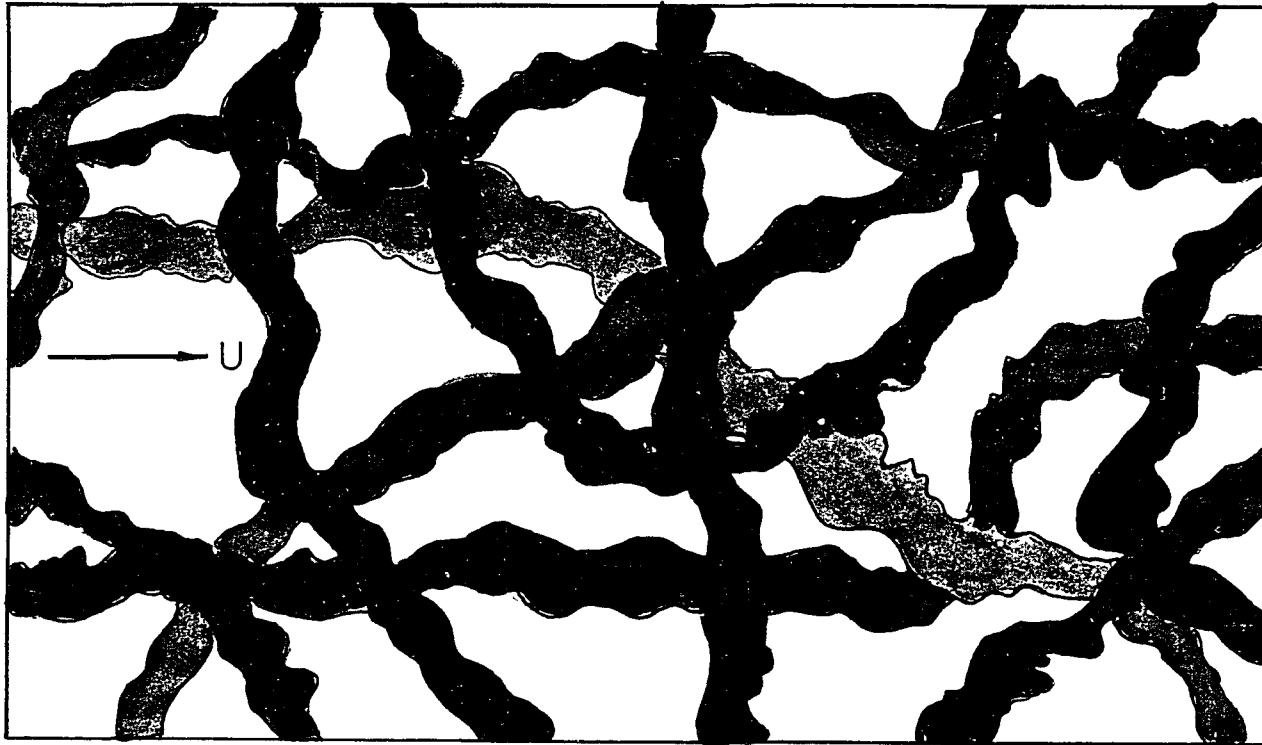


Figure 53. τ_m as Function of r_1 for Maximum $R_I(r_1, r_2, 0, \tau)$.



FINE-STRUCTURE REGION



NO-FINE-STRUCTURE REGION

Figure 54. Qualitative Sketch of Fine-structure Regions

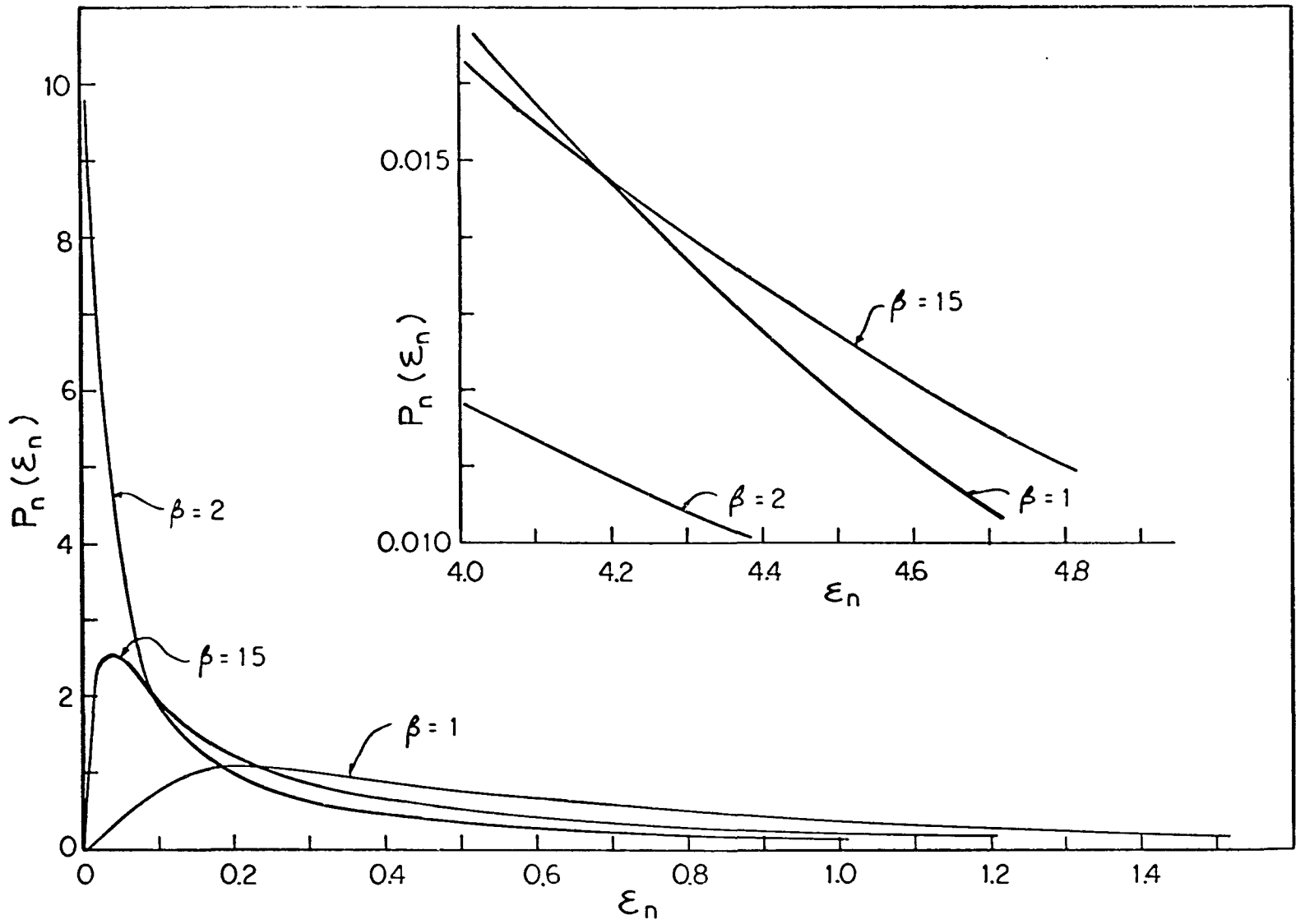


Figure 55. Log-normal Probability Density Curves

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